Jasper De Bock & Gert de Cooman

27 July 2011



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Research group





Jasper De Bock



Gert de Cooman



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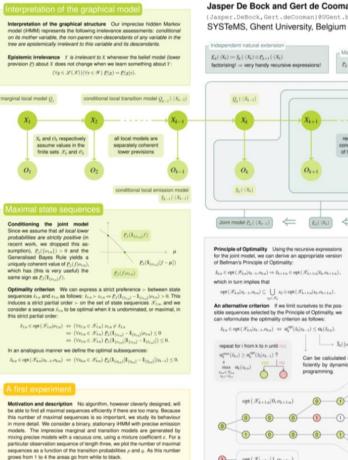








Filip Hermans Erik Quaeghebeur Keivan Shariatmadar Arthur Van Camp





Results We see that there are large regions of transition probability space where the number of maximal elements remains fairly small. The plots also display quite interesting behaviour. If we let the imprecision grow, by using higher r, the areas with multiple maximal sequences become larger. They are expanded versions of the lines of indifference that occur in the precise case.

Jasper De Bock and Gert de Cooman (Jasper.DeBock,Gert.deCooman)@UGent.be

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in the figure below. The se-

quences of green nodes are the

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Marginal extension $\underline{P}_{k}(\cdot|X_{k-1}) := \underline{Q}_{k}(\underline{E}_{k}(\cdot|X_{k})|X_{k-1})$

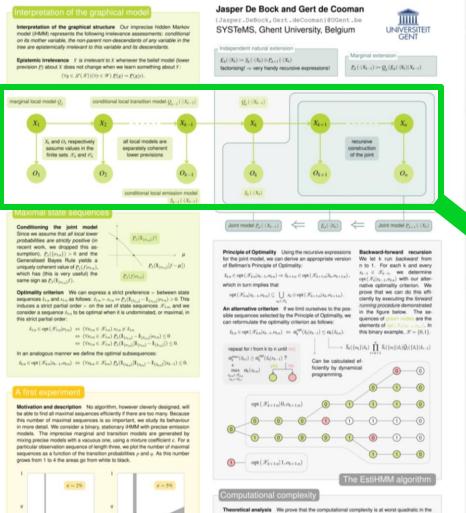
X, recursive construction of the joint O_n O_{k+1} $E_{\delta}(\cdot|\mathbf{X}_{\delta})$ \leftarrow Joint model Part (-(X4) Principle of Optimality Using the recursive expressions Backward-forward recursion for the joint model, we can derive an appropriate version We let k run backward from n to 1. For each k and every $x_{k-1} \in \mathscr{X}_{k-1}$, we determine $f_{k:n} \in \operatorname{opt} \left(\mathscr{X}_{k:n} | s_{k-1}, o_{k:n} \right) \Rightarrow f_{k+1:n} \in \operatorname{opt} \left(\mathscr{X}_{k+1:n} | f_k, o_{k+1:n} \right),$ opt $(\mathcal{X}_{i}|_{x_{k-1},\sigma_{k+1}})$ with our alternative optimality criterion. We prove that we can do this effi $opt(\mathscr{X}_{k:k}|x_{k-1}, o_{k:k}) \subseteq \bigcup x_k \oplus opt(\mathscr{X}_{k+1:k}|x_k, o_{k+1:k}).$ ciently by executing the forward running procedure demonstrated

 $\hat{x}_{k:s} \in \operatorname{opt}\left(\mathscr{X}_{k:s}|x_{k-1}, o_{k:s}\right) \ \Leftrightarrow \ \boldsymbol{a}_{k}^{\operatorname{opt}}(\hat{x}_{k}|x_{k-1}) \leq \boldsymbol{a}_{k}(\hat{x}_{k:s}).$ this binary example, $\mathscr{X} = \{0, 1\}$. $\longrightarrow \overline{S}_k(\{\sigma_k\}|\hat{t}_k) \prod_{i=1}^{n} \overline{S}_i(\{\sigma_i\}|\hat{t}_i)\overline{Q}_i(\{\hat{t}_i\}|\hat{t}_{i-1})$ Can be calculated ef ficiently by dynamical 0 0 programming. 1 -0 1 0 0 (1) opt $(\mathcal{X}_{k+1:n}|1, o_{k+1:n})$ The EstiHMM algorith

Theoretical analysis We prove that the computational complexity is at worst quadratic in the length of the Markov chain, cubic in the number of states, and roughly speaking linear in the number of maximal sequences (each backward step in the backward-forward loop has a linear complexity in the number of maximal elements at that stage).

Empirical confirmation In order to demonstrate that our algorithm is indeed quite efficient we let it determine the maximal sequences for a random output sequence of length 100. The HMM we use to determine the maximal sequences is generated by mixing precise local models with a vacuous one, using a mixture coefficient ϵ . For $\epsilon = 2\%$, there are 5 maximal sequences and it takes 0.2 seconds to calculate them. If we let c grow to 5%, we get 764 maximal sequences and these can be determined in 32 seconds. This demonstrates that the complexity is indeed linear in the number of solutions and that the algorithm can efficiently calculate the maximal sequences even for long output sequences.

State sequence prediction in imprecise hidden Markov models



Results We see that there are large regions of transition probability space where the number of maximal elements remains fairly small. The plots also

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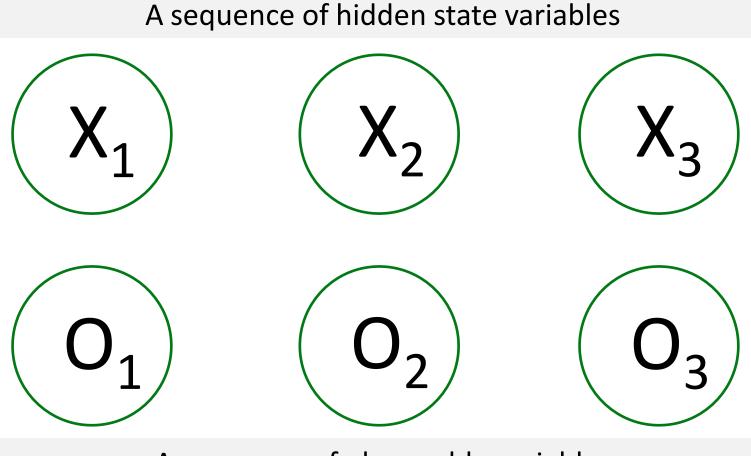
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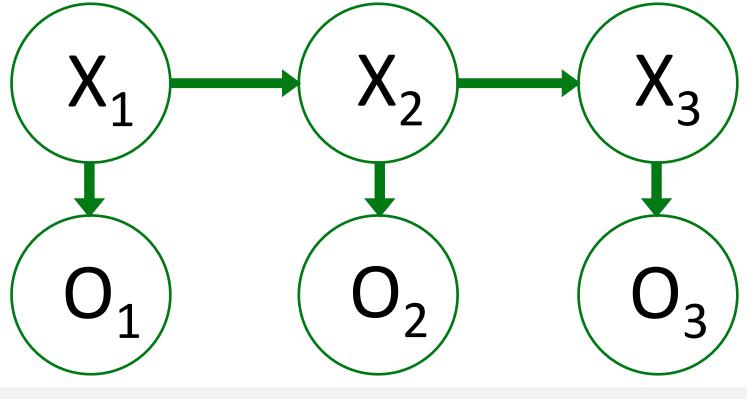
State sequence prediction in imprecise hidden Markov models

The imprecise hidden Markov model

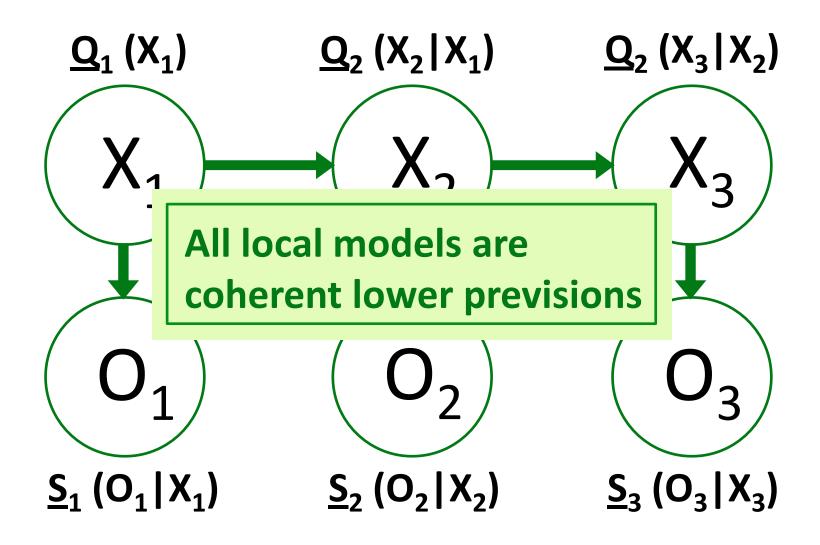


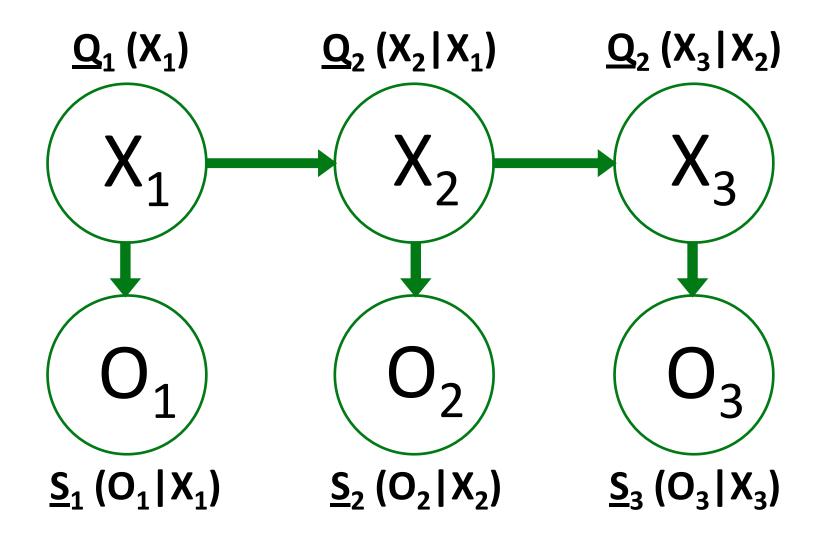
A sequence of observable variables

A sequence of hidden state variables



A sequence of observable variables





State sequence prediction in imprecise hidden Markov models Jasper De Bock and Gert de Cooman terpretation of the graphical mode (Jasper.DeBock,Gert.deCooman)@UGent.be Interpretation of the graphical structure Our imprecise hidden Markov SYSTeMS, Ghent University, Belgium UNIVERSITEIT model (iHMM) represents the following irrelevance assessments: conditional GENT on its mother variable, the non-parent non-descendants of any variable in the tree are epistemically irrelevant to this variable and its descendants. Independent natural extension Marginal extension Epistemic irrelevance Y is irrelevant to X whenever the belief model (lower $\underline{E}_{k}(\cdot|X_{k}) := \underline{S}_{k}(\cdot|X_{k}) \otimes \underline{P}_{k+1}(\cdot|X_{k})$ $\underline{P}_{k}(\cdot|X_{k-1}) := \underline{Q}_{k}(\underline{E}_{k}(\cdot|X_{k})|X_{k-1})$ prevision P) about X does not change when we learn something about Y: factorising! -> very handy recursive expressions $(\forall g \in \mathscr{L}(\mathscr{X}))(\forall y \in \mathscr{F}) P(g) = P(g|y).$ marginal local model Q, conditional local transition model $Q_{k-1}(\cdot|X_{k-2})$ X, and O, respectively all local models are assume values in the separately coherent onstructio finite sets \mathcal{R}_i and \mathcal{O}_i lower previsions O_k O. O1-1 O_{k+1} S. (-(X_1) onditional local emission model $S_{k-1}(-|X_{k-1})$ \leftarrow $E_{\ell}(\cdot|X_{\ell})$ Joint model PA (-|Xi-1) Joint model Part (-(X4) Conditioning the joint model Since we assume that all local lower P. (Line) probabilities are strictly positive (in recent work, we dropped this as-Principle of Optimality Using the recursive expressions Backward-forward recursion sumption), $P_1(\{\sigma_{1,n}\}) > 0$ and the Generalised Bayes Rule yields a for the joint model, we can derive an appropriate version We let k run backward from of Bellman's Principle of Optimality: $\underline{P}_t(\mathbf{1}_{\{r_{1,n}\}}[f-\mu])$ n to 1. For each k and every uniquely coherent value of P₁(f|o1,*) $x_{k-1} \in \mathscr{X}_{k-1}$, we determine which has (this is very useful) the $I_{k,n} \in \operatorname{opt} \left(\mathscr{X}_{k,n} | \mathbf{x}_{k-1}, \mathbf{o}_{k,n} \right) \Rightarrow I_{k+1,n} \in \operatorname{opt} \left(\mathscr{X}_{k+1,n} | I_k, \mathbf{o}_{k+1,n} \right),$ P. (Dera) opt $(\mathcal{X}_{i}|_{x_{k-1},\sigma_{k+1}})$ with our alter-which in turn implies that native optimality criterion. We prove that we can do this effi-Optimality criterion We can express a strict preference > between state $opt(\mathscr{X}_{k:k}|x_{k-1}, o_{k:k}) \subseteq \bigcup x_k \oplus opt(\mathscr{X}_{k+1:k}|x_k, o_{k+1:k}).$

sequences $\hat{x}_{1:n}$ and $x_{1:n}$ as follows: $\hat{x}_{1:n} \succ x_{1:n} \Leftrightarrow P_1(\mathbf{I}_{\{j_{1:n}\}} - \mathbf{I}_{\{x_{1:n}\}}|o_{1:n}) > 0$. This induces a strict partial order \succ on the set of state sequences \mathscr{X}_{1m} and we consider a sequence *i*_{1,n} to be optimal when it is undominated, or maximal, in this strict partial order:

 $\hat{x}_{1:n} \in \operatorname{opt}(\mathscr{X}_{1:n}|\sigma_{1:n}) \iff (\forall x_{1:n} \in \mathscr{X}_{1:n}) x_{1:n} \neq \hat{x}_{1:n}$

 $\Leftrightarrow \ (\forall x_{1:a} \in \mathscr{X}_{1:a}) \ \underline{P}_1(\mathbf{I}_{\{x_{1:a}\}} - \mathbf{I}_{\{x_{1:a}\}} | o_{1:a}) \leq 0$ \Leftrightarrow $(\forall x_{l:n} \in \mathcal{X}_{l:n}) \underline{P}_1(\mathbf{I}_{\{\sigma_{l:n}\}}[\mathbf{I}_{\{\sigma_{l:n}\}} - \mathbf{I}_{\{\delta_{l:n}\}}]) \leq 0.$

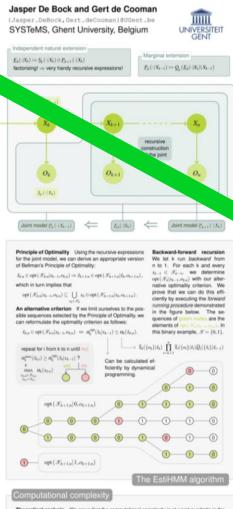
In an analogous manner we define the optimal subsequences.

 $i_{k,n} \in \operatorname{opt}\left(\mathscr{X}_{k,n} | x_{k-1}, o_{k,n}\right) \iff \left(\forall x_{k,n} \in \mathscr{X}_{k,n}\right) \underbrace{P_k}(\mathbf{I}_{\{o_{k,n}\}} [\mathbf{I}_{\{o_{k,n}\}} - \mathbf{I}_{\{\delta_{k,n}\}}] | x_{k-1}) \leq 0.$

Motivation and description No algorithm, however cleverly designed, will be able to find all maximal sequences efficiently if there are too many. Because this number of maximal sequences is so important, we study its behaviour in more detail. We consider a binary, stationary iHMM with precise emission models. The imprecise marginal and transition models are generated by mixing precise models with a vacuous one, using a mixture coefficient r. For a particular observation sequence of length three, we plot the number of maximal sequences as a function of the transition probabilities p and q. As this number grows from 1 to 4 the areas go from white to black.



Results We see that there are large regions of transition probability space where the number of maximal elements remains fairly small. The plots also display quite interesting behaviour. If we let the imprecision grow, by using higher r, the areas with multiple maximal sequences become larger. They are expanded versions of the lines of indifference that occur in the precise case.



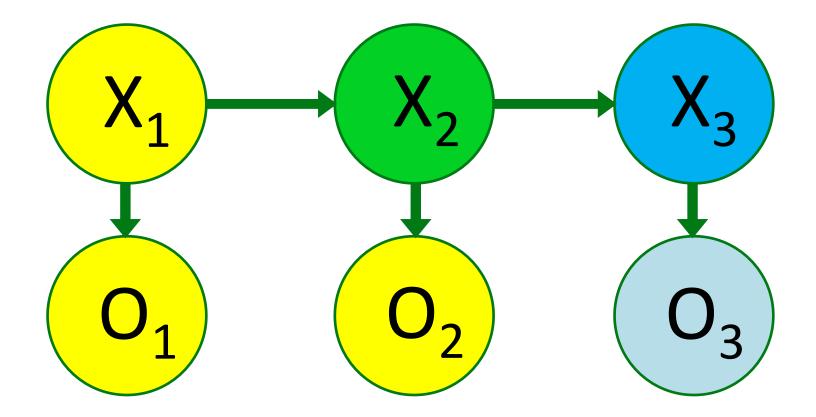
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Empirical confirmation In order to demonstrate that our algorithm is indeed quite efficient we let it determine the maximal sequences for a random output sequence of length 100. The HMM we use to determine the maximal sequences is generated by mixing precise local models with a vacuous one, using a mixture coefficient ϵ . For $\epsilon = 2\%$, there are 5 maximal sequences and it takes 0.2 seconds to calculate them. If we let c grow to 5%, we get 764 maximal sequences and these can be determined in 32 seconds. This demonstrates that the complexity is indeed linear in the number of solutions and that the algorithm can efficiently calculate the maximal sequences even for long output sequences.

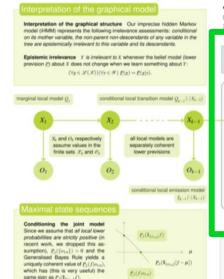
State sequence prediction in imprecise hidden Markov models

Epistemic Irrelevance

Epistemic irrelevance



Conditional on its **mother variable**, the **non-parent non**descendants of any variable in the tree are epistemically irrelevant to this variable and its descendants



Optimality criterion. We can express a strict preference — between state sequences $h_{i,n}$ and $\pi_{i,n}$ as follows: $\pi_{i,n} \to \pi_{i,n} \otimes P_i(a_{i,n,n}) = h_{i,n,n}(a_{i,n,n}) = 0$. This induces a strict partial order \succ on the set of state sequences $\pi_{i,n}$ and we consider a sequence $\pi_{i,n}$ to be optimal when it is undominated, or maximal, in this strict partial order:

- $\hat{x}_{1:n} \in \operatorname{opt} \left(\mathscr{X}_{1:n} | \sigma_{1:n} \right) \iff \left(\forall x_{1:n} \in \mathscr{X}_{1:n} \right) x_{1:n} \neq \hat{x}_{1:n}$
 - $\begin{array}{l} \Leftrightarrow \ \left(\forall x_{1:n} \in \mathscr{X}_{1:n} \right) \underbrace{P_1(\mathbf{I}_{\{x_{1:n}\}} \mathbf{I}_{\{\hat{x}_{1:n}\}} | \sigma_{1:n}) \leq 0 \\ \Leftrightarrow \ \left(\forall x_{1:n} \in \mathscr{X}_{1:n} \right) \underbrace{P_1(\mathbf{I}_{\{x_{1:n}\}} | \mathbf{I}_{\{x_{1:n}\}} \mathbf{I}_{\{\hat{x}_{1:n}\}} |) \leq 0. \end{array}$

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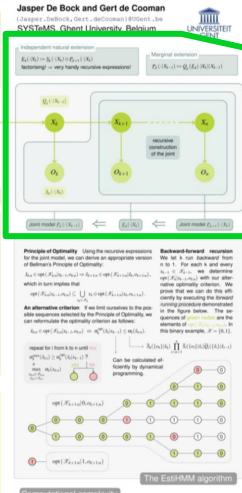
 $f_{k,a} \in \operatorname{opt}\left(\mathscr{X}_{k,a}|x_{k-1}, o_{k,a}\right) \iff (\forall x_{k,a} \in \mathscr{X}_{k,a}) \ \underline{P}_k(\mathbf{I}_{\{o_{k-1}\}}|\mathbf{I}_{\{o_{k-1}\}} - \mathbf{I}_{\{b_{k-1}\}}]|x_{k-1}) \leq 0.$

A first experiment

Motivation and description. No algorithm, however cleverly designed, will be able to find all maximal sequences efficiently three are too many. Because this number of maximal sequences is so important, we study its behaviour in more detail. We consider a binary, stationary iHMM with precise emission models. The improvise marginal and transition models are generated by mixing precise models with a vacous one, using a mixture coefficient *e*. For a particular coservation sequence of length three, we point the number of maximal sequences as a function of the transition probabilities *p* and *g*. As this number grows from 11 to the areaso for monitor binary.



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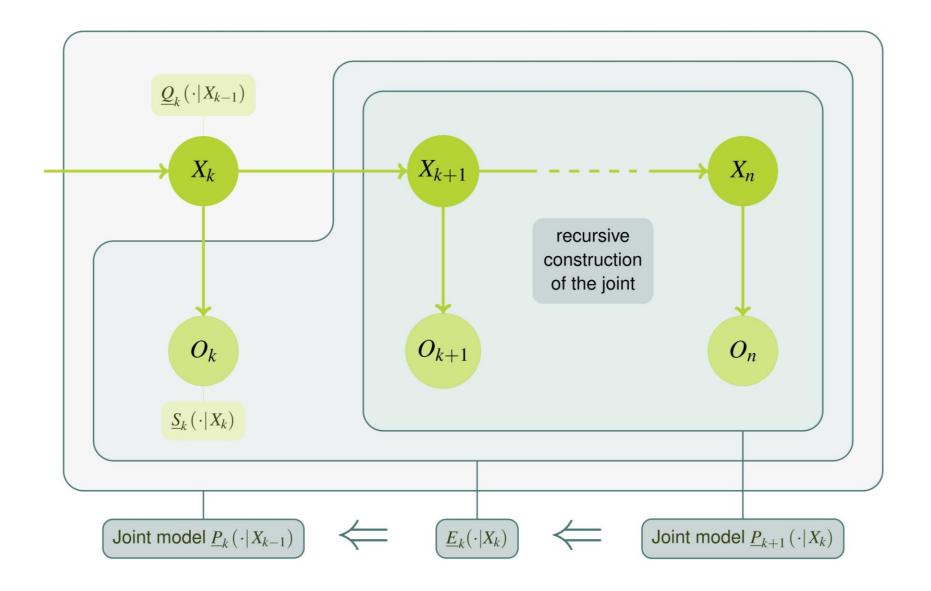
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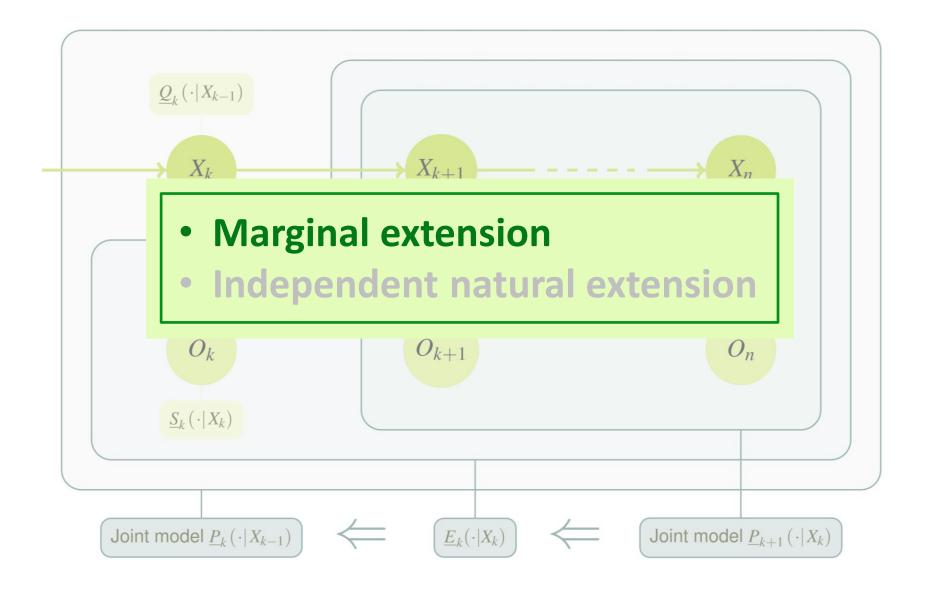
State sequence prediction in imprecise hidden Markov models

Recursive construction of a joint model for the imprecise hidden Markov model

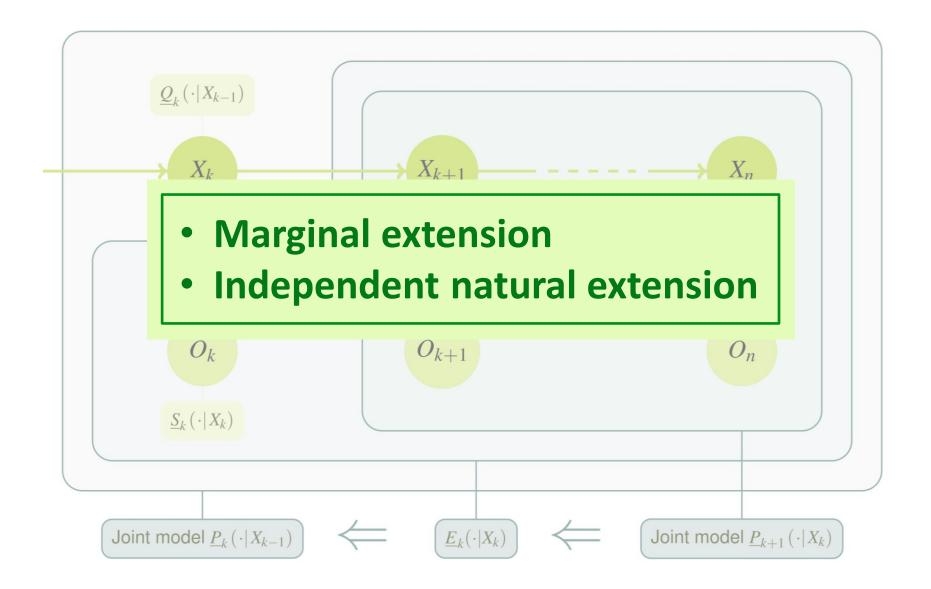
Recursive construction of a joint model

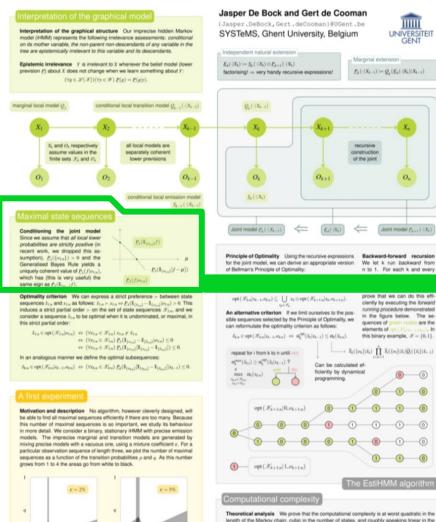


Recursive construction of a joint model



Recursive construction of a joint model





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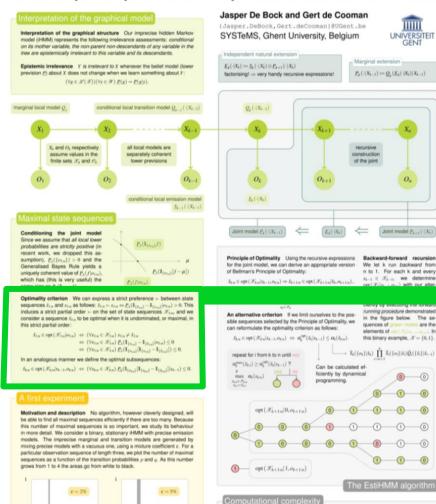
State sequence prediction in imprecise hidden Markov models

Conditioning the model on the observations

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Conditioning the model on the observations

Generalised Bayes rule: An extension of the Bayes rule to imprecise probabilities



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Joint model Part (-(X4)

Backward-forward recursion

We let k run backward from

n to 1. For each k and every

 $s_{k-1} \in \mathscr{X}_{k-1}$, we determine

running procedure demonstrated

in the figure below. The se-

quences of green nodes are the

this binary example, $\mathcal{X} = \{0, 1\}$.

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State sequence prediction in imprecise hidden Markov models

Maximal state sequences

We predict the state sequence by calculating a set of optimal sequences

Notion of optimality: maximality

Strict partial ordening:

 $\hat{x}_{1:n} \succ x_{1:n} \Leftrightarrow \underline{P}_1(\mathbb{I}_{\{\hat{x}_{1:n}\}} - \mathbb{I}_{\{x_{1:n}\}} | o_{1:n}) > 0.$

Maximal state sequences:

 $\begin{aligned} \hat{x}_{1:n} &\in \text{opt}\left(\mathscr{X}_{1:n} | o_{1:n}\right) \\ \Leftrightarrow & (\forall x_{1:n} \in \mathscr{X}_{1:n}) \ x_{1:n} \not\succ \hat{x}_{1:n} \\ \Leftrightarrow & (\forall x_{1:n} \in \mathscr{X}_{1:n}) \ \underline{P}_1(\mathbb{I}_{\{x_{1:n}\}} - \mathbb{I}_{\{\hat{x}_{1:n}\}} | o_{1:n}) \leq 0 \end{aligned}$

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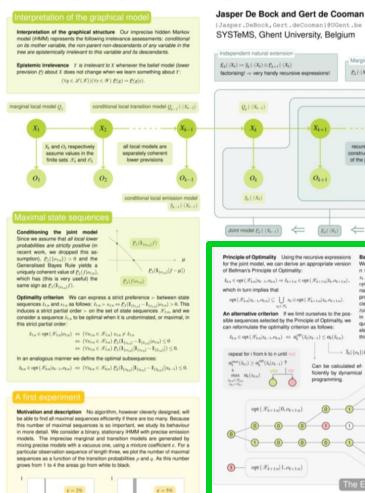
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Maximal state sequences:

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\begin{aligned} \hat{x}_{1:n} &\in \operatorname{opt}\left(\mathscr{X}_{1:n} | o_{1:n}\right) \\ \Leftrightarrow & \left(\forall x_{1:n} \in \mathscr{X}_{1:n}\right) x_{1:n} \not\succ \hat{x}_{1:n} \\ \Leftrightarrow & \left(\forall x_{1:n} \in \mathscr{X}_{1:n}\right) \underline{P}_1(\mathbb{I}_{\{x_{1:n}\}} - \mathbb{I}_{\{\hat{x}_{1:n}\}} | o_{1:n}) \leq 0 \\ \Leftrightarrow & \left(\forall x_{1:n} \in \mathscr{X}_{1:n}\right) \underline{P}_1(\mathbb{I}_{\{o_{1:n}\}}[\mathbb{I}_{\{x_{1:n}\}} - \mathbb{I}_{\{\hat{x}_{1:n}\}}]) \leq 0. \end{aligned}
```

State sequence prediction in imprecise hidden Markov models

Maximal state sequences



Results We see that there are large regions of transition probability space where the number of maximal elements remains fairly small. The plots also

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Joint model $P_{k+1}(\cdot|X_k)$

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Marginal extension

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constructio of the joint

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State sequence prediction in imprecise hidden Markov models

EstiHMM:

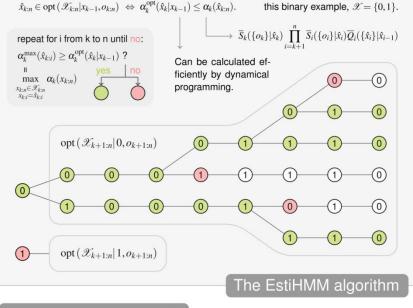
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 $\hat{x}_{k:n} \in \operatorname{opt}\left(\mathscr{X}_{k:n}|x_{k-1}, o_{k:n}\right) \Rightarrow \hat{x}_{k+1:n} \in \operatorname{opt}\left(\mathscr{X}_{k+1:n}|\hat{x}_{k}, o_{k+1:n}\right),$

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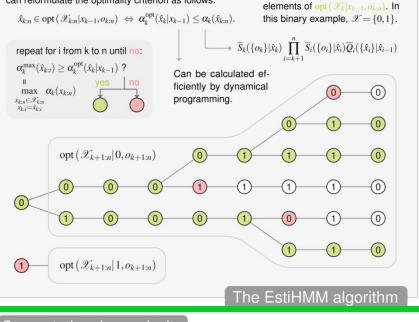
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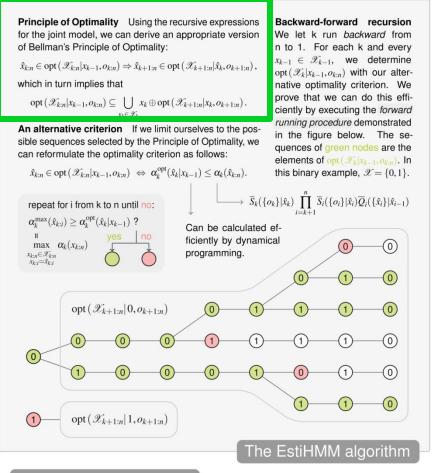
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EstiHMM: an efficient algorithm to determine the maximal sequences

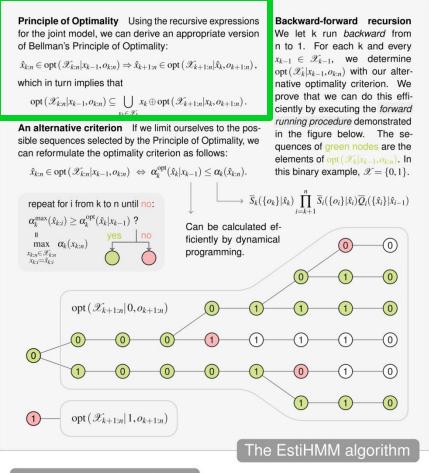


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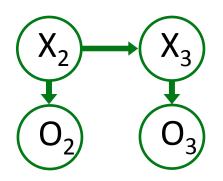


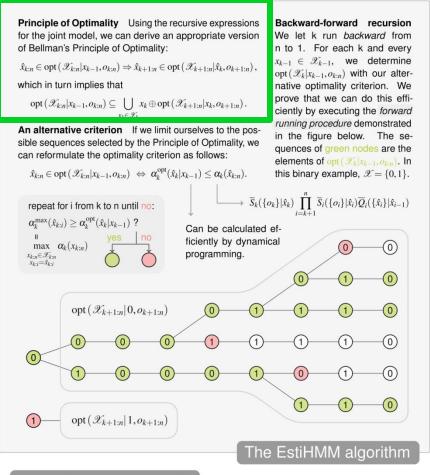
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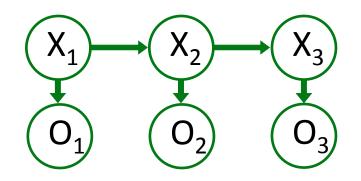


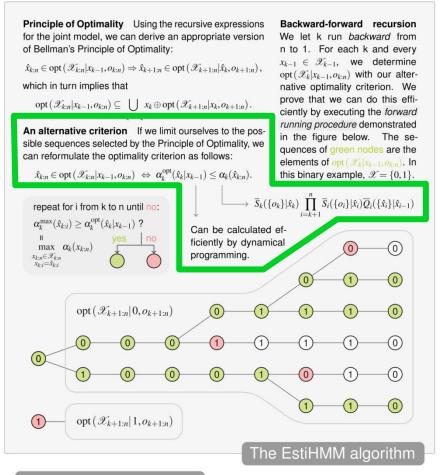
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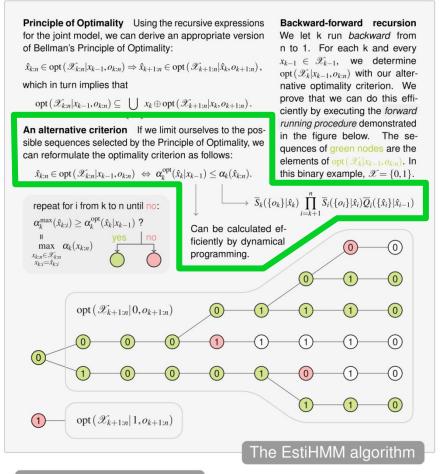
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EstiHMM: an efficient algorithm to determine the maximal sequences

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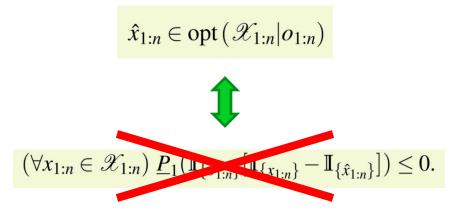


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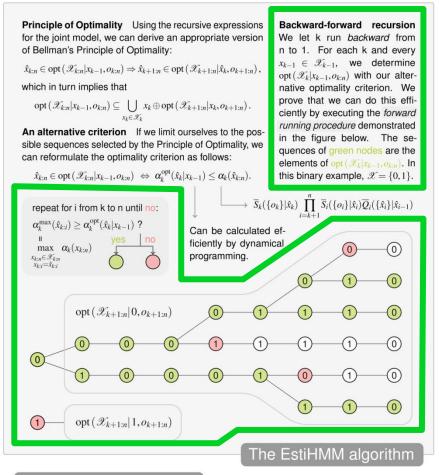
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 $\alpha_k^{\text{opt}}(\hat{x}_k|x_{k-1}) \leq \alpha_k(\hat{x}_{k:n})$



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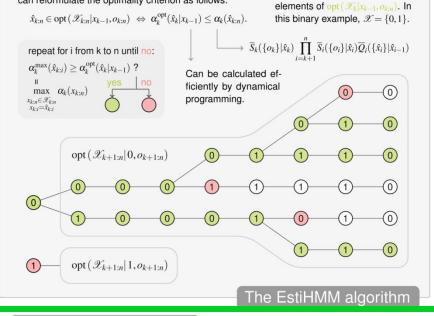
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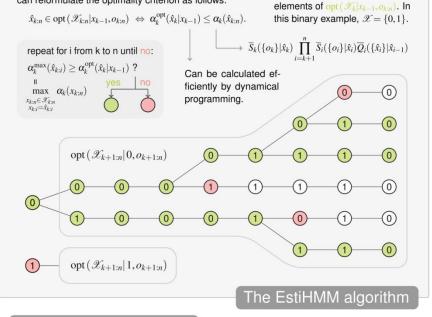
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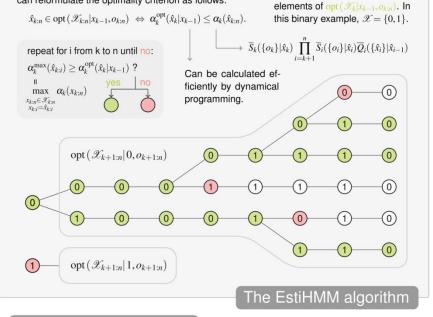
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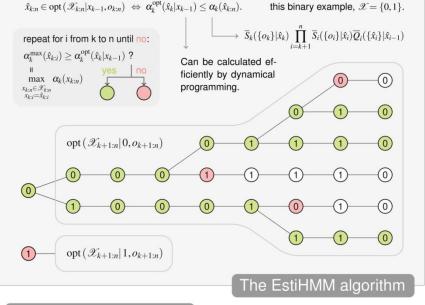
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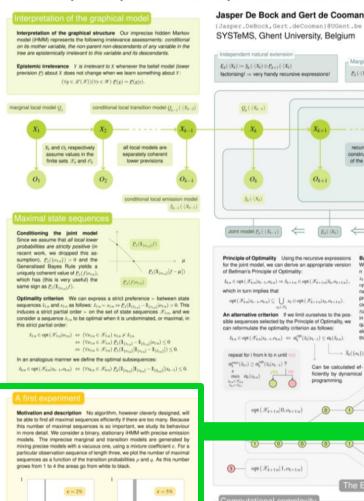
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EstiHMM: an efficient algorithm to determine the maximal sequences

- Principle of optimality
- Deriving an alternative optimality criterion
- A recursive approach

Complexity

- Theoretical analysis
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Results We see that there are large regions of transition probability space where the number of maximal elements remains fairly small. The plots also

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Joint model Part (-(X4)

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