

Imprecise Markov chains

From basic theory to applications I

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UTOPIAÆ

Uncertainty
Treatment and
Optimisation in
Aerospace
Engineering



Handling the unknown at the edge of tomorrow



Imprecise discrete-time Markov chains

Precise probability models

Mass functions and expectations

Assume we are **uncertain** about:

- ▶ the value or a variable X
- ▶ in a set of possible values \mathcal{X} .

This is usually modelled by a **probability mass function** p on \mathcal{X} :

$$p(x) \geq 0 \text{ and } \sum_{x \in \mathcal{X}} p(x) = 1;$$

With p we can associate an **expectation operator** E_p :

$$E_p(f) := \sum_{x \in \mathcal{X}} p(x)f(x) \text{ where } f: \mathcal{X} \rightarrow \mathbb{R}.$$

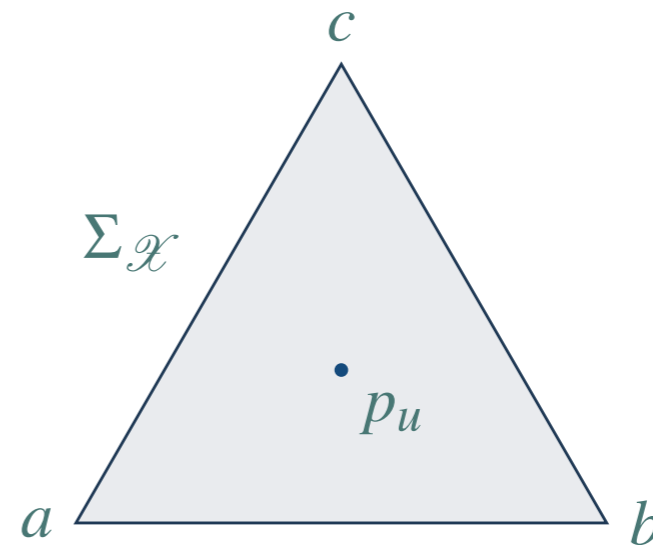
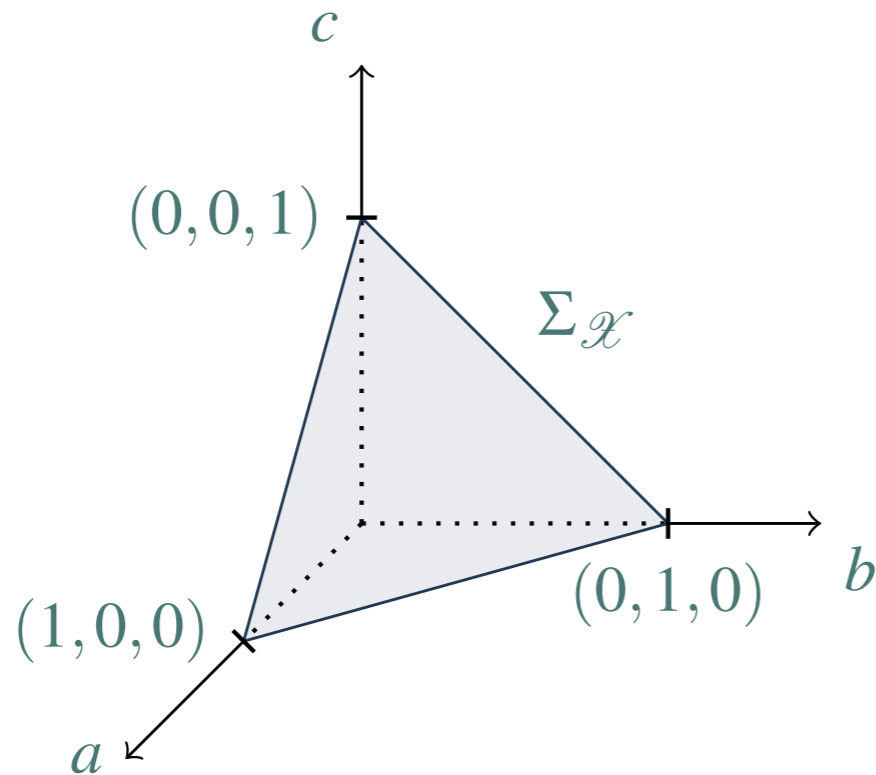
If $A \subseteq \mathcal{X}$ is an **event**, then its **probability** is given by

$$P_p(A) = \sum_{x \in A} p(x) = E_p(I_A).$$

The simplex of all probability mass functions

Consider the **simplex** $\Sigma_{\mathcal{X}}$ of all mass functions on \mathcal{X} :

$$\Sigma_{\mathcal{X}} := \left\{ p \in \mathbb{R}_+^{\mathcal{X}} : \sum_{x \in \mathcal{X}} p(x) = 1 \right\}.$$



Geometrical interpretation of expectation

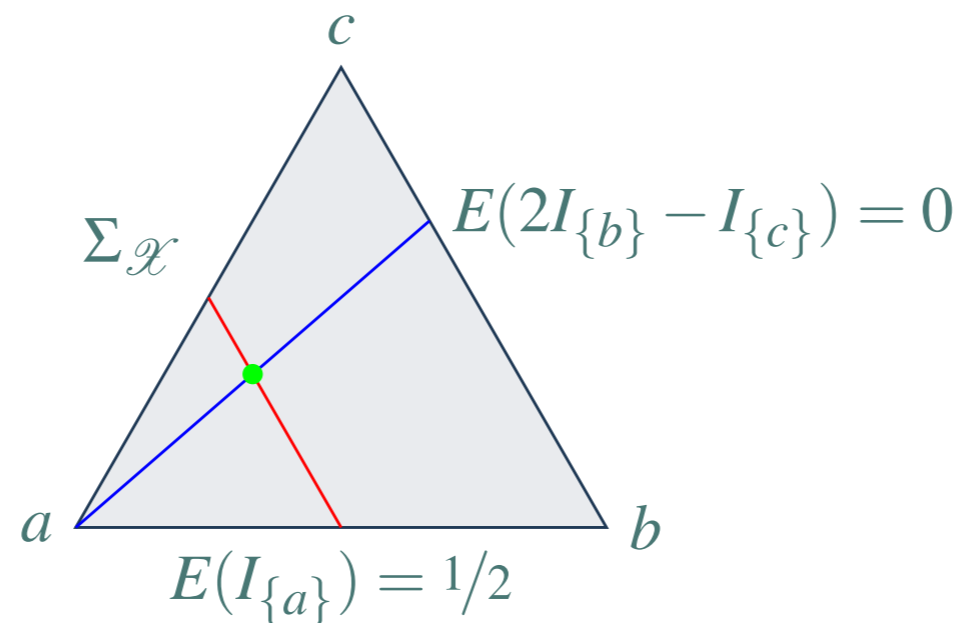
Assessments lead to constraints

Specifying an expectation $E(f)$ for a map $f: \mathcal{X} \rightarrow \mathbb{R}$

$$\sum_{x \in \mathcal{X}} p(x) f(x) = E(f)$$

imposes a **linear constraint** on the possible values for p in $\Sigma_{\mathcal{X}}$.

It corresponds to intersecting the simplex $\Sigma_{\mathcal{X}}$ with a **hyperplane**, whose direction depends on f :



Imprecise probability models

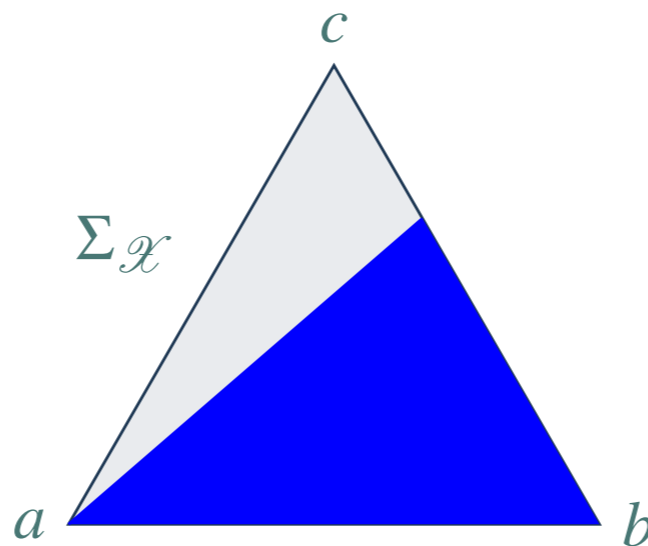
Linear inequality constraints

More flexible assessments

Impose linear inequality constraints on p in $\Sigma_{\mathcal{X}}$:

$$\underline{E}(f) \leq \sum_{x \in \mathcal{X}} p(x)f(x) \quad \text{or} \quad \sum_{x \in \mathcal{X}} p(x)f(x) \leq \bar{E}(f).$$

Corresponds to intersecting $\Sigma_{\mathcal{X}}$ with affine semi-spaces:



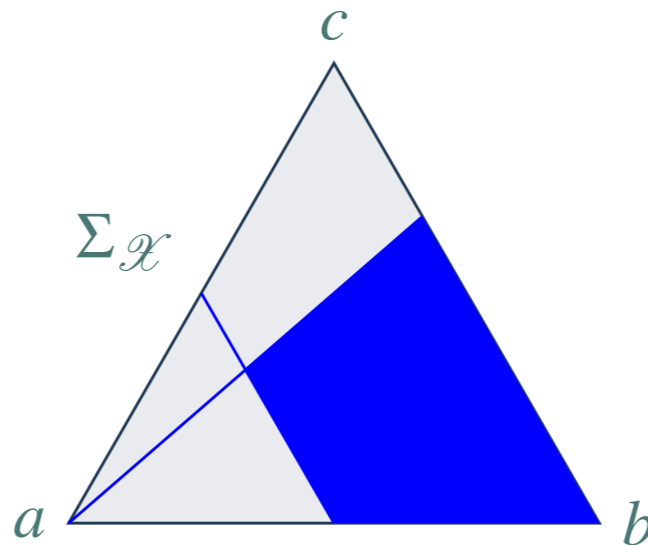
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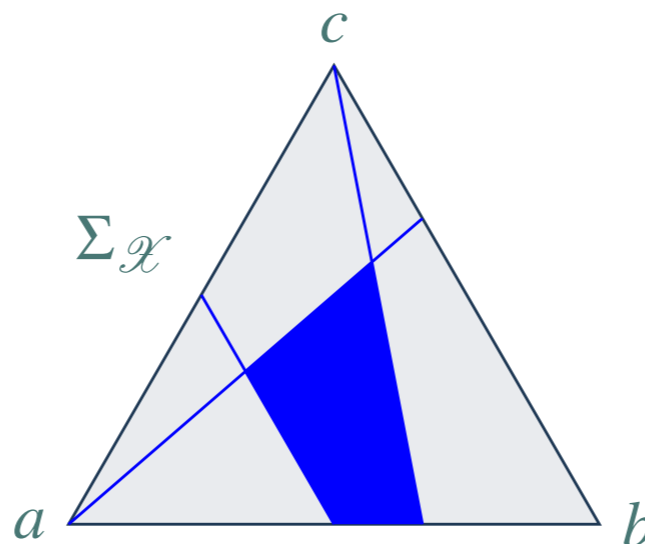
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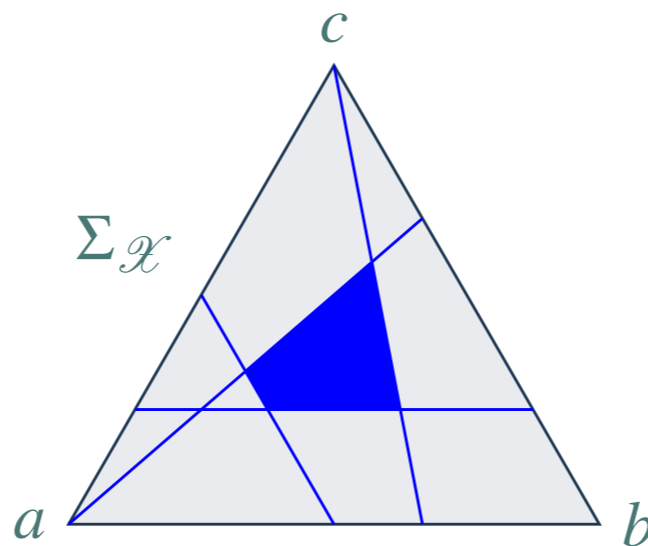
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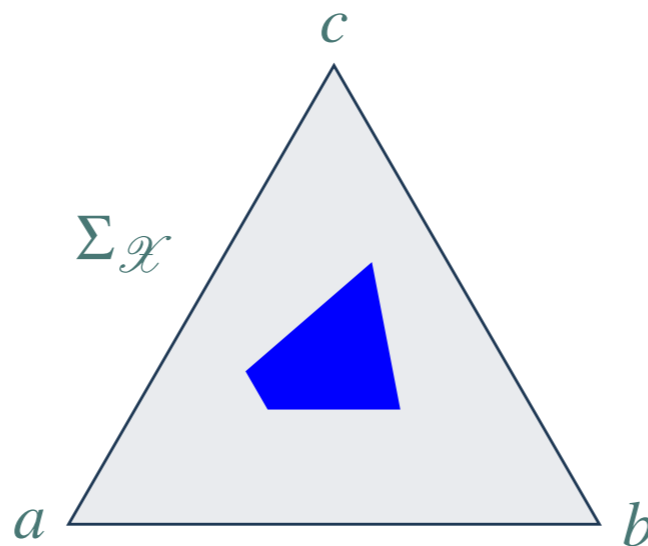
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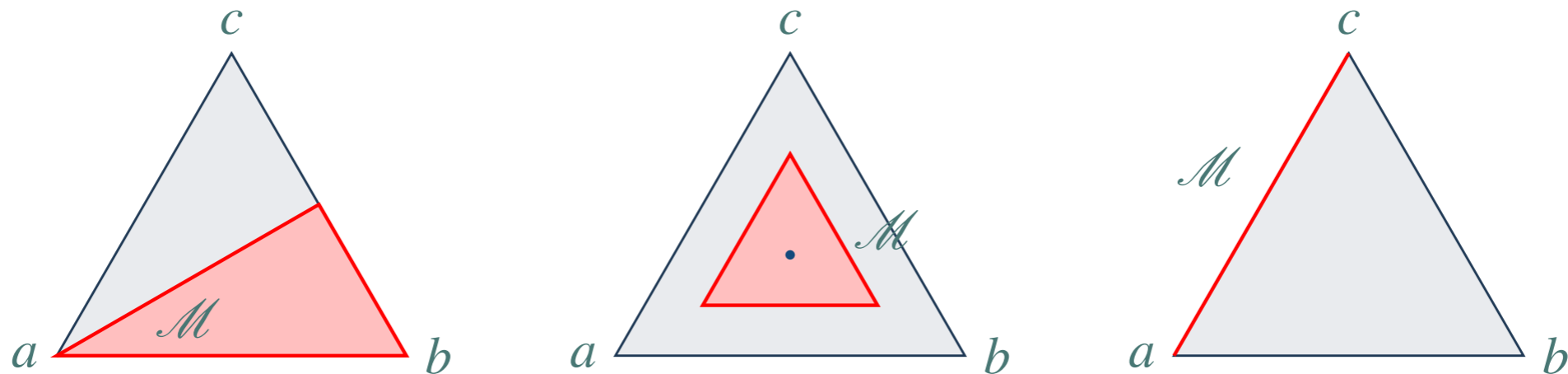


Credal sets

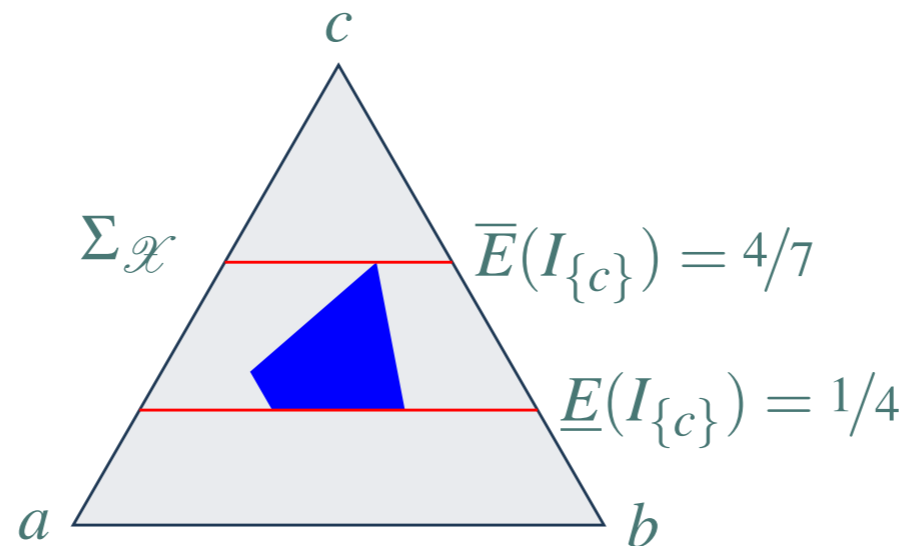
Any such number of assessments leads to a credal set \mathcal{M} .

Definition

A credal set \mathcal{M} is a convex closed subset of $\Sigma_{\mathcal{X}}$.



Lower and upper expectations



Equivalent model

Consider the set $\mathcal{L}(\mathcal{X}) = \mathbb{R}^{\mathcal{X}}$ of all real-valued maps on \mathcal{X} . We define two real functionals on $\mathcal{L}(\mathcal{X})$: for all $f: \mathcal{X} \rightarrow \mathbb{R}$

$$\underline{E}_{\mathcal{M}}(f) = \min \{E_p(f) : p \in \mathcal{M}\} \text{ lower expectation}$$
$$\bar{E}_{\mathcal{M}}(f) = \max \{E_p(f) : p \in \mathcal{M}\} \text{ upper expectation.}$$

Observe that

$$\bar{E}_{\mathcal{M}}(f) = -\underline{E}_{\mathcal{M}}(-f).$$

Basic properties of upper expectations

Definition

We call a real functional \bar{E} on $\mathcal{L}(\mathcal{X})$ an **upper expectation** if it satisfies the following properties:

for all f and g in $\mathcal{L}(\mathcal{X})$ and all real $\lambda \geq 0$:

1. $\bar{E}(f) \leq \max f$ [**boundedness**];
2. $\bar{E}(f + g) \leq \bar{E}(f) + \bar{E}(g)$ [**sub-additivity**];
3. $\bar{E}(\lambda f) = \lambda \bar{E}(f)$ [**non-negative homogeneity**].

Theorem

A real functional \bar{E} is an upper expectation if and only if it is the upper envelope of some credal set \mathcal{M} .

Proof.

Use $\mathcal{M} = \{p \in \Sigma_{\mathcal{X}} : (\forall f \in \mathcal{L}(\mathcal{X})) (E_p(f) \leq \bar{E}(f))\}$. □

Discrete-time uncertain processes

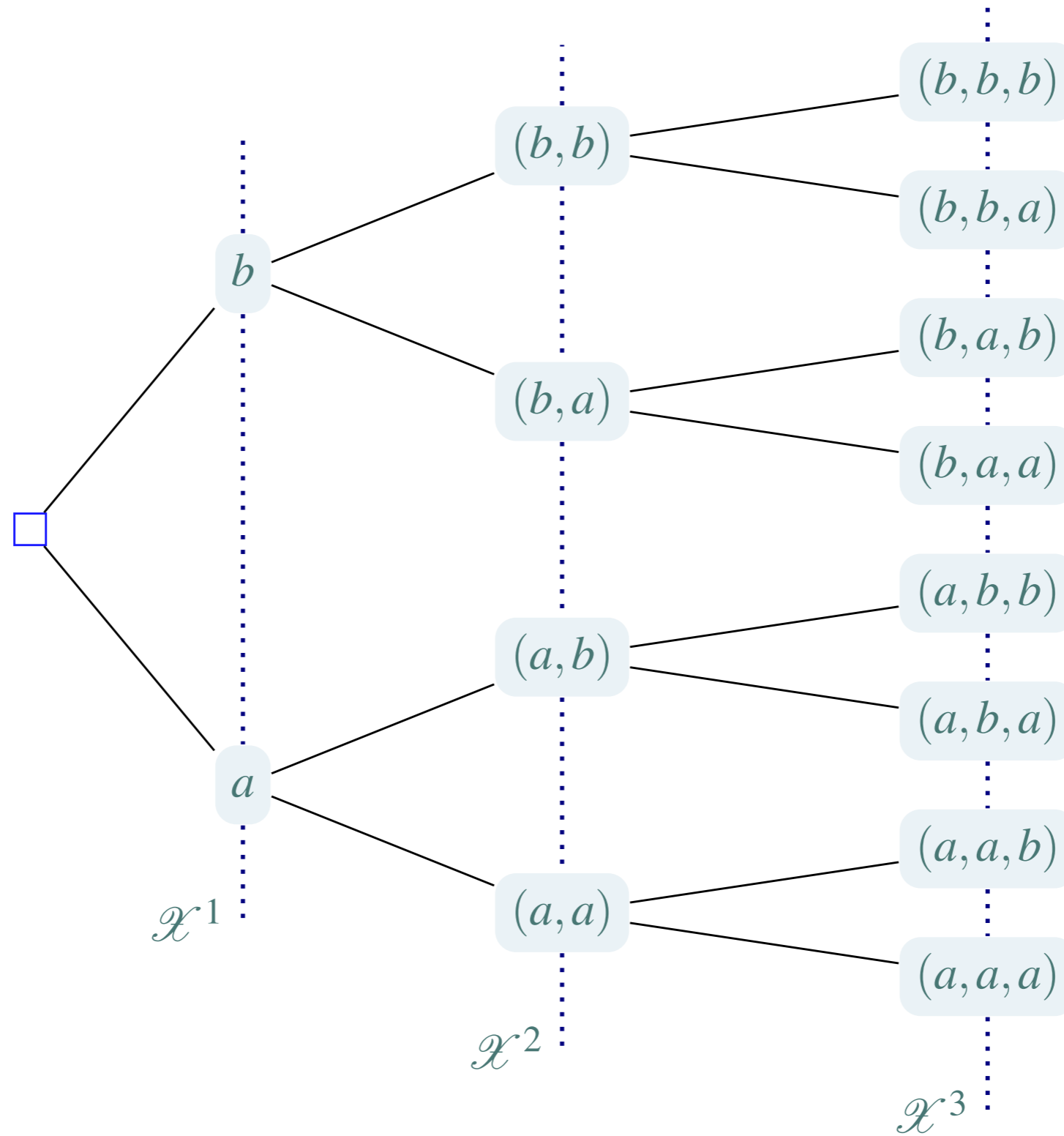
Precise probability trees

We consider an **uncertain process** with variables $X_1, X_2, \dots, X_n, \dots$ assuming values in a finite set of **states** \mathcal{X} .

This leads to a **standard event tree** with nodes

$$s = (x_1, x_2, \dots, x_n), \quad x_k \in \mathcal{X}, \quad n \geq 0.$$

Precise probability trees



Precise probability trees

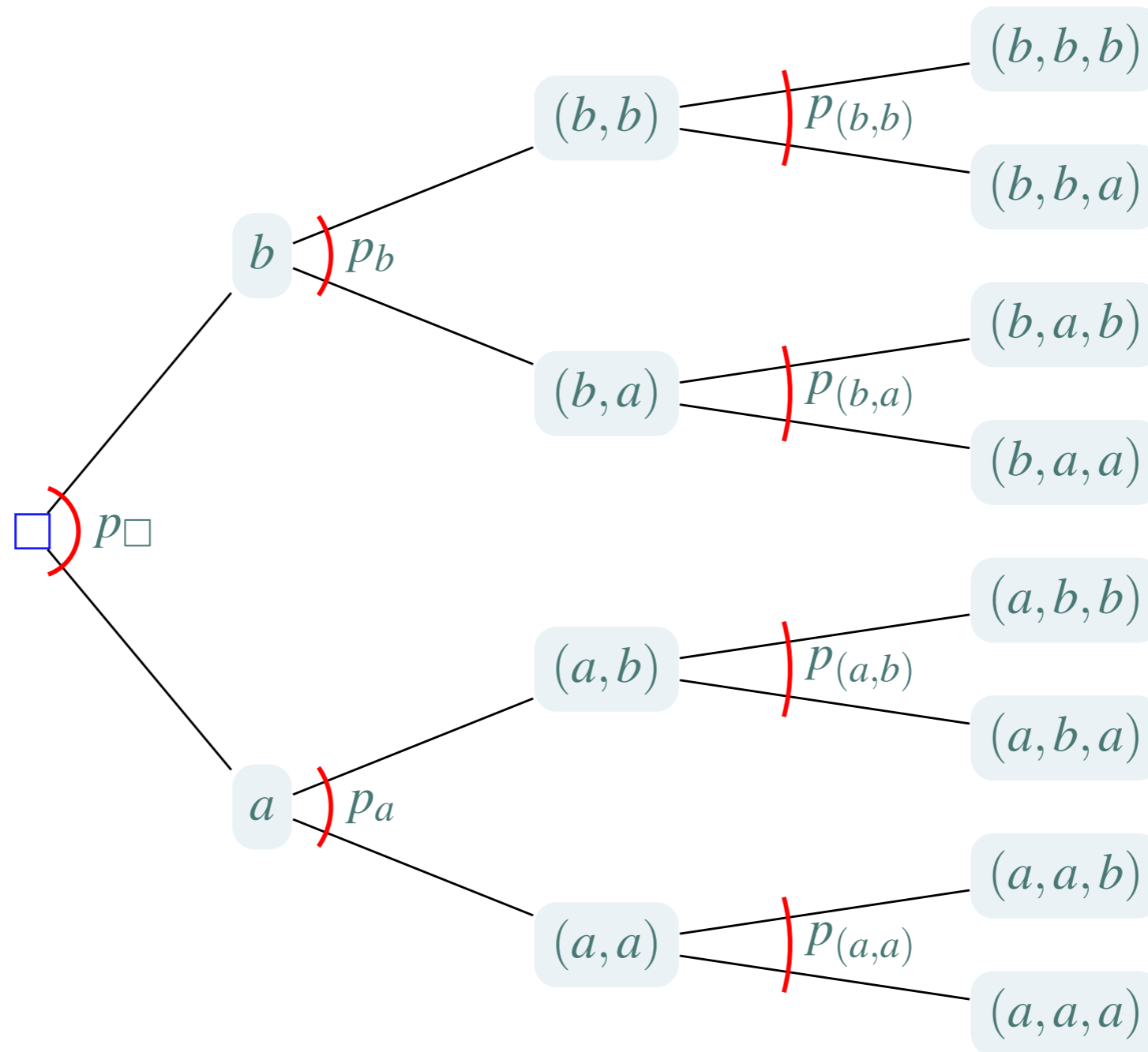
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$$s = (x_1, x_2, \dots, x_n), \quad x_k \in \mathcal{X}, \quad n \geq 0.$$

The standard event tree becomes a **probability tree** by attaching to each node s a local **probability mass function** p_s on \mathcal{X} with associated **expectation operator** E_s .

Precise probability trees



Calculating global expectations from local ones

Consider a function $g: \mathcal{X}^n \rightarrow \mathbb{R}$ of the first n variables:

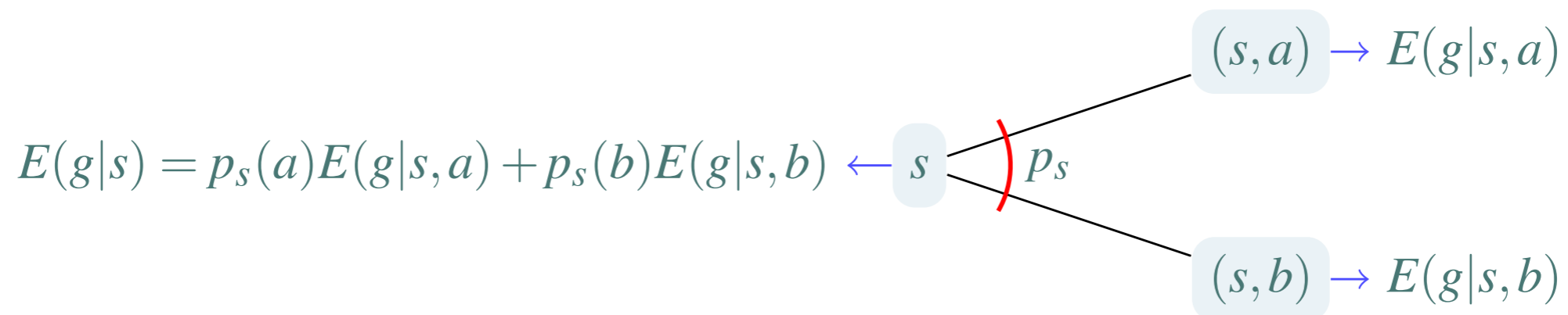
$$g = g(X_1, X_2, \dots, X_n)$$

We want to calculate its expectation $E(g|s)$ in $s = (x_1, \dots, x_k)$.

Theorem (Law of Iterated Expectation)

Suppose we know $E(g|s, x)$ for all $x \in \mathcal{X}$, then we can calculate $E(g|s)$ by *backwards recursion* using the local model p_s :

$$E(g|s) = \underbrace{E_s}_{\text{local}}(E(g|s, \cdot)) = \sum_{x \in \mathcal{X}} p_s(x) E(g|s, x).$$



Calculating global expectations from local ones

All expectations $E(g|x_1, \dots, x_k)$ in the tree can be calculated from the local models as follows:

1. start in the final cut \mathcal{X}^n and let:

$$E(g|x_1, x_2, \dots, x_n) = g(x_1, x_2, \dots, x_n);$$

2. do backwards recursion using the Law of Iterated Expectation:

$$E(g|x_1, \dots, x_k) = \underbrace{E_{(x_1, \dots, x_k)}}_{\text{local}}(E(g|x_1, \dots, x_k, \cdot))$$

3. go on until you get to the root node \square , where:

$$E(g|\square) = E(g).$$

Christiaan Huygens (1656–1657)

Si vincat qui proximus duobus punctis alibi praestiteri
 calculus ita se habebit ut x portio debita lusori B
 et quod depositum sit, quod vocetur n.

$$\begin{array}{l}
 \text{od} \\
 x \propto \left\{ \begin{array}{l} y \text{ debita } 1 \text{ so} \\ \xi \text{ deb. } - \text{od} \end{array} \right. \left\{ \begin{array}{l} y \text{ --- } \frac{n}{c} \\ \xi \text{ --- } x \\ \xi \text{ --- } \frac{x^2}{2c} \end{array} \right. \begin{array}{l} \\ \\ 2x \frac{\partial x}{\partial c} \\ y x \frac{\partial n + cx}{\partial c} \end{array}
 \end{array}$$

$$\begin{array}{l}
 \text{od} \\
 x \propto \left\{ \begin{array}{l} d \text{ --- } \frac{\partial n + cx}{\partial c} \\ c \text{ --- } \frac{\partial x}{\partial c} \end{array} \right. \text{ Ergo } x \propto \frac{\partial \partial n + 2 \partial cx}{cc + 2 \partial c + \partial d}
 \end{array}$$

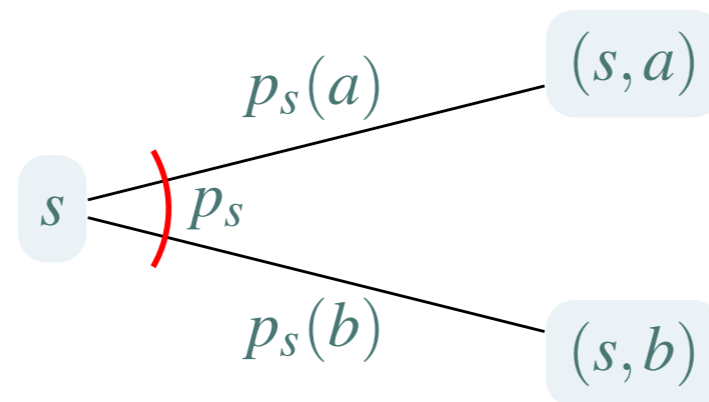
Cum B habeat $\frac{\partial n}{cc + \partial d}$, habebit
 A $\frac{ccn}{cc + \partial d}$ quia simul addita pars debet
 facta n. Ergo pars B ad partem A ut ∂d ad cc .

$$\frac{ccx + 2 \partial cx + \partial \partial x}{x} \propto \frac{\partial \partial n}{cc + \partial d} \text{ portio lusori B.}$$

Sets of mass functions

Major restrictive assumption

Until now, we have assumed that we have **sufficient information** in order to specify, in each node s , a probability mass function p_s on the set \mathcal{X} of possible values for the next state.



More general uncertainty models

We consider **credal sets** as more general uncertainty models: **closed convex subsets** of $\Sigma_{\mathcal{X}}$.

Imprecise probability trees

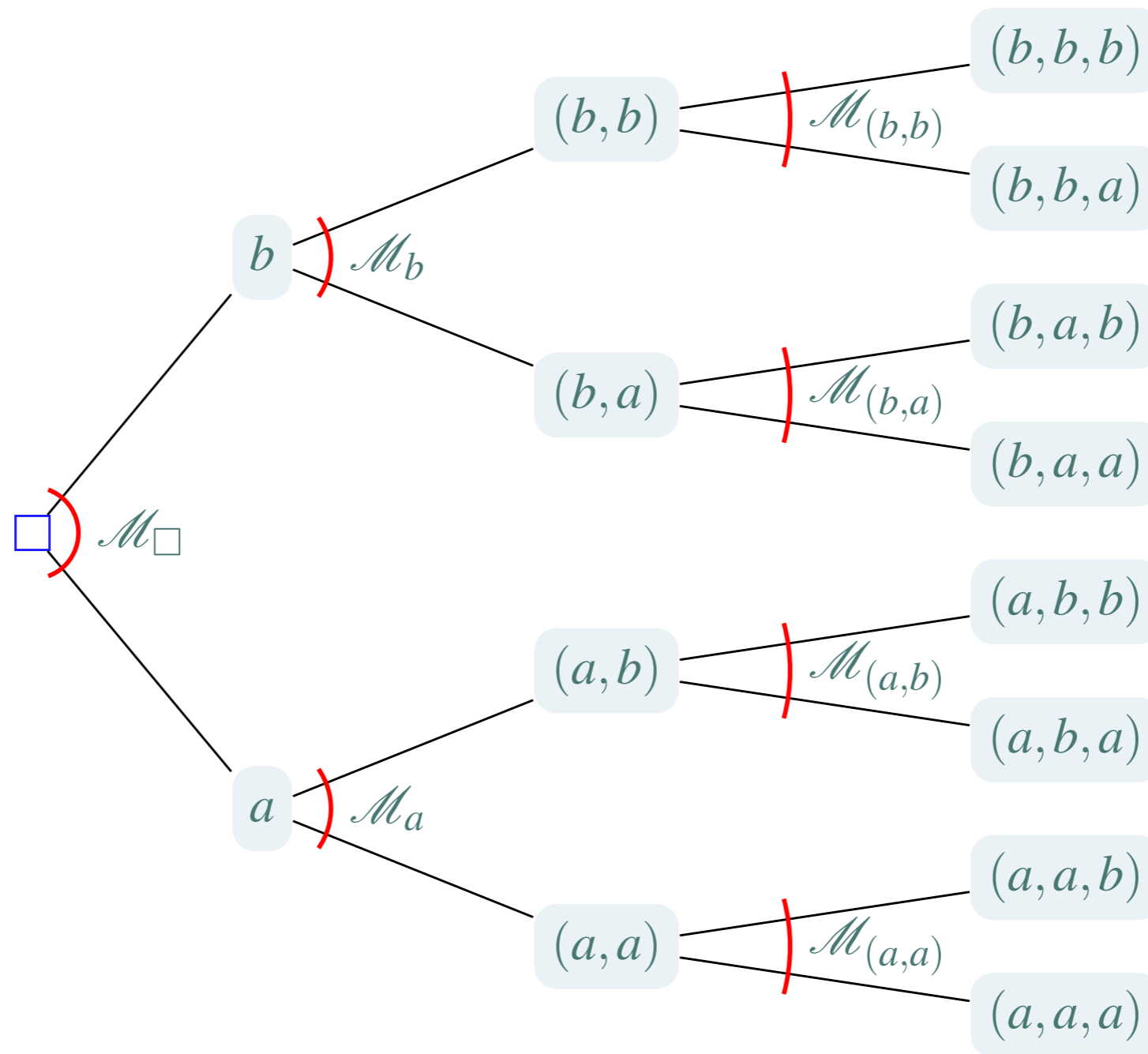
Definition and interpretation

Definition

An **imprecise probability tree** is a probability tree where in each node s the local uncertainty model is an **imprecise probability model** \mathcal{M}_s , or equivalently, its associated **upper expectation** \bar{E}_s :

$$\bar{E}_s(f) = \max \{E_p(f) : p \in \mathcal{M}_s\} \text{ for all real maps } f \text{ on } \mathcal{X}.$$

Definition and interpretation



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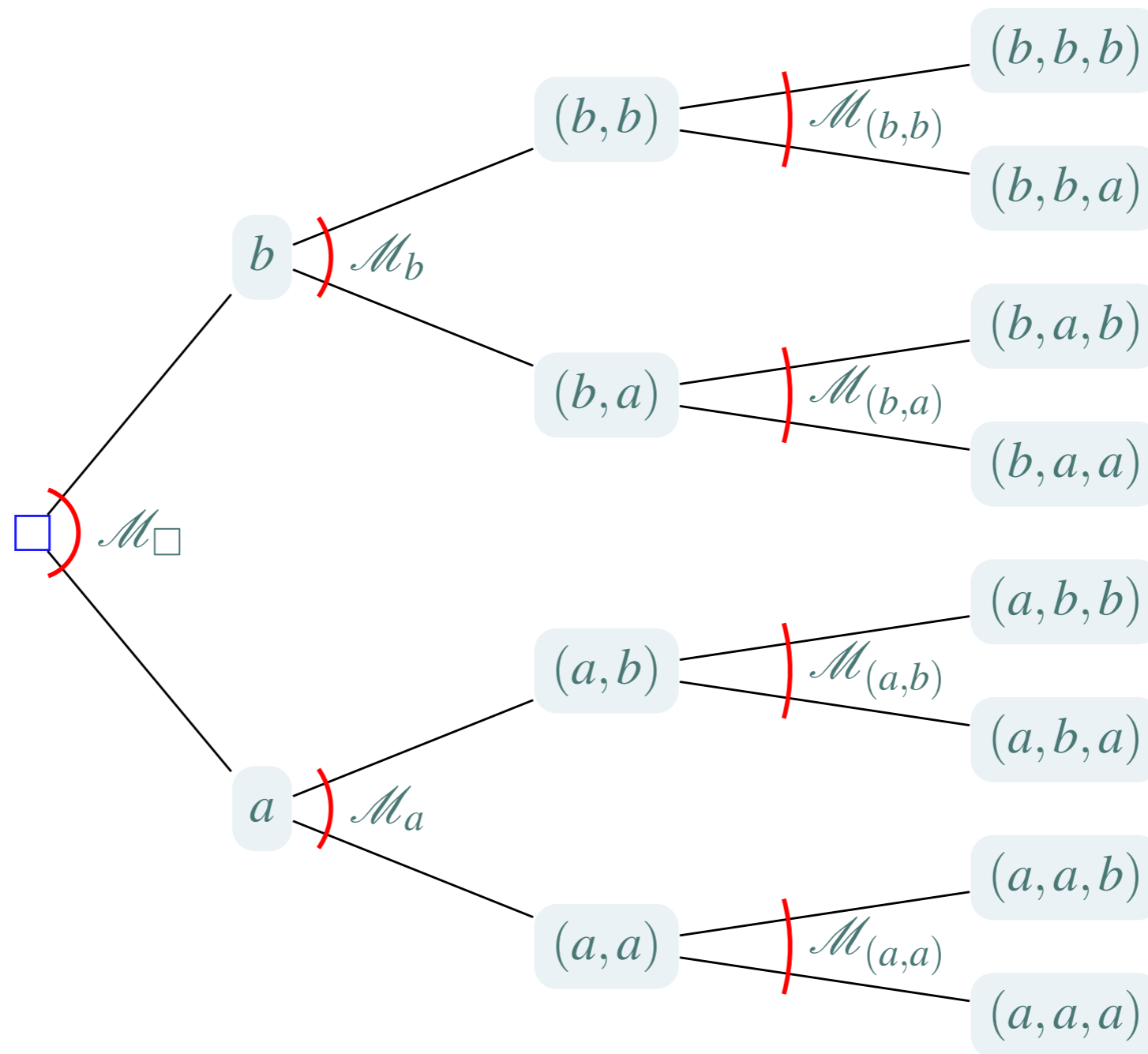
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An imprecise probability tree can be seen as an infinity of **compatible** precise probability trees: choose in each node s a probability mass function p_s from the set \mathcal{M}_s .

Definition and interpretation



Associated lower and upper expectations

For each real map $g = g(X_1, \dots, X_n)$, each node $s = (x_1, \dots, x_k)$, and each such compatible precise probability tree, we can calculate the expectation

$$E(g|x_1, \dots, x_k)$$

using the backwards recursion method described before.

By varying over each compatible probability tree, we get a closed real interval:

$$[\underline{E}(g|x_1, \dots, x_k), \bar{E}(g|x_1, \dots, x_k)]$$

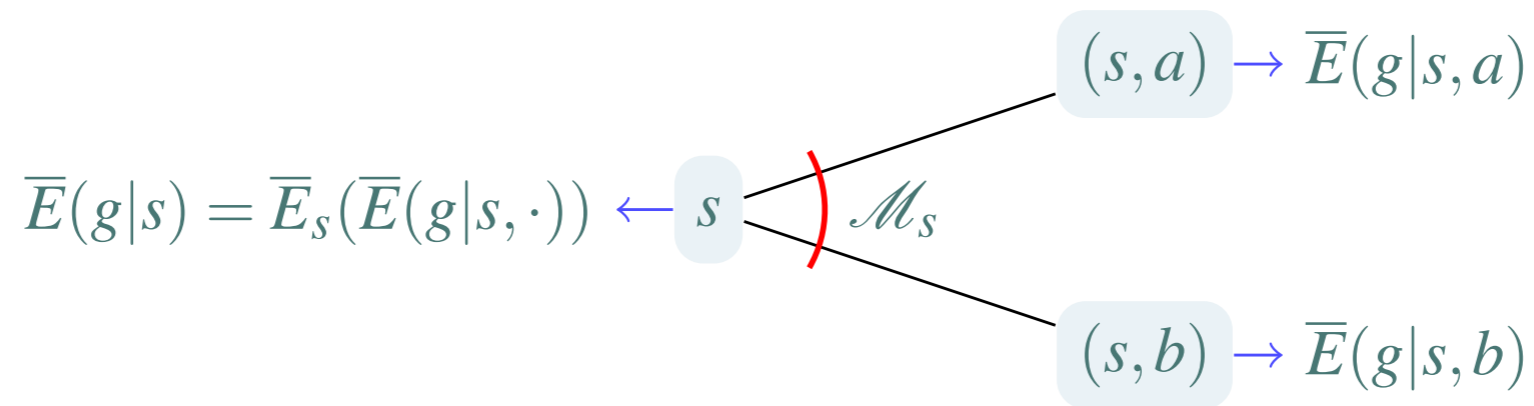
We want a better, more efficient method to calculate these lower and upper expectations $\underline{E}(g|x_1, \dots, x_k)$ and $\bar{E}(g|x_1, \dots, x_k)$.

The Law of Iterated Expectation

Theorem (Law of Iterated Expectation)

Suppose we know $\bar{E}(g|s, x)$ for all $x \in \mathcal{X}$, then we can calculate $\bar{E}(g|s)$ by *backwards recursion* using the local model \bar{E}_s :

$$\bar{E}(g|s) = \underbrace{\bar{E}_s}_{\text{local}}(\bar{E}(g|s, \cdot)) = \max_{p_s \in \mathcal{M}_s} \sum_{x \in \mathcal{X}} p_s(x) \bar{E}(g|s, x).$$



The complexity of calculating the $\bar{E}(g|s)$, as a function of n , is therefore essentially the same as in the precise case!

Imprecise Markov chains

Precise Markov chains: definition

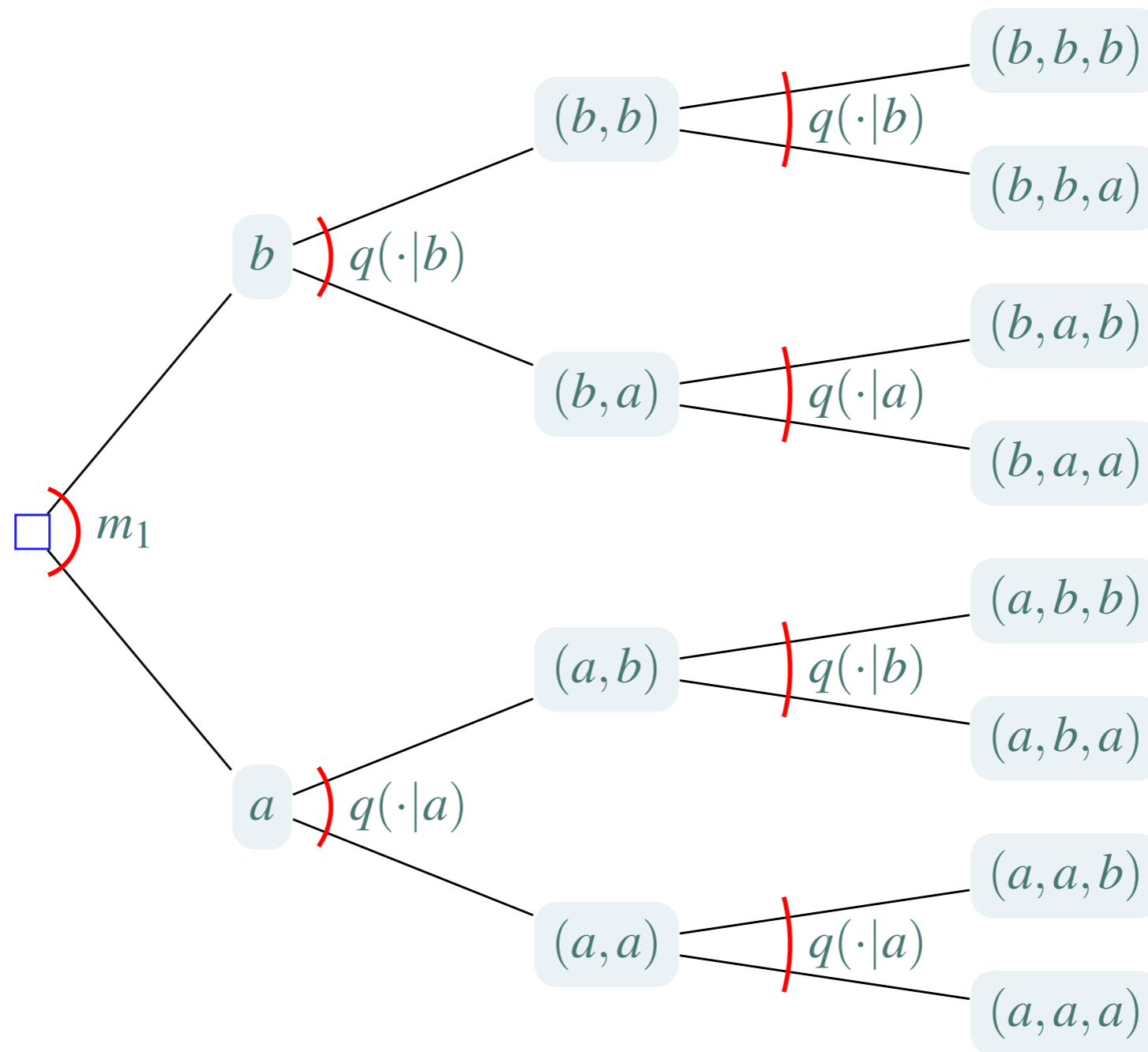
Definition

The uncertain process is a stationary **precise Markov chain** when all \mathcal{M}_s are singletons (precise), and

1. $\mathcal{M}_\square = \{m_1\}$,
2. the **Markov Condition** is satisfied:

$$\mathcal{M}_{(x_1, \dots, x_n)} = \{q(\cdot | x_n)\}.$$

Precise Markov chains: definition



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$$\mathcal{M}_{(x_1, \dots, x_n)} = \{q(\cdot | x_n)\}.$$

For each $x \in \mathcal{X}$, the transition mass function $q(\cdot | x)$ corresponds to an expectation operator:

$$E(f|x) = \sum_{z \in \mathcal{X}} q(z|x) f(z).$$

Precise Markov chains: transition operators

Definition

Consider the linear transformation T of $\mathcal{L}(\mathcal{X})$, called **transition operator**:

$$T: \mathcal{L}(\mathcal{X}) \rightarrow \mathcal{L}(\mathcal{X}): f \mapsto Tf$$

where Tf is the real map given by, for any $x \in \mathcal{X}$:

$$Tf(x) := E(f|x) = \sum_{z \in \mathcal{X}} q(z|x)f(z)$$

T is the dual of the linear transformation with **Markov matrix** M , with elements $M_{xy} := q(y|x)$.

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Then the **Law of Iterated Expectation** yields:

$$E_n(f) = E_1(T^{n-1}f), \text{ and dually, } m_n = M^{n-1}m_1.$$

Complexity is linear in the number of time steps n .

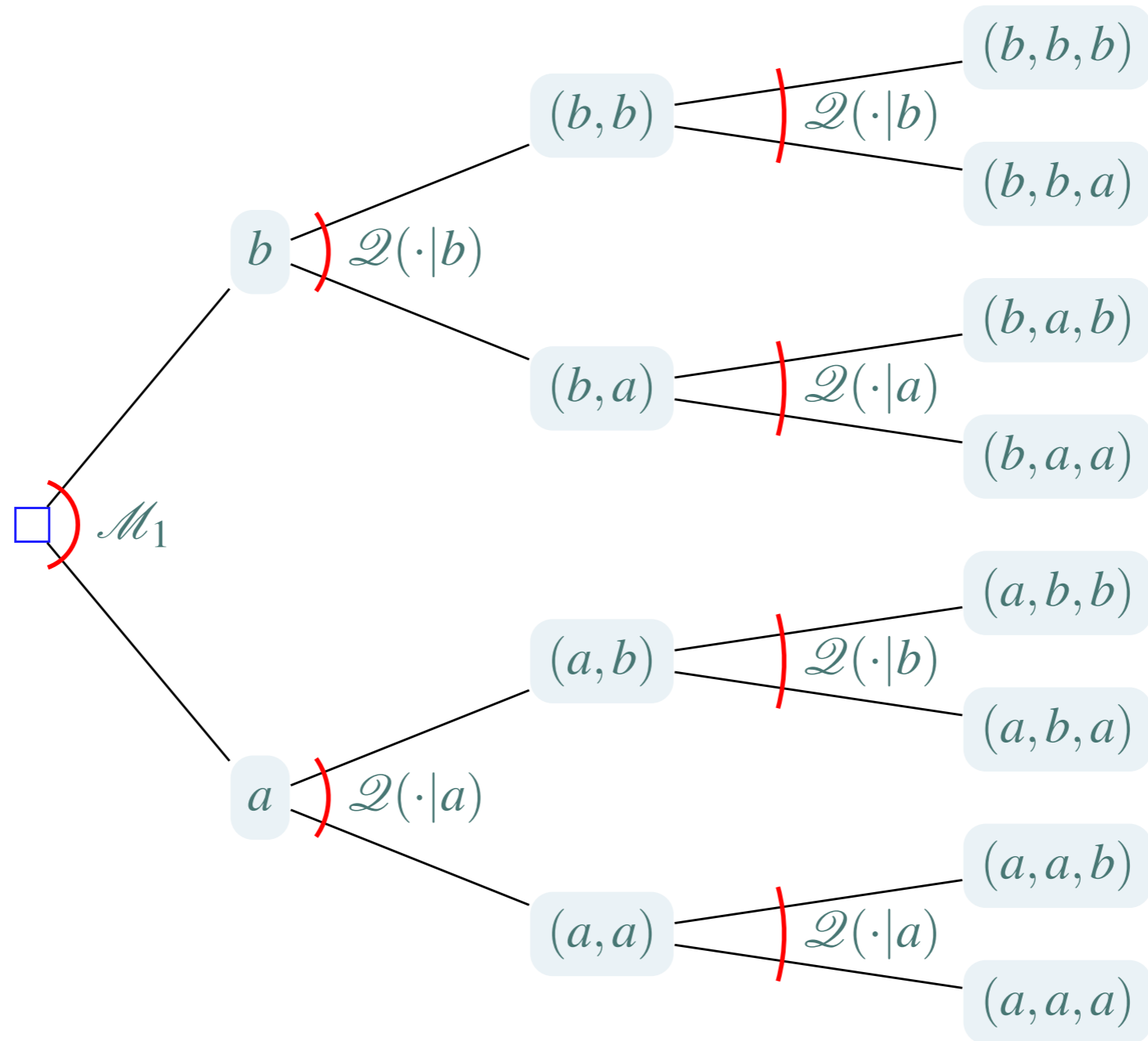
Imprecise Markov chains: definition

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The uncertain process is a stationary **imprecise Markov chain** when the **Markov Condition** is satisfied:

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Imprecise Markov chains: definition



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$$\mathcal{M}_{(x_1, \dots, x_n)} = \mathcal{Q}(\cdot | x_n).$$

An imprecise Markov chain can be seen as an infinity of probability trees.

For each $x \in \mathcal{X}$, the local transition model $\mathcal{Q}(\cdot | x)$ corresponds to **lower** and **upper expectation operators**:

$$\underline{E}(f|x) = \min \{E_p(f) : p \in \mathcal{Q}(\cdot | x)\}$$
$$\bar{E}(f|x) = \max \{E_p(f) : p \in \mathcal{Q}(\cdot | x)\}.$$

Imprecise Markov chains: transition operators

Definition

Consider the **non-linear** transformations $\underline{\mathbb{T}}$ and $\overline{\mathbb{T}}$ of $\mathcal{L}(\mathcal{X})$, called **lower** and **upper transition operators**:

$$\underline{\mathbb{T}}: \mathcal{L}(\mathcal{X}) \rightarrow \mathcal{L}(\mathcal{X}): f \mapsto \underline{\mathbb{T}}f$$

$$\overline{\mathbb{T}}: \mathcal{L}(\mathcal{X}) \rightarrow \mathcal{L}(\mathcal{X}): f \mapsto \overline{\mathbb{T}}f$$

where the real maps $\underline{\mathbb{T}}f$ and $\overline{\mathbb{T}}f$ are given by:

$$\begin{aligned} \underline{\mathbb{T}}f(x) &:= \underline{E}(f|x) = \min \{E_p(f) : p \in \mathcal{Q}(\cdot|x)\} \\ \overline{\mathbb{T}}f(x) &:= \overline{E}(f|x) = \max \{E_p(f) : p \in \mathcal{Q}(\cdot|x)\} \end{aligned}$$

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Then the **Law of Iterated Expectation** yields:

$$\underline{E}_n(f) = \underline{E}_1(\underline{\mathbb{T}}^{n-1}f) \text{ and } \overline{E}_n(f) = \overline{E}_1(\overline{\mathbb{T}}^{n-1}f).$$

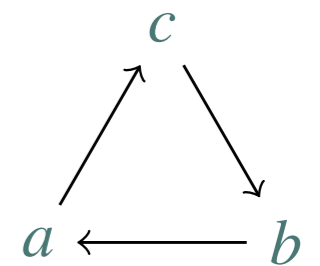
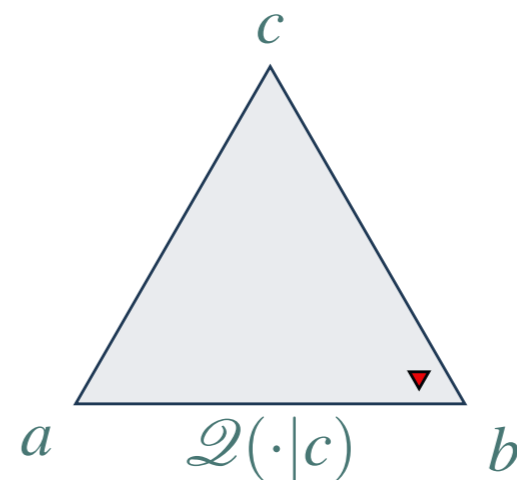
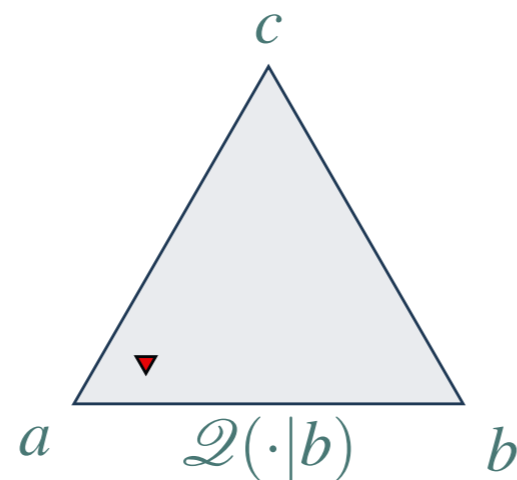
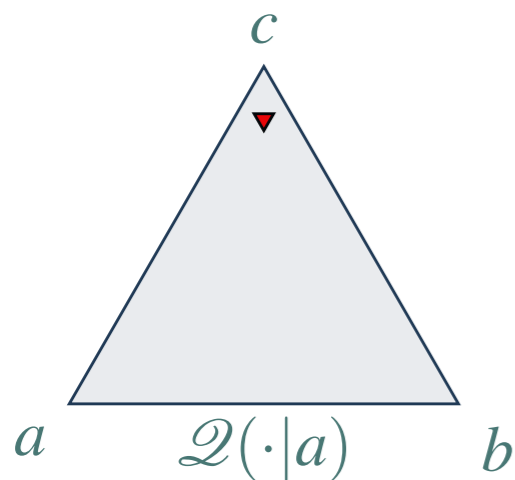
Complexity is still linear in the number of time steps n .

Lower and upper mass functions

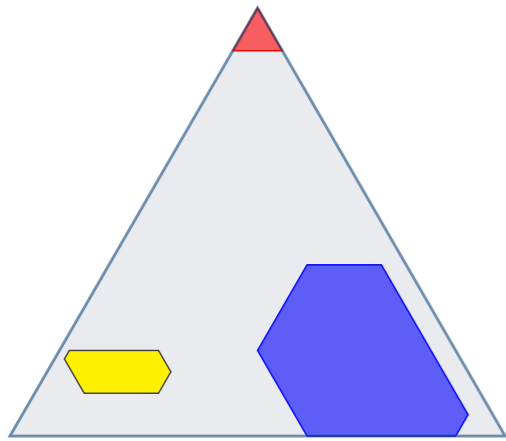
Another example with $\mathcal{X} = \{a, b, c\}$

$$[\underline{\mathbf{T}}I_{\{a\}} \quad \underline{\mathbf{T}}I_{\{b\}} \quad \underline{\mathbf{T}}I_{\{c\}}] = \begin{bmatrix} \underline{q}(a|a) & \underline{q}(b|a) & \underline{q}(c|a) \\ \underline{q}(a|b) & \underline{q}(b|b) & \underline{q}(c|b) \\ \underline{q}(a|c) & \underline{q}(b|c) & \underline{q}(c|c) \end{bmatrix} = 1/200 \begin{bmatrix} 9 & 9 & 162 \\ 144 & 18 & 18 \\ 9 & 162 & 9 \end{bmatrix}$$

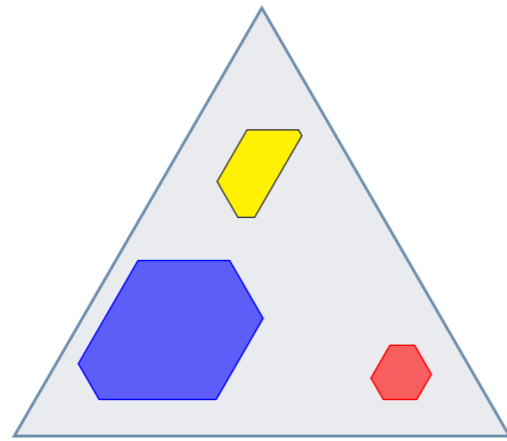
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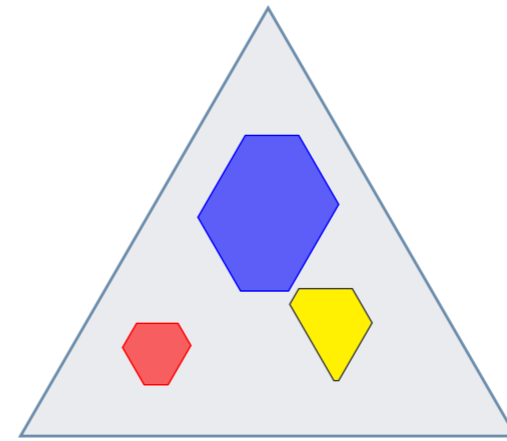
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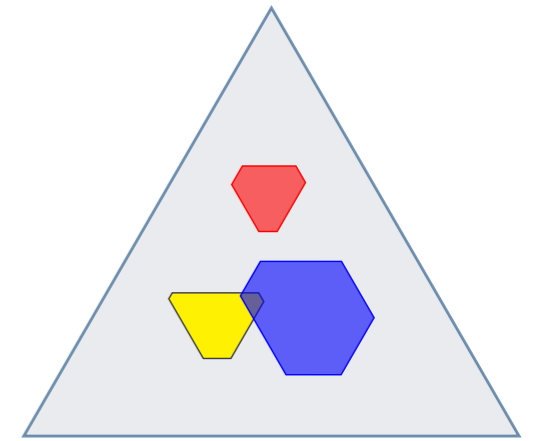
$n = 1$



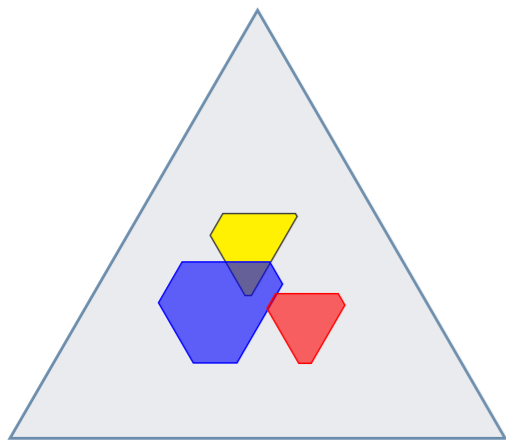
$n = 2$



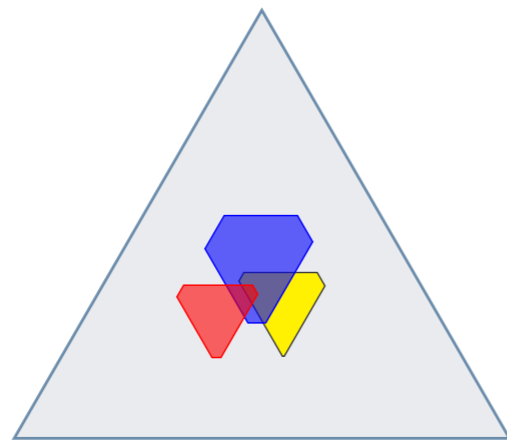
$n = 3$



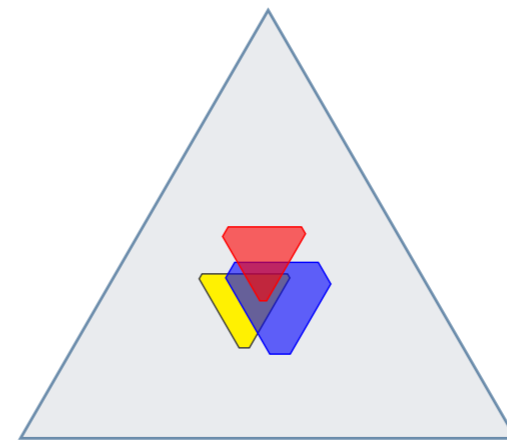
$n = 4$



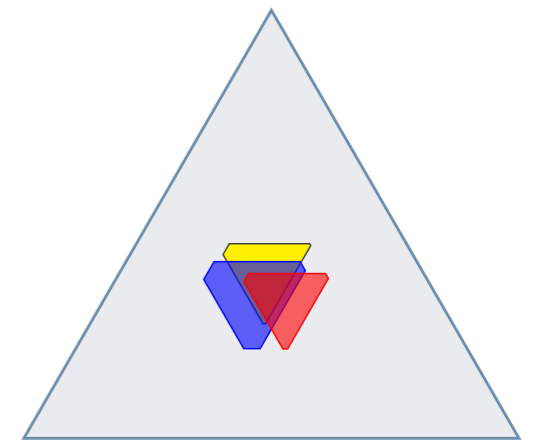
$n = 5$



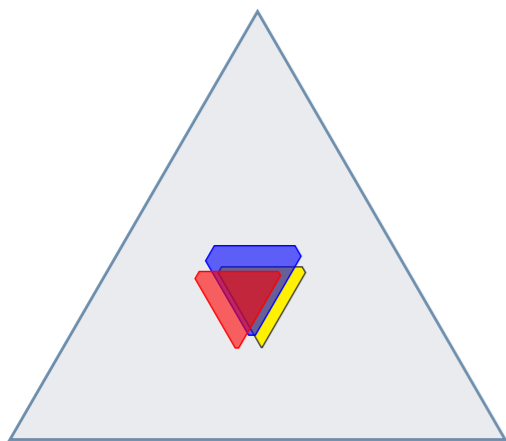
$n = 6$



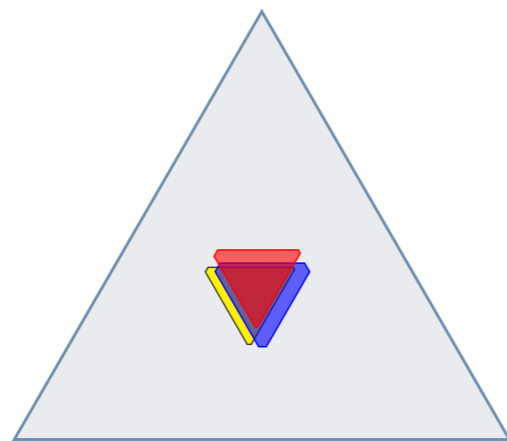
$n = 7$



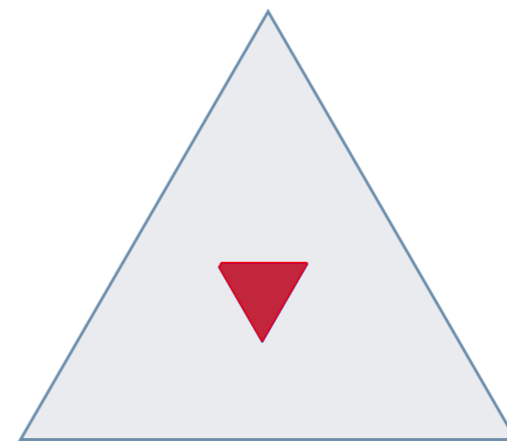
$n = 8$



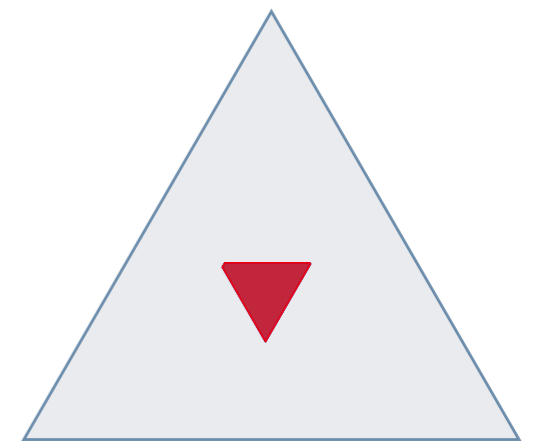
$n = 9$



$n = 10$



$n = 22$



$n = 1000$

**A non-linear
Perron–Frobenius Theorem**

Generalising the linear case

Theorem (De Cooman, Hermans and Quaeghebeur, 2008)

Consider a stationary imprecise Markov chain with finite state set \mathcal{X} and an upper transition operator \bar{T} . Suppose that \bar{T} is *regular*, meaning that there is some $n > 0$ such that $\min \bar{T}^n I_{\{x\}} > 0$ for all $x \in \mathcal{X}$. Then for every initial upper expectation \bar{E}_1 , the upper expectation $\bar{E}_n = \bar{E}_1 \circ \bar{T}^{n-1}$ for the state at time n converges point-wise to the same upper expectation \bar{E}_∞ :

$$\lim_{n \rightarrow \infty} \bar{E}_n(h) = \lim_{n \rightarrow \infty} \bar{E}_1(\bar{T}^{n-1} h) := \bar{E}_\infty(h)$$

for all h in $\mathcal{L}(\mathcal{X})$. Moreover, the corresponding limit upper expectation \bar{E}_∞ is the *only* \bar{T} -invariant upper expectation on $\mathcal{L}(\mathcal{X})$, meaning that $\bar{E}_\infty = \bar{E}_\infty \circ \bar{T}$.

Ergodicity

And in that case also

Theorem (De Cooman, De Bock and Lopatatzidis, 2016)


Consider a stationary imprecise Markov chain with finite state set \mathcal{X} and an upper transition operator \bar{T} . Suppose that \bar{T} is Perron–Frobenius-like with stationary upper expectation \bar{E}_∞ . Then

$$\underline{E}_\infty(h) \leq \liminf_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n h(X_k) \leq \limsup_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n h(X_k) \leq \bar{E}_\infty(h)$$


almost surely for all h in $\mathcal{L}(\mathcal{X})$.

Literature

Literature

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