

Markov chains

An introduction

Consider a generic continuous-time stochastic process $(X_t)_{t \in \mathbb{R}_{\geq 0}}$, where for all $t \in \mathbb{R}_{\geq 0}$ the state X_t is a random variable that takes values x in the finite state space \mathcal{X} . We provide \mathcal{X} with some ordering, such that any real-valued function f on \mathcal{X} can be identified with a row vector. We furthermore let $\mathcal{L}(\mathcal{X})$ denote the set of all real-valued functions on \mathcal{X} . Then any linear operator $T: \mathcal{L}(\mathcal{X}) \rightarrow \mathcal{L}(\mathcal{X})$ can be identified with a matrix.

Precise Markov chains

The stochastic process $(X_t)_{t \in \mathbb{R}_{\geq 0}}$ is a *precise (continuous-time) Markov chain* (pMC) if it satisfies the *Markov property*: where $n \geq 0$ is an integer and $\{t_1, \dots, t_n, s, t\}$ is a strictly increasing sequence of non-negative time points. The *transition matrix* T_s^t thus defined satisfies

$$\begin{aligned} [T_s^t f](x_s) &= \mathbb{E}(f(X_t) | X_s = x_s) \\ &= \mathbb{E}(f(X_t) | X_{t_1} = x_{t_1}, \dots, X_{t_n} = x_{t_n}, X_s = x_s). \end{aligned} \quad (\text{P1})$$

A pMC is called *stationary* if it satisfies $T_t^{t+\Delta} = T_0^\Delta =: T_\Delta$ for all $t, \Delta \in \mathbb{R}_{\geq 0}$. In this case, there is a unique *transition rate matrix* Q —a matrix with non-negative off-diagonal elements and rows that sum up to zero—such that

$$(\forall t \in \mathbb{R}_{\geq 0}) T_\Delta = T_t^{t+\Delta} \approx I + \Delta Q \quad \text{for } \Delta \text{ suff. small.}$$

Furthermore, T_t then satisfies the differential equation

$$\frac{d}{dt} T_t = Q T_t, \quad \text{with } T_0 = I. \quad (\text{P2})$$

Similarly, for any non-stationary pMC there is a time-dependent transition rate matrix Q_t such that

$$(\forall t \in \mathbb{R}_{\geq 0}) T_t^{t+\Delta} \approx I + \Delta Q_t \quad \text{for } \Delta \text{ suff. small.}$$

Imprecise Markov chains

It is often infeasible to precisely specify the transition rate matrix Q of a stationary pMC. Furthermore, assuming stationarity is not always justified. Therefore, we here consider the case where the (time-dependent) transition rate matrix Q_t of a (non-stationary) pMC is only known to be contained in some (non-empty and bounded) set \mathcal{Q} . In other words, we consider the set $\mathbb{P}_{\mathcal{Q}}$ of all pMCs that are consistent with \mathcal{Q} , in the sense that

$$(\forall t \in \mathbb{R}_{\geq 0}) (\exists Q_t \in \mathcal{Q}) T_t^{t+\Delta} \approx I + \Delta Q_t \quad \text{for } \Delta \text{ suff. small.}$$

This set $\mathbb{P}_{\mathcal{Q}}$ characterises an *imprecise (continuous-time) Markov chain* (iMC) as follows. Analogous to (P1), we define a *lower transition operator* \underline{T}_s^t as

$$\begin{aligned} [\underline{T}_s^t f](x_s) &:= \underline{\mathbb{E}}(f(X_t) | X_s = x_s) \\ &= \underline{\mathbb{E}}(f(X_t) | X_{t_1} = x_{t_1}, \dots, X_{t_n} = x_{t_n}, X_s = x_s), \end{aligned} \quad (\text{I1})$$

where $\underline{\mathbb{E}}(\cdot)$ is the minimum of the conditional expectations that are induced by the set of consistent processes.

In case \mathcal{Q} has separately specified rows, Krak et al. (2017) show that $\underline{T}_t^{t+\Delta} = \underline{T}_0^\Delta =: \underline{T}_\Delta$ for all $t, \Delta \in \mathbb{R}_{\geq 0}$. Moreover, they show that \underline{T}_Δ is the unique operator that satisfies

$$\frac{d}{dt} \underline{T}_t = \underline{Q} \underline{T}_t, \quad \text{with } \underline{T}_0 = I. \quad (\text{I2})$$

In (I2), \underline{Q} is the so-called *lower transition rate operator* of \mathcal{Q} , which, for any $f \in \mathcal{L}(\mathcal{X})$ and $x \in \mathcal{X}$, is defined as

$$[\underline{Q}f](x) := \min \{ [Qf](x) : Q \in \mathcal{Q} \}. \quad (\text{I3})$$

Ergodicity

We are often interested in the long-term limit behaviour of stationary pMCs and iMCs. For iMCs, a special case is when

$$\lim_{t \rightarrow +\infty} [\underline{T}_t f](x) = \underline{E}_\infty(f) \quad \text{for all } f \in \mathcal{L}(\mathcal{X}) \text{ and } x \in \mathcal{X}.$$

If this is the case, then the iMC is said to be *ergodic* and $\underline{E}_\infty(f)$ is called the *limit lower expectation*. Similarly, a stationary pMC is ergodic if

$$\lim_{t \rightarrow +\infty} [T_t f](x) = E_\infty(f) \quad \text{for all } f \in \mathcal{L}(\mathcal{X}) \text{ and } x \in \mathcal{X},$$

where E_∞ is now called the limit expectation.

Handling state space explosion in Markov chains

How lumping introduces imprecision (almost) inevitably

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State space explosion

Precise Markov chains (or pMCs, as introduced in [Markov chains: An introduction](#)) are used ubiquitously to model systems with uncertain dynamics. Consider a stationary pMC and assume that we are interested in making inferences of the form

$$\lim_{t \rightarrow +\infty} \mathbb{E}(f(X_t)) = \lim_{t \rightarrow +\infty} \sum_{x \in \mathcal{X}} \pi_0(x) \mathbb{E}(f(X_t) | X_0 = x) = \lim_{t \rightarrow +\infty} \sum_{x \in \mathcal{X}} \pi_0(x) [T_t f](x),$$

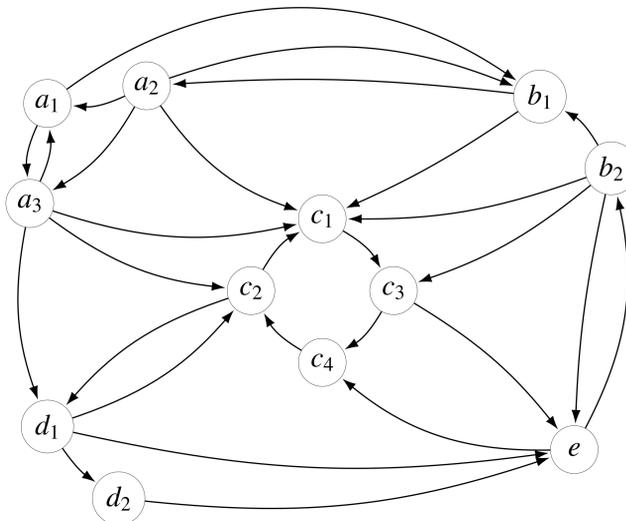
where f is a real-valued function on \mathcal{X} and π_0 is an initial probability distribution. If the pMC is ergodic, then we immediately obtain that

$$\lim_{t \rightarrow +\infty} \mathbb{E}(f(X_t)) = E_\infty(f) = \sum_{x \in \mathcal{X}} \pi_\infty(x) f(x),$$

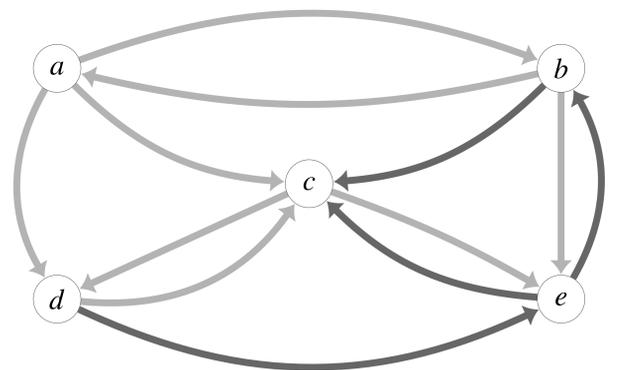
regardless of π_0 . It is well-known that π_∞ is the unique probability distribution on \mathcal{X} that satisfies the equilibrium condition

$$\pi_\infty Q = 0. \quad (1)$$

In case the number of states $|\mathcal{X}|$ is relatively small, this linear system of equations can be efficiently solved analytically or numerically. However, in many applications—see for instance [Modelling spectrum assignment in a two-service flexi-grid optical link](#)—the number of states grows **exponentially** with respect to the dimensions of the system! **This state space explosion makes (1) practically unsolvable for large systems.**



Lumping



Lumped pMC is stationary

So far, lumping states in order to make determining the steady-state distribution feasible was limited to the well-known special case where the *lumped process is a stationary pMC*. This occurs if the transition rate matrix Q of the original (stationary and ergodic) pMC satisfies

$$(\forall \hat{x}, \hat{y} \in \hat{\mathcal{X}}) \min_{x \in \hat{x}} \sum_{y \in \hat{y}} Q(x, y) = \hat{Q}(\hat{x}, \hat{y}) = \hat{Q}_\infty(\hat{x}, \hat{y}) = \max_{x \in \hat{x}} \sum_{y \in \hat{y}} Q(x, y). \quad (2)$$

Lumped pMC is not stationary

We now consider an original stationary pMC of which the transition rate matrix Q does not satisfy (2). In this case, all we can say for sure about the lumped pMC—without determining the actual distribution $\mathbb{P}(X_t = x)$ of the original pMC—is that, for all $t \in \mathbb{R}_{\geq 0}$,

$$(\forall \hat{x}, \hat{y} \in \hat{\mathcal{X}}, \hat{x} \neq \hat{y}) \min_{x \in \hat{x}} \sum_{y \in \hat{y}} Q(x, y) \leq \hat{Q}(\hat{x}, \hat{y}) \leq \max_{x \in \hat{x}} \sum_{y \in \hat{y}} Q(x, y). \quad (3)$$

We collect all transition rate matrices that satisfy (3) in the set $\hat{\mathcal{Q}}$, and let \hat{Q} denote the associated lower transition rate operator. By the theory of [imprecise Markov chains](#) (or iMCs), the lumped pMC is then contained in the set $\mathbb{P}_{\hat{\mathcal{Q}}}$. Consequently, we are guaranteed that

$$\underline{\mathbb{E}}(\hat{f}(\hat{X}_t) | \hat{X}_s = \hat{x}) \leq \mathbb{E}(f(X_t) | X_s \in \hat{x}) \leq \bar{\mathbb{E}}(\hat{f}(\hat{X}_t) | \hat{X}_s = \hat{x}) := -\underline{\mathbb{E}}(-\hat{f}(\hat{X}_t) | \hat{X}_s = \hat{x}).$$

It can be moreover shown that the obtained iMC is ergodic, whence

$$\underline{E}_\infty(\hat{f}) \leq E_\infty(f) \leq \bar{E}_\infty(\hat{f}) := -\underline{E}_\infty(-\hat{f}). \quad (4)$$

Want to know how to efficiently approximate $\underline{\mathbb{E}}(\hat{f}(\hat{X}_t) | \hat{X}_s = \hat{x})$ or $\underline{E}_\infty(\hat{f})$ up to some guaranteed maximal error? See [iMCs: Efficient computational methods with guaranteed error bounds](#).

Lumping reduces the number of states

One way to reduce the number of states is to **lump** together states. For example, in [Modelling spectrum assignment in a two-service flexi-grid optical link](#) we lump together states that correspond to the same higher-order description. In any case, this lumping of states yields the *lumped state space* $\hat{\mathcal{X}}$, which is a partition of \mathcal{X} .

The lumped stochastic process $(\hat{X}_t)_{t \in \mathbb{R}_{\geq 0}}$, which has state space $\hat{\mathcal{X}}$, is derived from the original stochastic process $(X_t)_{t \in \mathbb{R}_{\geq 0}}$ using the relation

$$(\forall \hat{x} \in \hat{\mathcal{X}}) \hat{X}_t = \hat{x} \Leftrightarrow X_t \in \hat{x}.$$

Throughout this poster, we only consider real-valued functions on $\hat{\mathcal{X}}$ that are constant over the elements of the partition, as such a function f can be trivially identified with a real-valued function \hat{f} on $\hat{\mathcal{X}}$.

Consider a stationary and ergodic pMC with state space \mathcal{X} and transition rate matrix Q . Assume, for the sake of simplicity, that the pMC is *irreducible*, in the sense that $\mathbb{P}(X_t = x) > 0$ for all $t > 0$ and all $x \in \mathcal{X}$. Given an initial distribution π_0 for the original pMC, we find that the lumped process $(\hat{X}_t)_{t \in \mathbb{R}_{\geq 0}}$ is a pMC with (time-dependent) transition rate matrix

$$\hat{Q}_t(\hat{x}, \hat{y}) = \frac{\sum_{x \in \hat{x}} \mathbb{P}(X_t = x) \sum_{y \in \hat{y}} Q(x, y)}{\sum_{x \in \hat{x}} \mathbb{P}(X_t = x)}.$$

Moreover, regardless of the initial distribution,

$$\lim_{t \rightarrow +\infty} \mathbb{E}(\hat{f}(\hat{X}_t)) = \sum_{\hat{x} \in \hat{\mathcal{X}}} \hat{\pi}_\infty(\hat{x}) \hat{f}(\hat{x}).$$

In this expression, $\hat{\pi}_\infty$ is the unique distribution on $\hat{\mathcal{X}}$ that satisfies

$$\hat{\pi}_\infty \hat{Q}_\infty = 0$$

where

$$\hat{Q}_\infty(\hat{x}, \hat{y}) := \frac{\sum_{x \in \hat{x}} \pi_\infty(x) \sum_{y \in \hat{y}} Q(x, y)}{\sum_{x \in \hat{x}} \pi_\infty(x)}.$$

In general, we can only precisely determine the (long-term limit of the) temporal evolution of the probability distribution over the lumps if we first determine the (long-term limit of the) temporal evolution of the probability distribution over the states of the original pMC.

Alternative bounds on $E_\infty(f)$

Assume we are only interested in determining (guaranteed bounds for) the limit expectation $E_\infty(f)$ of some real-valued function f on \mathcal{X} . An alternative to computing the limit lower expectation of the induced iMC is the following.

For any $A \subseteq \hat{\mathcal{X}}$ and $\hat{x} \in \hat{\mathcal{X}}$, we define

$$\hat{Q}_L(\hat{x}, A) := \min_{x \in \hat{x}} \sum_{y \in A} \sum_{y \in \hat{\mathcal{X}}} Q(x, y) \quad \text{and} \quad \hat{Q}_U(\hat{x}, A) := \max_{x \in \hat{x}} \sum_{y \in A} \sum_{y \in \hat{\mathcal{X}}} Q(x, y).$$

Let \mathcal{A} be a collection of subsets of $\hat{\mathcal{X}}$. If we let

$$\hat{\Pi}_{\mathcal{A}} := \{ \hat{\pi} \text{ a probability distribution on } \hat{\mathcal{X}} : (\forall A \in \mathcal{A}) \sum_{\hat{x} \in \hat{\mathcal{X}}} \hat{\pi}(\hat{x}) \hat{Q}_L(\hat{x}, A) \leq 0 \leq \sum_{\hat{x} \in \hat{\mathcal{X}}} \hat{\pi}(\hat{x}) \hat{Q}_U(\hat{x}, A) \},$$

then

$$\min_{\hat{\pi} \in \hat{\Pi}_{\mathcal{A}}} \sum_{\hat{x} \in \hat{\mathcal{X}}} \hat{\pi}(\hat{x}) \hat{f}(\hat{x}) \leq E_\infty(f) \leq \max_{\hat{\pi} \in \hat{\Pi}_{\mathcal{A}}} \sum_{\hat{x} \in \hat{\mathcal{X}}} \hat{\pi}(\hat{x}) \hat{f}(\hat{x}), \quad (5)$$

where the optimisations can be solved using a linear program.

How to pick \mathcal{A} and the tightness of the bounds of (5) compared to the bounds (4) of the iMC is the subject of ongoing research.