

ISIPTA '13

Allowing for probability zero in...

credal networks under epistemic irrelevance

...using sets of desirable gambles

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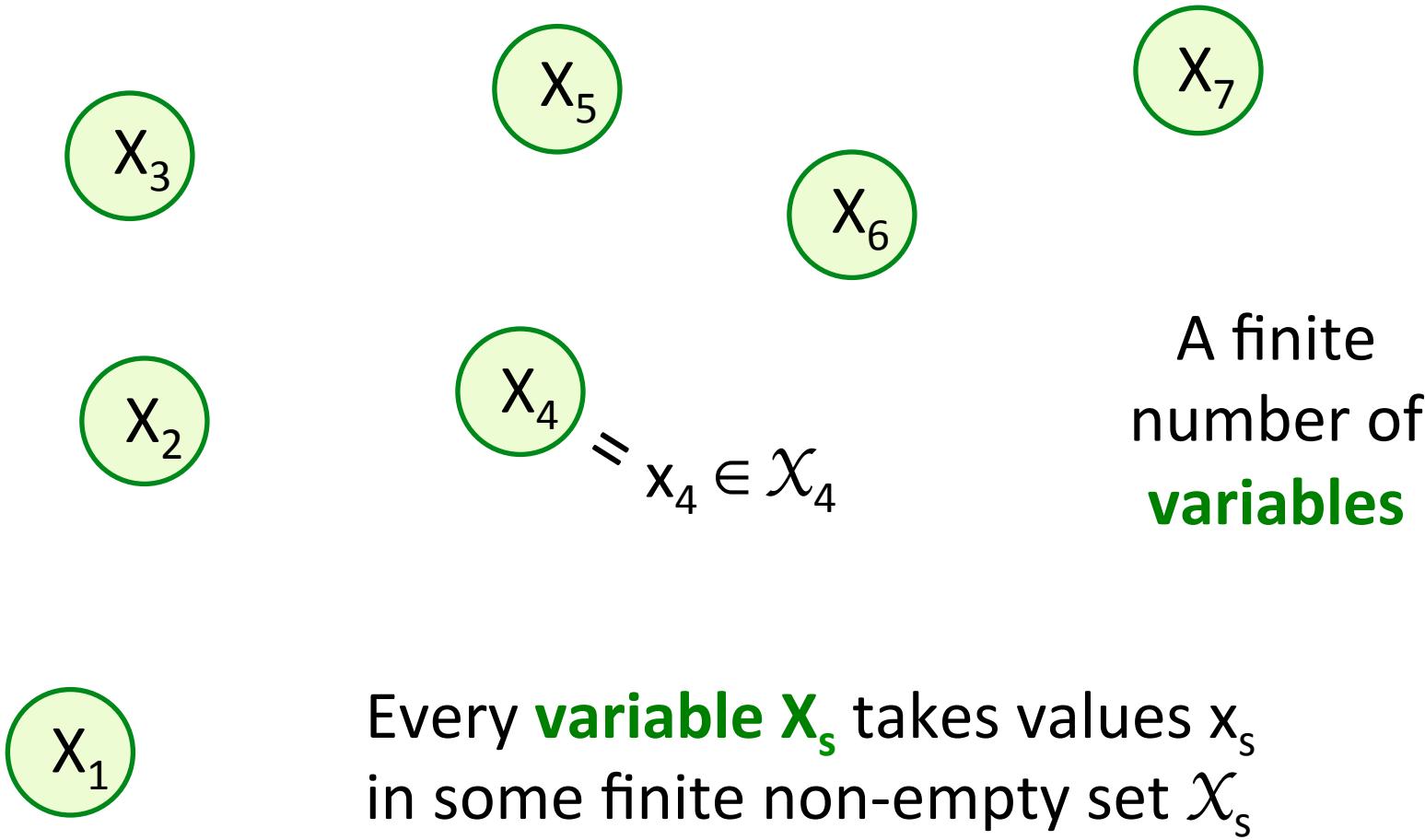


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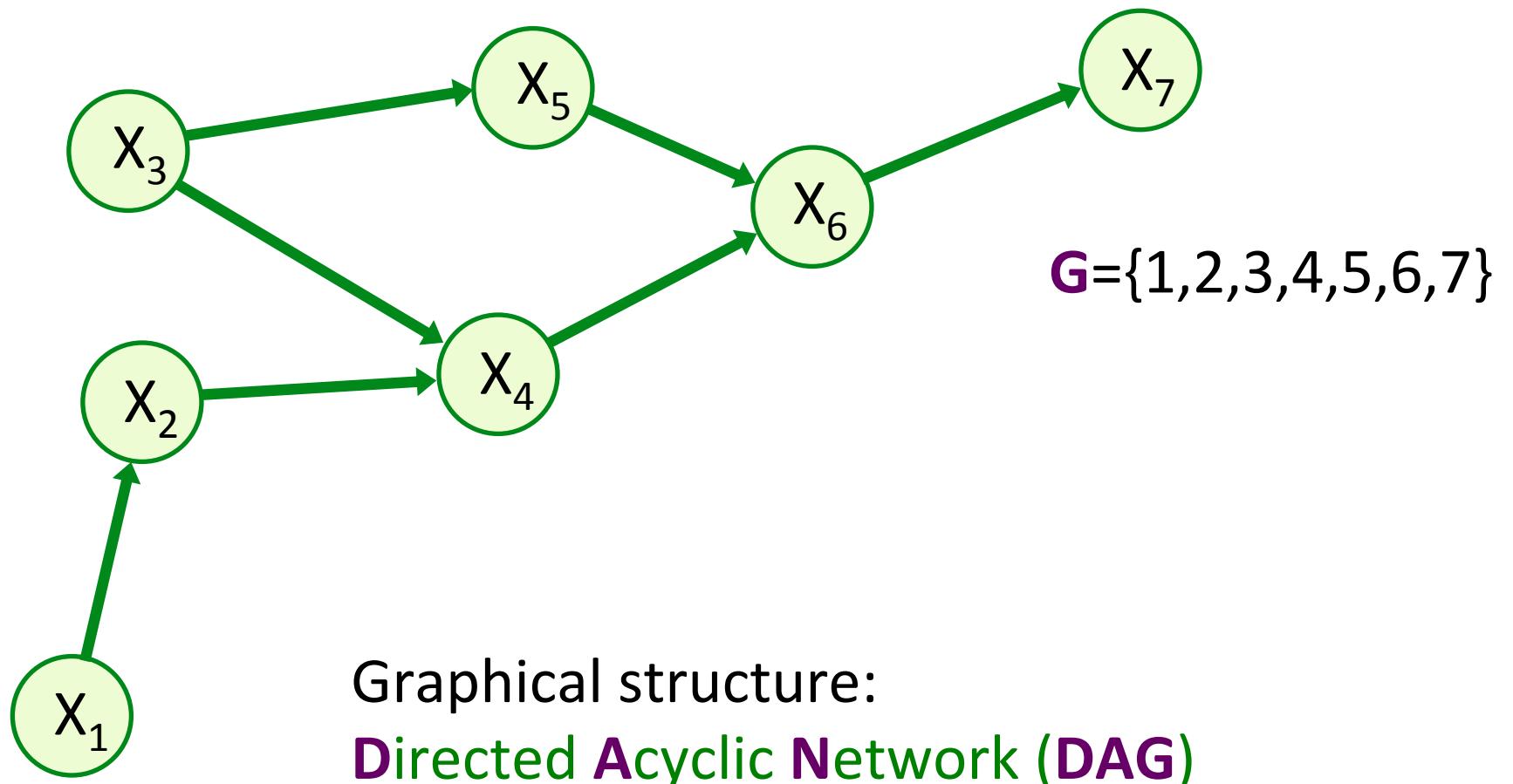


Márcio A. Diniz

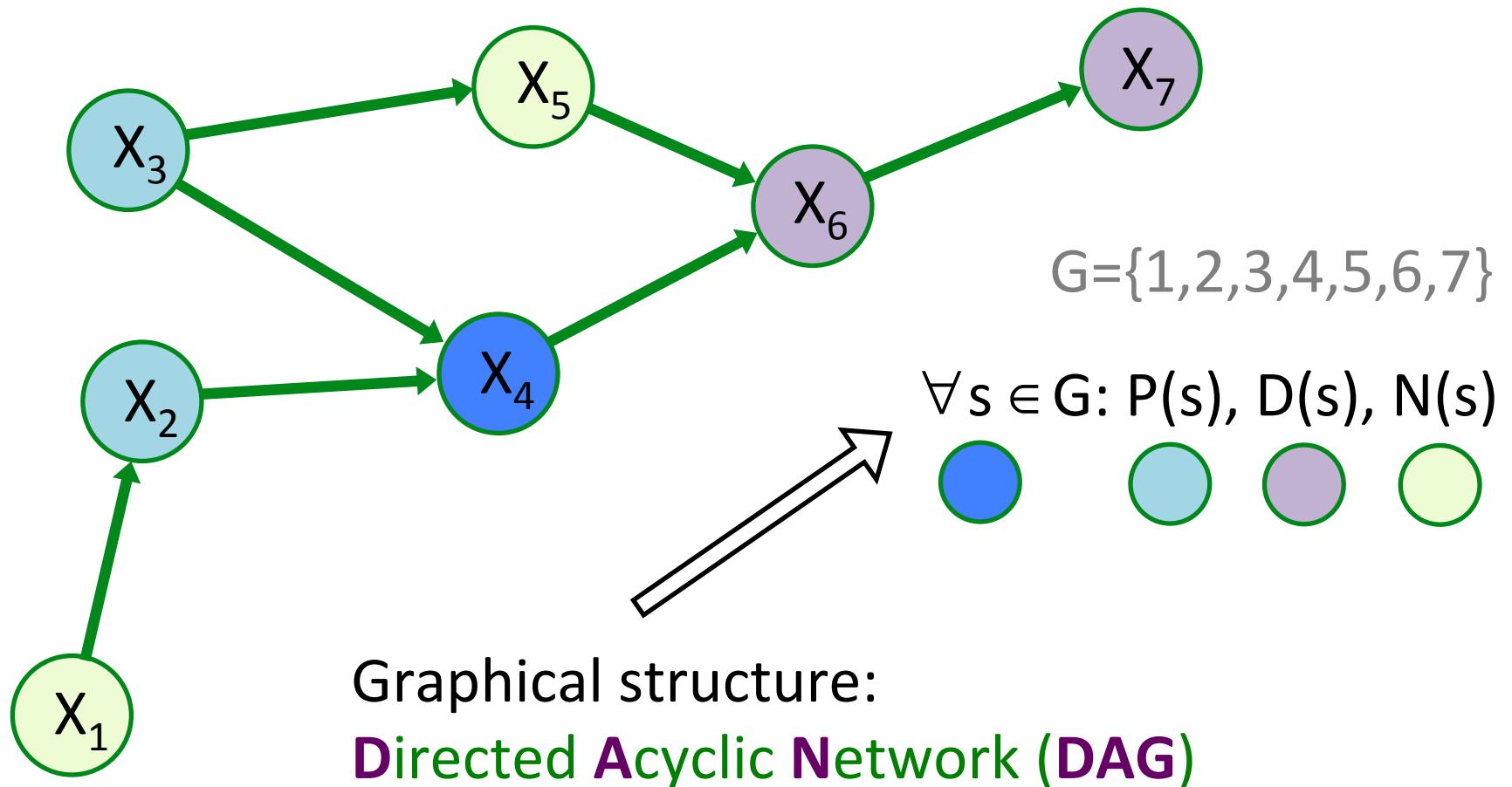
Credal networks: basic setup



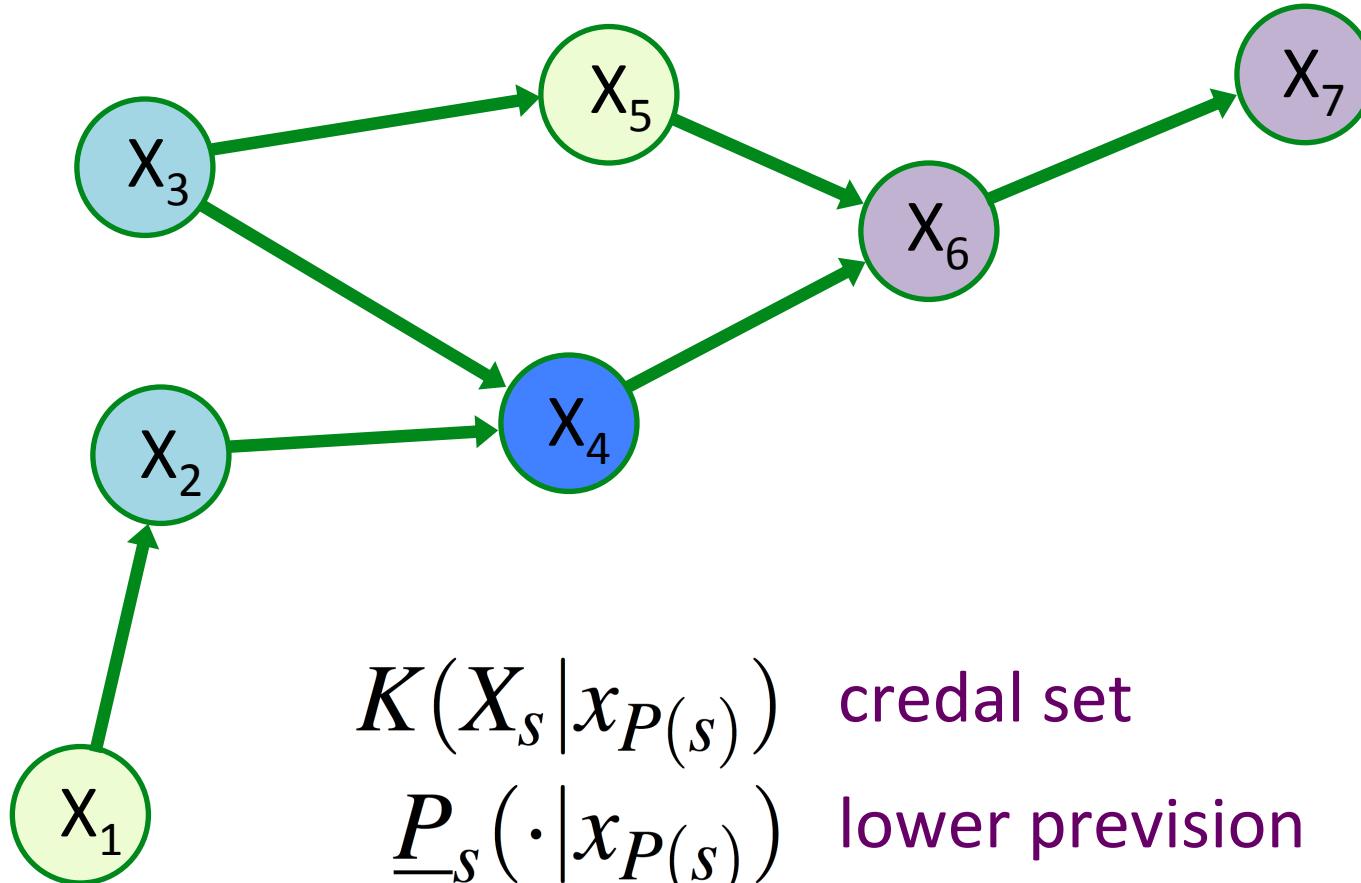
Credal networks: basic setup



Credal networks: basic setup



Credal networks: local uncertainty models

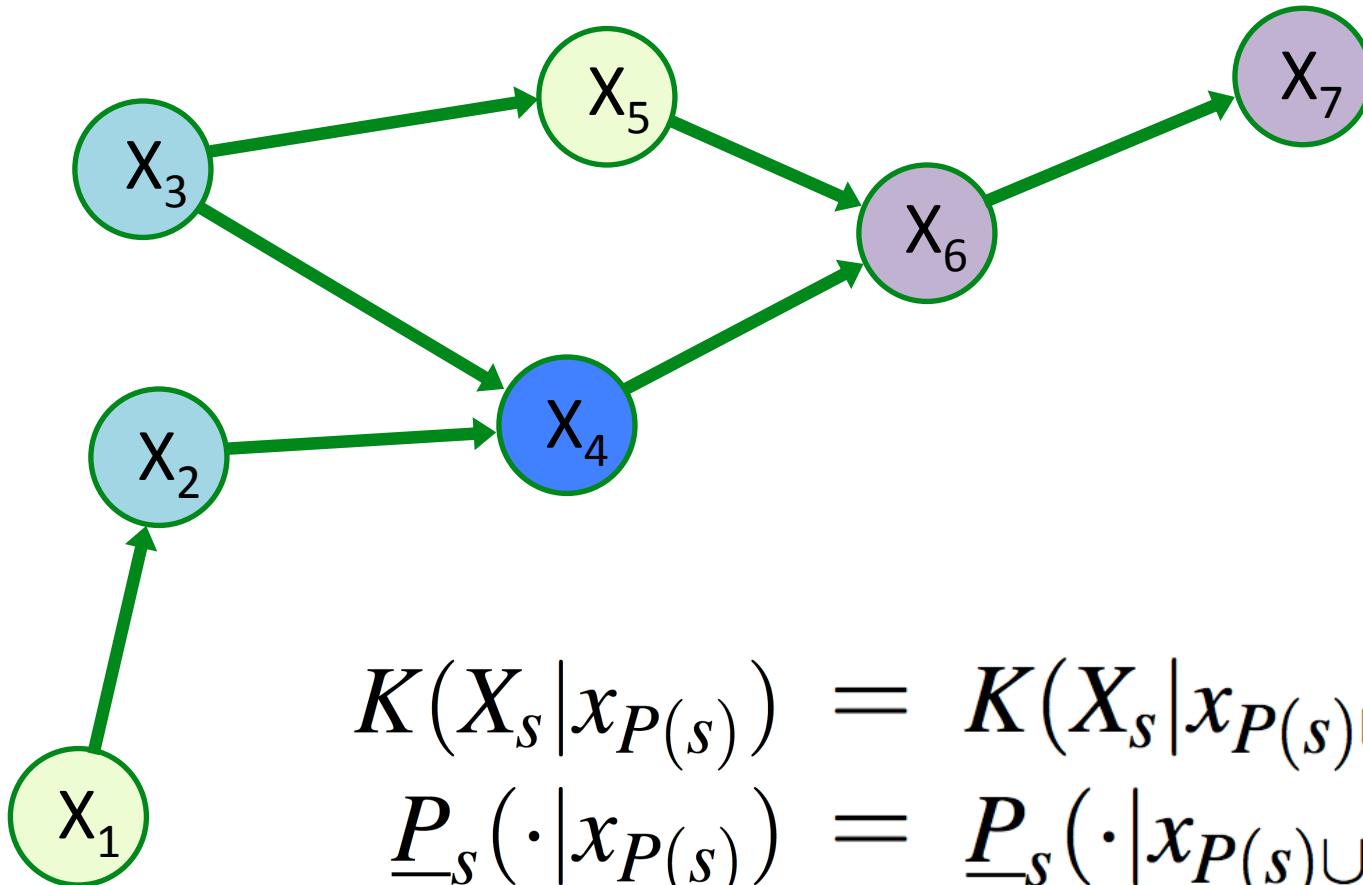


$K(X_s | x_{P(s)})$ credal set

$\underline{P}_s(\cdot | x_{P(s)})$ lower prevision

$\mathcal{D}_s | x_{P(s)}$ set of desirable gambles

Credal networks: epistemic irrelevance

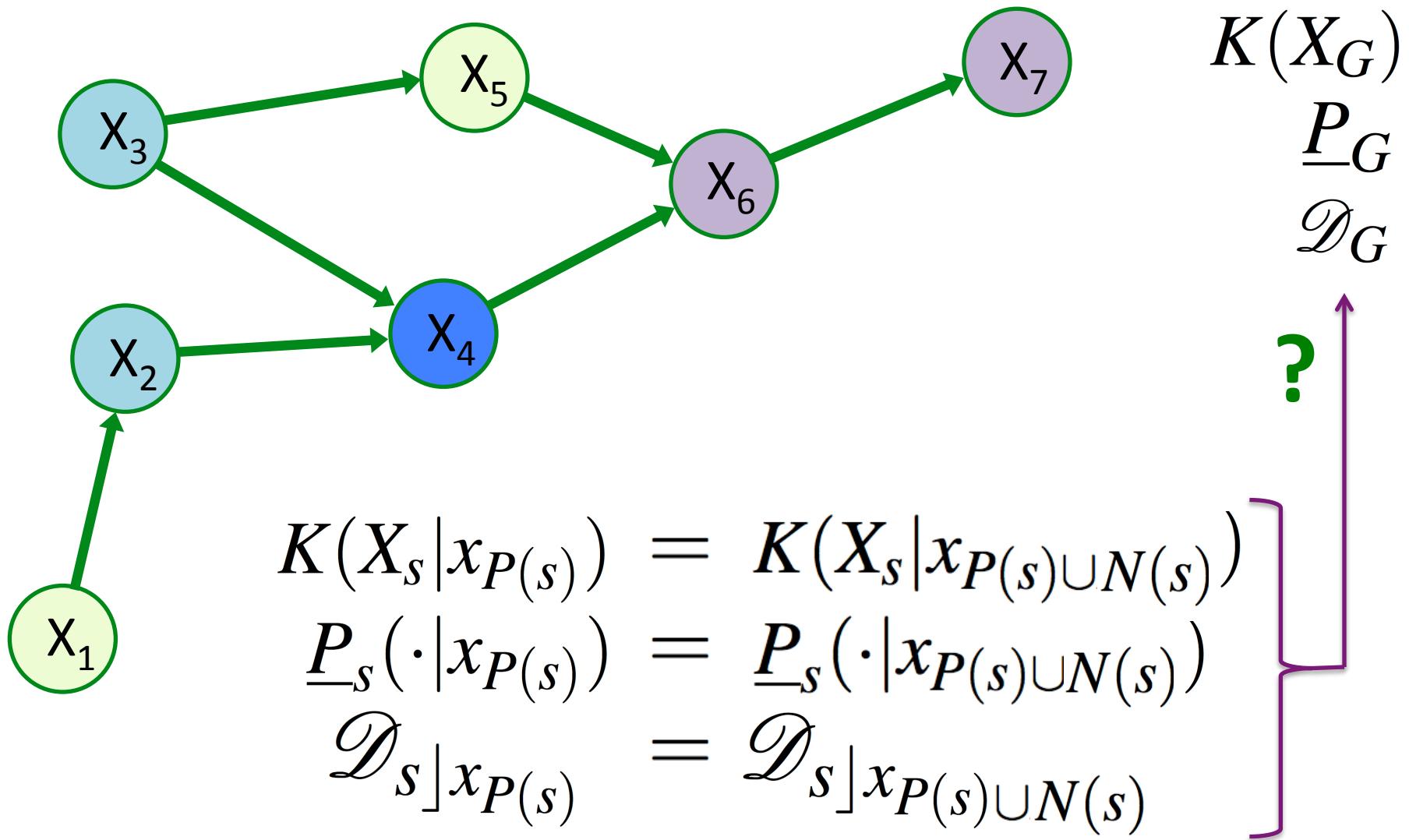


$$K(X_s | x_{P(s)}) = K(X_s | x_{P(s) \cup N(s)})$$

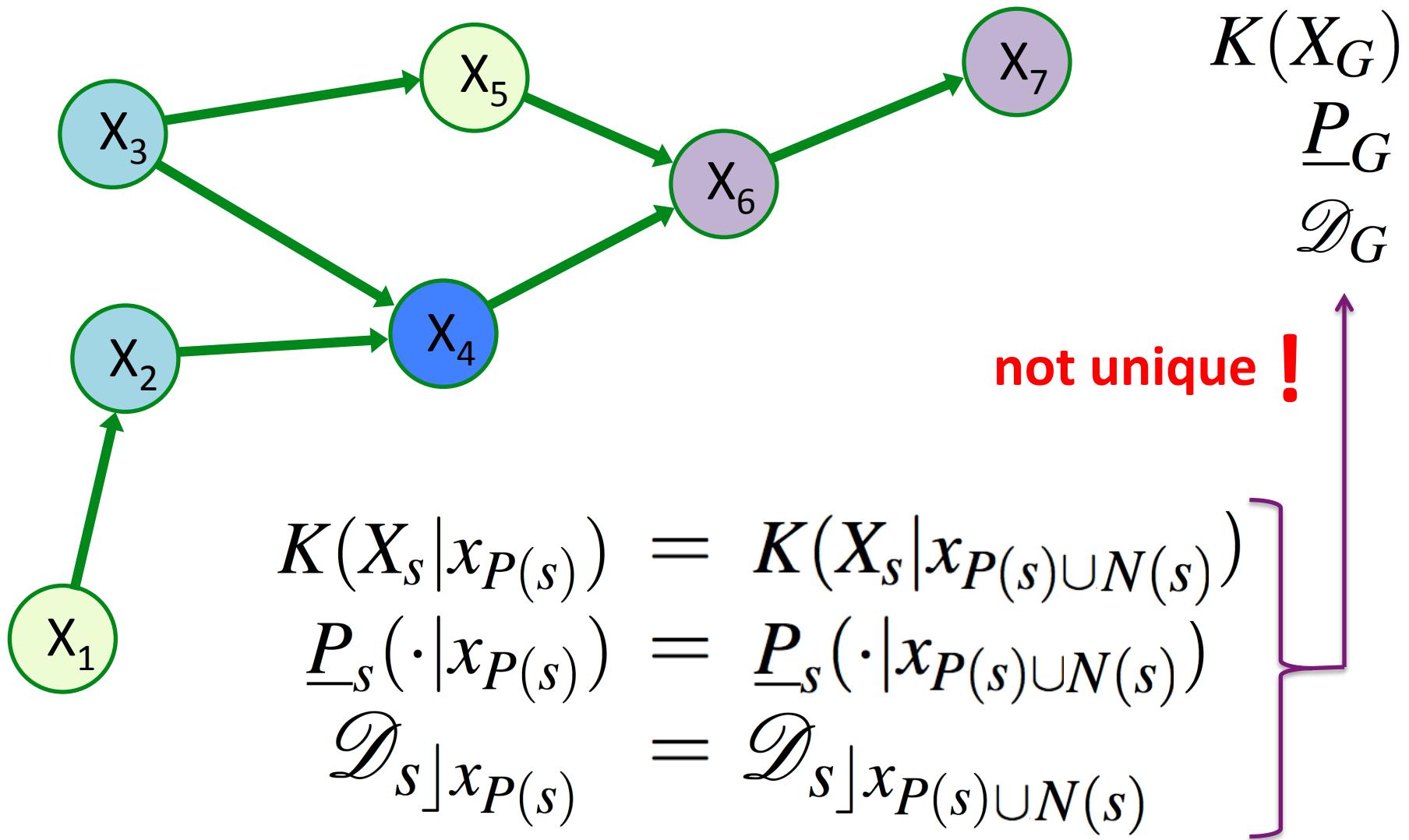
$$\underline{P}_s(\cdot | x_{P(s)}) = \underline{P}_s(\cdot | x_{P(s) \cup N(s)})$$

$$\mathcal{D}_s | x_{P(s)} = \mathcal{D}_s | x_{P(s) \cup N(s)}$$

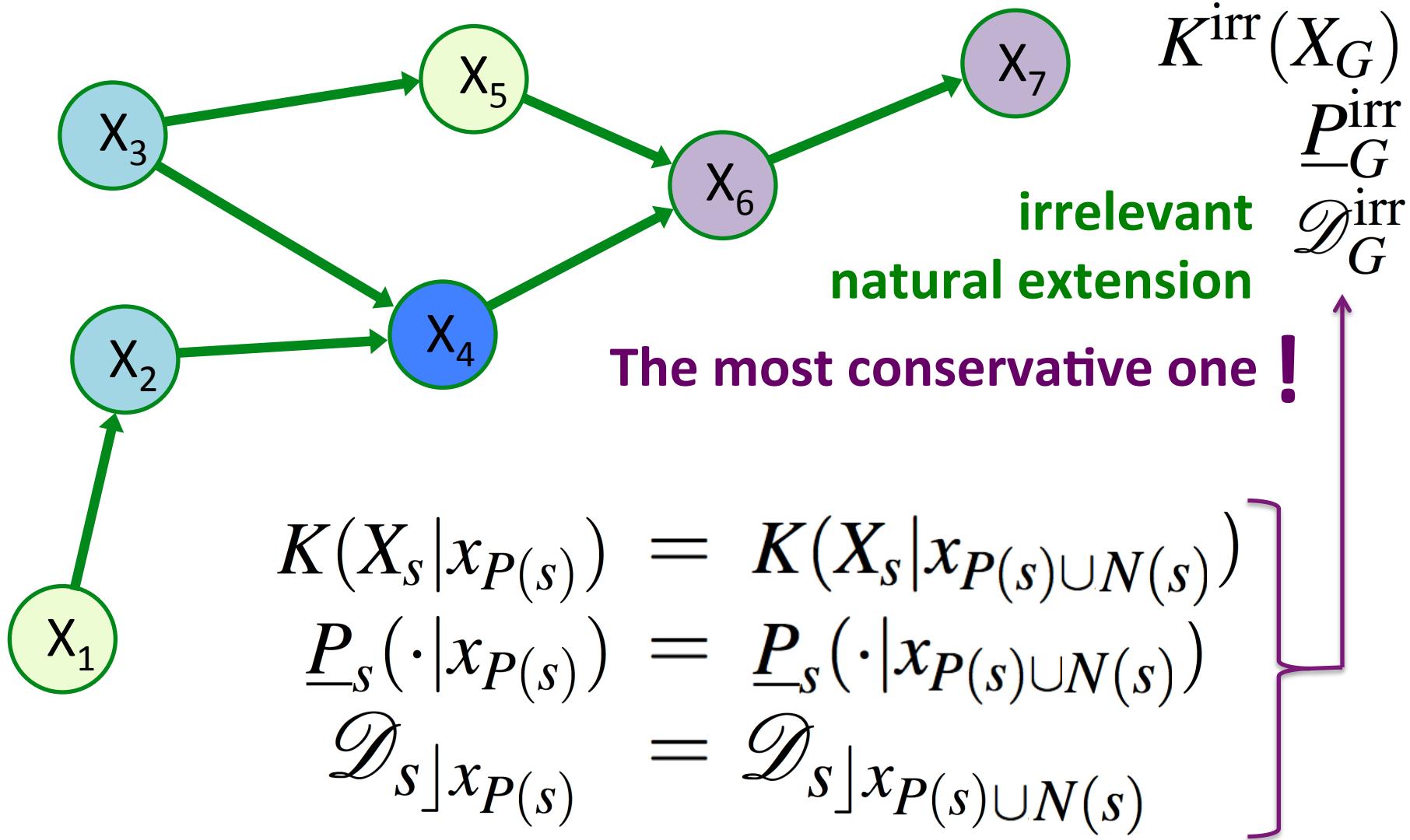
Credal networks: a joint model



Credal networks: a joint model

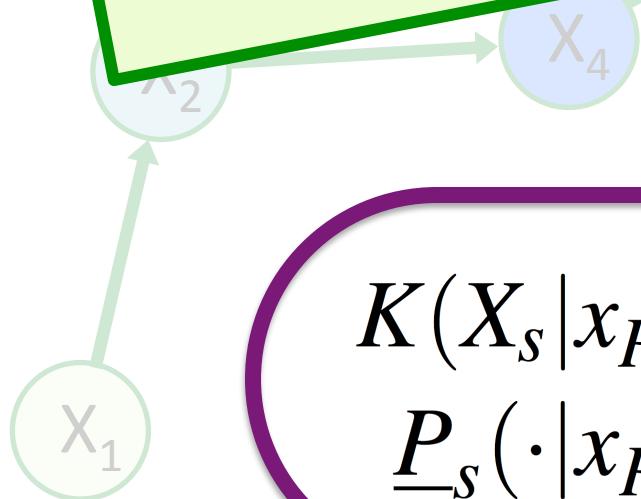


Credal networks: a joint model



Credal networks: a joint model

Allowing for probability zero
in credal networks under
epistemic irrelevance



The model

$$K(X_s | x_{P(s)}) = K(X_s | x_{P(s) \cup N(s)})$$

$$\underline{P}_s(\cdot | x_{P(s)}) = \underline{P}_s(\cdot | x_{P(s) \cup N(s)})$$

$$\mathcal{D}_s[x_{P(s)}] = \mathcal{D}_s[x_{P(s) \cup N(s)}]$$



$$K^{\text{irr}}(X_G)$$

$$\underline{P}^{\text{irr}}_G$$

$$\mathcal{D}^{\text{irr}}_G$$

$$\uparrow$$

Allowing for probability zero in credal networks under epistemic irrelevance

Abstract of the paper

We generalise Cozman's (2000) concept of a credal network under epistemic irrelevance to the case where lower (and upper) probabilities are allowed to be zero. Our main definition is expressed in terms of coherent lower previsions (BMT) and imposes epistemic irrelevance by means of strong coherence rather than element-wise Bayes's rule (I, II, III). We also present a number of alternative representations for the resulting joint model, both in terms of lower previsions and credal sets, amongst which an intuitive characterisation of the joint credal set by means of linear constraints (IV). We then apply our method to a simple case: the independent natural extension for two binary variables (V). This allows us to, for the first time, find analytical expressions for the extreme points of this special type of independent product.

Basic modelling tools (BMT)

We will model a subject's beliefs about the value that a variable X assumes in some set \mathcal{X} by means of two different, although mathematically equivalent, imprecise-probabilistic methods. The approach that is perhaps best known is to use a credal set $K(X)$, defined as a closed convex subset of $\Sigma_{\mathcal{X}}$, which is the set containing all probability mass functions on \mathcal{X} . The second approach is to use the associated and coherent lower prevision $E^*(\mathcal{X})$, where \mathcal{X} is the set of all gambles on \mathcal{X} : it is given by $E^*(\mathcal{X}) = \min\{E_f(\mathcal{X}): p(X) \in K(X)\}$, where E_f is the expectation operator (precision) for the probability mass function $p(X)$. The credal set of such a coherent lower precision is given by $K(X) = \{p \in \Sigma_{\mathcal{X}}: (Y \in \mathcal{X}) \rightarrow E_p(Y) \geq E^*(Y)\}$, thereby establishing the mathematical equivalence.

Linear constraints (IV)

It is well known that each local credal set $K(X_{\text{parent}})$ ($i \in \mathcal{G}$) and $\mathcal{X}_{\text{parent}}$ can be reduced to a local uncertainty constraint and a set of linear homogeneous inequalities of the form $\sum_{j \in \mathcal{G} \setminus \{i\}} p_{ij} x_{j,y_{ij}} - p_i x_{i,y_i} \geq 0$. We show that, even without the positivity assumption (I), these local constraints can be used to derive an intuitive characterisation of the irrelevant natural extension $K''(X)$ (II) in terms of linear constraints. $K''(X)$ is the solution set to the global utility constraint and, for all $i \in \mathcal{G}$, $x_{i,y_{ij}} \in \mathcal{X}_{\text{parent}}$ and $y \in T_i(x_{i,y_i})$, a linear homogeneous inequality $\sum_{j \in \mathcal{G} \setminus \{i\}} p_{ij} x_{j,y_{ij}} - p_i x_{i,y_i} \geq 0$.

Independent natural extension for two binary variables (V)

$K(X_1)$
 $X_1 = \{x_1, x_2\}$

$K(X_2)$
 $X_2 = \{x_3, x_4\}$



For the simple model network above, $K''(X)$ is the so-called independent natural extension of $K(X_1)$ and $K(X_2)$. Every $K(X_i)$, with $i \in \{1, 2\}$, is fully determined by the lower probability $p_i(x_i)$ and upper probability $p_i(x_i')$ of 'heads'. The probability of 'tails' is bounded by $p(x_i) \leq 1 - p(x_i')$ and $p(x_i) \geq 1 - p(x_i')$. Using the linear constraints in (IV), we have derived analytical expressions for the extreme points of $K''(X)$. They can be found using the table and diagram to the right; see our paper for more details.

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Local

Uncertainty models (I)

With every node x_i ($i \in \mathcal{G}$), we associate some finite, non-empty set \mathcal{X}_i . This set of all nodes is denoted by variable X taking values in empty set \mathcal{X} . The set of all nodes for every subset $G \subseteq \mathcal{G}$, the joint values in $\mathcal{X}_G := \prod_{i \in G} \mathcal{X}_i$. For every $x \in \mathcal{X}_G$, we denote its parents by $\text{parent}(x)$ and its children by $\text{child}(x)$. Similar to what is done in classical Bayesian networks, we attach local uncertainty models to the nodes. For all $i \in \mathcal{G}$ and every instantiation x_i of X_i , we require a credal equivalence $K(X_i|x_{\text{parent}})$ ($K''(X_i)$) (BMT).

Imposing epistemic irrelevance (II)

We provide the practical structure of the network with the following interpretation: for any node $i \in \mathcal{G}$, its non-parent non-descendant variables $X_{\text{irr},i}$ are epistemically irrelevant to X_i conditional on $X_{\text{parent},i}$. (In our paper, we also require this for subsets of $N(i)$.) We do not impose these additional assessments on this poster because we have recently discovered that, at least for the unconditional joint model, they are redundant.) Put more mathematically, and using $\text{int}(x)$ as a shorthand notation for $P(x) \cup V(x)$, we require that:

$$K(X_i|x_{\text{parent},i}) = K(X_i|x_{\text{parent},i} \cup \text{int}(x_{\text{irr},i})) \quad (1)$$

the right hand side of these equations being provided by the local models (I). In order to translate this into a property of a joint model $K(X_G)$, it is often assumed that for every $p(X_i) \in K(X_i)$, all events have strictly positive probability (Cozman 2000). Under this assumption, $K(X_G)$ can be conditioned by means of element-wise Bayes's rule (applying Bayes's rule to every $p(X_i) \in K(X_i)$), thereby making it possible to impose Eq. (1).

We drop this positivity assumption by using an approach based on lower previsions, replacing Eq. (1) by the equivalent (BMT) statement that:

$$E_p(X_{\text{irr},i}) = E_p(x_{\text{irr},i}) \text{ for all } i \in \mathcal{G} \text{ and } x_{\text{irr},i} \in \mathcal{X}_{\text{irr},i} \quad (2)$$

where, again, the right hand side is provided by the local models (I). Since, without the positivity assumption, conditioning is not uniquely defined, we use a different method for making the conditional models in Eq. (2) consistent with the joint model E_p : we require them to be (strongly) coherent. We prove that, in our particular case, this is equivalent to requiring that

$$E_p(E_{\text{irr},i}[x] - E_p(x_{\text{irr},i})) = 0 \text{ and } E_p(E_{\text{irr},i}[x] - E_p(x_{\text{irr},i})) = 0 \quad (3)$$

for all $i \in \mathcal{G}$, $x_{\text{irr},i} \in \mathcal{X}_{\text{irr},i}$ and $x \in \mathcal{X}$. This formula is known as Generalised Bayes's Rule (GBR) and is equivalent to element-wise Bayes's rule if the positivity assumption is satisfied. It should therefore be clear that our approach is an extension of the one by Cozman (2000), coinciding with it under the positivity assumption.

The properties that we impose on our network (1-3) can be satisfied by multiple coherent lower previsions E_p on $\mathcal{X}(X_G)$. However, amongst them, there is a unique most conservative (pointwise smallest) one. We call it the *irrelevant natural extension* of the network and denote it by E_p^{**} . We show that E_p^{**} is the pointwise smallest coherent lower prevision on $\mathcal{X}(X_G)$ such that for all $x \in \mathcal{X}$, $E_p^{**}(x) = E_p(x_{\text{irr},i})$ and $g \in \mathcal{G}(X)$,

$$E_p^{**}(E_{\text{irr},i}[x] - E_p(x_{\text{irr},i})) = 0.$$

We also prove the following simple characterisation of the corresponding credal set $K''(X_G)$: it consists of all probability mass functions $p(X_i)$ on X_i for which for all $i \in \mathcal{G}$ and $x_{\text{parent},i} \in \mathcal{X}_{\text{parent},i}$, there are a real number $\lambda \geq 0$ and a $p(X_i|x_{\text{parent},i}) \in K(X_i|x_{\text{parent},i})$ such that

$$\sum_{x_{\text{irr},i} \in \mathcal{X}_{\text{irr},i}} p(x_{\text{irr},i}|x_{\text{parent},i}) = \lambda p(X_i|x_{\text{parent},i}),$$

where we use $\mathcal{X}(x)$ to denote the set consisting of the descendants of the node x .

We believe that most of the marginalisation and graph-like properties that are presented on our other poster can be translated to the current framework. Combined with linear programming (IV), this might allow for efficient inference algorithms.



Definition by means of strong coherence

Allowing for probability zero in credal networks under epistemic irrelevance

Abstract of the paper

We generalise Cozman's (2000) concept of a credal network under epistemic irrelevance to the case where lower (and upper) probabilities are allowed to be zero. Our main definition is expressed in terms of coherent lower previsions (BMT) and imposes epistemic irrelevance by means of strong coherence rather than element-wise Bayes's rule (I, II, III). We also present a number of alternative representations for the resulting joint model, both in terms of lower previsions and credal sets, amongst which an intuitive characterisation of the joint credal set by means of linear constraints (IV). We then apply our method to a simple case: the independent natural extension for two binary variables (V). This allows us to, for the first time, find analytical expressions for the extreme points of this special type of independent product.

Basic modelling tools (BMT)

We will model a subject's beliefs about the value that a variable X assumes in some set \mathcal{X} by means of two different, although mathematically equivalent, imprecise-probabilistic methods. The approach that is perhaps best known is to use a credal set $K(X)$, defined as a closed convex subset of $\Sigma_{\mathcal{X}}$, which is the set containing all probability mass functions on \mathcal{X} . The second approach and coherent lower previsions are based on $\mathcal{W}(\mathcal{X})$, where $\mathcal{W}(\mathcal{X})$ is the set of all gambles on \mathcal{X} . It is given by $\mathcal{E}(f) = \min\{p_f(x) : p(x) \in K(X)\}$ for all $f \in \mathcal{W}(\mathcal{X})$, where p_f is the expectation operator (or precision) for the probability mass function p . The credal set of such a coherent lower prevision p is given by $K(p) = \{p \in \Sigma_{\mathcal{X}} : (Y \in \mathcal{Y}) \rightarrow p(Y) \geq \mathcal{E}(Y)\}$, thereby establishing the mathematical equivalence.

Linear constraints (IV)

It is well known that each local credal set $K(X_{p(i)})$ ($i \in G$) and $x_{p(i)} \in \Sigma_{X_{p(i)}}$ is the solution set to a local uncertainty constraint and a set of linear homogeneous inequalities of the form $\sum_{x \in X_{p(i)}} p_i(x)x_{p(i)}(x) \geq 0$, where x takes values in some (possibly infinite, but often finite) set $\Sigma_{X_{p(i)}} \subseteq \mathcal{W}(\mathcal{X}_{p(i)})$. We show that, even without the positivity assumption (I), these local constraints can be used to derive an intuitive characterisation of the irrelevant natural extension $K''(X)$ (II) in terms of linear constraints. $K''(X)$ is the solution set to the global uniform constraint and, for all $i \in G$, $x_{p(i)} \in \Sigma_{X_{p(i)}}$ and $y \in \Sigma_{X_{p(i)}} \setminus \{x_{p(i)}\}$, a linear homogeneous inequality $\sum_{x \in X_{p(i)}} \sum_{y \in X_{p(i)} \setminus \{x_{p(i)}\}} p_i(x)p_i(y)x_{p(i)}(x)y_{p(i)} \geq 0$.

Independent natural extension for two binary variables (V)

$K(X_1)$
 $X_1 = \{x_1, x_2\}$
 $\mathcal{X}_1 = \{\emptyset, x_1, x_2\}$

$K(X_2)$
 $X_2 = \{x_3, x_4\}$
 $\mathcal{X}_2 = \{\emptyset, x_3, x_4\}$



For the simple model network above, $K''(X)$ is the so-called independent natural extension of $K(X_1)$ and $K(X_2)$. Every $K(X_i)$, with $i \in \{1, 2\}$, is fully determined by the lower probability $p_i(x)$ and upper probability $p_i(y)$ of 'heads'. The probability of 'tails' is bounded by $p_i(x) \leq 1 - p_i(y)$ and $p_i(y) \leq 1 - p_i(x)$. Using the linear constraints in (IV), we have derived analytical expressions for the extreme points of $K''(X)$. They can be found using the table and diagram to the right; see our paper for more details.

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Local uncertainty models (I)

With every node i of a finite directed acyclic graph (DAG), we associate a variable X_i taking values in some finite, non-empty set \mathcal{X}_i . The set of all nodes is denoted by G . For every subset $\mathcal{S} \subseteq G$, the joint variable $X_{\mathcal{S}}$ takes values in $\mathcal{S} \times \dots \times \mathcal{X}_{G \setminus \mathcal{S}}$. For every $i \in G$, we denote by $\mathcal{P}(i)$ the set consisting of the parent nodes of i . Similar to what is done in classical Bayesian networks, we attach local uncertainty models to the nodes of the network, conditional on the value of their parents. For all $i \in G$ and every instantiation $x_{\mathcal{P}(i)} \in \Sigma_{\mathcal{P}(i)}$ of $X_{\mathcal{P}(i)}$, we require a credal set $K(X_{\{i\}}, x_{\mathcal{P}(i)})$ of $\mathcal{W}(\mathcal{X}_i)$ (BMT).

Imposing epistemic irrelevance (II)

We provide the practical structure of the network with the following interpretation: for any node $i \in G$, its non-parent non-descendant variables $X_{\mathcal{N}^c(i)}$ are epistemically irrelevant to X_i conditional on $X_{\mathcal{P}(i)}$. (In our paper, we also provide this for subsets of $N^c(i)$.) We do not impose these additional assess-

ments on X_i if $X_i \in \mathcal{N}^c(i)$, where $\mathcal{N}^c(i) = \mathcal{X}_i \times \mathcal{X}_{G \setminus \{i\}}$. It is given by $\mathcal{E}(f) = \min\{p_f(x) : p(x) \in K(X_i)\}$ for all $f \in \mathcal{W}(\mathcal{X}_i)$, where p_f is the expectation operator (or precision) for the probability mass function p . The credal set of such a coherent lower prevision p is given by

$$K(p) = \{p \in \Sigma_{\mathcal{X}_i} : (Y \in \mathcal{Y}) \rightarrow p(Y) \geq \mathcal{E}(Y)\},$$

the unconditional joint model, they are redundant.) Put more mathematically, and using $\mathcal{W}(x)$ as a shorthand notation for $\mathcal{W}(\mathcal{X}) \cup \{x\}$, we require that

$$\mathcal{E}(p_{\mathcal{N}^c(i)}) = K(X_{\{i\}}, x_{\mathcal{P}(i)}) \text{ for all } i \in G \text{ and } x_{\mathcal{P}(i)} \in \Sigma_{\mathcal{P}(i)}, \quad (1)$$

the right hand side of these equations being provided by the local models (I). In order to translate this into a property of a joint model $K(X_G)$, it is often assumed that for every $p(x) \in K(X_G)$, all events have strictly positive probability (Cozman 2000). Under this assumption, $K(X_G)$ can be conditioned by means of element-wise Bayes's rule (applying Bayes's rule to every $p(x) \in K(X_G)$), thereby making it possible to impose Eq. (1).

We drop this positivity assumption by using an approach based on lower previsions, replacing Eq. (1) by the equivalent (BMT) statement that

$$\mathcal{E}(p_{\mathcal{N}^c(i)}) = E(p_{\mathcal{N}^c(i)}) \text{ for all } i \in G \text{ and } x_{\mathcal{P}(i)} \in \Sigma_{\mathcal{P}(i)}, \quad (2)$$

where, again, the right hand side is provided by the local models (I). Since, without the positivity assumption, conditioning is not uniquely defined, we use a different method for making the conditional models in Eq. (2) consistent with the joint model p : we require them to be (strongly) coherent. We prove that, in our particular case, this is equivalent to requiring that

$$\mathbb{E}(p_{\mathcal{N}^c(i)}[x - E(p_{\mathcal{N}^c(i)})]) = 0 \text{ and } \mathbb{E}(p_{\mathcal{N}^c(i)}[x - E(p_{\mathcal{N}^c(i)})]) \geq 0$$

for all $i \in G$, $x_{\mathcal{P}(i)} \in \Sigma_{\mathcal{P}(i)}$ and $x \in \mathcal{X}_i$. This formula is known as Generalised Bayes's Rule (GBR) and is equivalent to element-wise Bayes's rule if the positivity assumption is satisfied. It should therefore be clear that our approach is an extension of the one by Cozman (2000), coinciding with it under the positivity assumption.

Irrelevant natural extension (III)

The properties that we impose on our network (1, II) can be satisfied by multiple coherent lower previsions $p_{\mathcal{C}}$ on $\mathcal{W}(X_G)$. However, amongst them, there is a unique most conservative (pointwise smallest) one. We call it the irrelevant natural extension of the network and denote it by $K''(X)$. We show that $K''(X)$ is the pointwise smallest coherent lower prevision on $\mathcal{W}(X_G)$ such that for all $x \in G$, $x_{\mathcal{P}(x)} \in \Sigma_{\mathcal{P}(x)}$ and $y \in \mathcal{X}_G$,

$$\mathbb{E}(p_{\mathcal{C}}[x - E(p_{\mathcal{C}})]) = 0.$$

We also prove the following simple characterisation of the corresponding credal set $K''(X)$: it consists of all probability mass function $p(x)$ on X_G for which for all $i \in G$ and $x_{\mathcal{P}(i)} \in \Sigma_{\mathcal{P}(i)}$, there are a real number $\lambda \geq 0$ and a $p(X_i|x_{\mathcal{P}(i)}) \in K(X_i|x_{\mathcal{P}(i)})$ such that

$$\mathbb{E}_{x_{\mathcal{P}(i)} \sim p_{\mathcal{C}}} p(X_i|x_{\mathcal{P}(i)}) p(X_{\mathcal{N}^c(i)}|x_{\mathcal{P}(i)}) = \lambda p(X_i|x_{\mathcal{P}(i)}),$$

where we use $\mathbb{E}[x]$ to denote the set consisting of the descendants of the node x .

We believe that most of the marginalisation and graphical properties that are presented on our other poster can be translated to the current framework. Combined with linear programming (IV), this might allow for efficient inference algorithms.

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Description in terms of linear constraints

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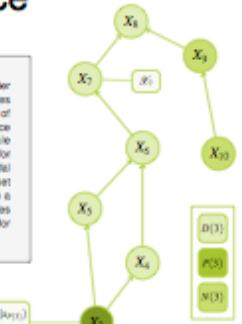
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Linear constraints (IV)

It is well known that each local credal set $K(X_{\text{parent}})$ ($x \in G$ and $x_{\text{parent}} \in \mathcal{X}_{\text{parent}}$) translates into a local uncertainty constraint and a set of linear homogeneous inequalities of the form $\sum_{x \in \mathcal{X}_{\text{parent}}} p(x_{\text{parent}}|y_x) \geq 0$, where y_x takes values in some (possibly infinite, but often finite) set $\Gamma(x_{\text{parent}}) \subset \mathcal{F}(x)$. We show that, even without the positivity assumption (I), these local constraints can be used to derive an intuitive characterisation of the irrelevant natural extension $K^*(X)$ (II) in terms of linear constraints. $K^*(X)$ is the solution set to the global utility constraint and, for all $x \in G$, $x_{\text{parent}} \in \mathcal{X}_{\text{parent}}$ and $y \in \Gamma(x_{\text{parent}})$, a linear homogeneous inequality $\sum_{x \in \mathcal{X}_{\text{parent}}} p(x_{\text{parent}}|y_x) p(x|x_{\text{parent}}, y_x) \geq 0$.

Independent natural extension for two binary variables (V)

For the simple model network above, $K(X_1)$ is the so-called independent natural extension of $K(X_1)$ and $K(X_2)$. Every $K(X_i)$, with $i \in \{1, 2\}$, is fully determined by the lower probability $p(h_i)$ and upper probability $p(u_i)$ of 'heads'. The probability of 'tails' is bounded by $p(u_i) - p(h_i)$ and $p(h_i) + 1 - p(u_i)$. Using the linear constraints in (IV), we have derived analytical expressions for the extreme points of $K^*(X)$. They can be found using the table and diagram to the right; see our paper for more details.



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Imposing epistemic irrelevance (II)

We provide the practical structure of the network with the following interpretation: for any node $x \in G$, its non-parent non-descendant variables $X_{G \setminus c}$ are epistemically irrelevant to X conditional on $X_{G \setminus c}$. (In our paper, we also require this for subsets of $N(x)$.) We do not impose these additional assessments on this poster because we have recently discovered that, at least for the unconditional joint model, they are redundant. Put more mathematically, and using $\text{int}(x)$ as a shorthand notation for $P(x) \cup M(x)$, we require that:

$$K(X_c|x_{G \setminus c}) = K(X_c|\text{int}(x)) \text{ for all } c \in G \text{ and } x_{G \setminus c} \in \mathcal{X}_{G \setminus c}. \quad (1)$$

the right hand side of these equations being provided by the local models (I). In order to translate this into a property of a joint model $K(X_G)$, it is often assumed that for every $p(x_G) \in K(X_G)$, all events have strictly positive probability (Cozman 2000). Under this assumption, $K(X_G)$ can be conditioned by means of element-wise Bayes's rule (applying Bayes's rule to every $x \in G \setminus c$), thereby making it possible to impose Eq. (1).

We drop this positivity assumption by using an approach based on lower previsions, replacing Eq. (1) by the equivalent (BMT) statement that:

$$E_c(x_{G \setminus c}) = E_c(x_{G \setminus c}) \text{ for all } c \in G \text{ and } x_{G \setminus c} \in \mathcal{X}_{G \setminus c}. \quad (2)$$

where, again, the right hand side is provided by the local models (I). Since, without the positivity assumption, conditioning is not uniquely defined, we use a different method for making the conditional models in Eq. (2) consistent with the joint model E_G : we require them to be (strongly) coherent. We prove that, in our particular case, this is equivalent to requiring that

$$E_G(\mathbb{I}_{\{x_{G \setminus c}\}}|x_{G \setminus c}) = 0 \text{ and } E_G(\mathbb{I}_{\{x_{G \setminus c}\}}|x_{G \setminus c}) = 0 \quad (3)$$

for all $c \in G$, $x_{G \setminus c} \in \mathcal{X}_{G \setminus c}$ and $x \in \mathcal{F}(x)$. This formula is known as Generalised Bayes's Rule (GBR) and is equivalent to element-wise Bayes's rule if the positivity assumption is satisfied. It should therefore be clear that our approach is an extension of the one by Cozman (2000), coinciding with it under the positivity assumption.



Irrelevant natural extension (III)

The properties that we impose on our network (I, II) can be satisfied by multiple coherent lower previsions E_G on $\mathcal{F}(X_G)$. However, amongst these, there is a least one, the unique most conservative (pointwise) coherent lower prevision. We call it the irrelevant natural extension of the network and denote it by E^*_G . We show that E^*_G is the solution of the linear program (IV) and is the unique solution of the linear program (V) (see the next section).

We also prove the following simple characterization of E^*_G : it consists of all probability mass functions $p(X_G)$ on X_G for which for all $x \in G$ and $x_{G \setminus c} \in \mathcal{X}_{G \setminus c}$, there are a real number $\lambda \geq 0$ and a $p(X_c|x_{G \setminus c}) \in K(X_c|x_{G \setminus c})$ such that

$$\sum_{x \in \mathcal{X}_{G \setminus c}} p(x_{G \setminus c}) p(x_c|x_{G \setminus c}, x_{G \setminus c}) = \lambda p(X_c|x_{G \setminus c}),$$

where we use $M(x)$ to denote the set consisting of the descendants of the node x .

We believe that most of the marginalisation and graphoid properties that are presented on our other poster can be translated to the current framework. Combined with linear programming (IV), this might allow for efficient inference algorithms.

Definition by means of strong coherence

Description in terms of linear constraints

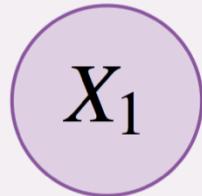
Independent natural extension for two binary variables



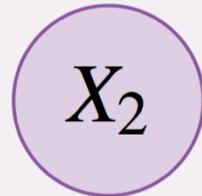
Independent natural extension for two binary variables

$$K^{\text{irr}}(X_G) = K(X_1) \otimes K(X_2)$$

$$K(X_1)$$



$$K(X_2)$$



$$\mathcal{X}_1 = \{h_1, t_1\}$$

$$\mathcal{X}_2 = \{h_2, t_2\}$$

$$\bar{p}(t_1)p(h_1, h_2) - \underline{p}(h_1)p(t_1, h_2) \geq 0$$

$$-\underline{p}(t_1)p(h_1, h_2) + \bar{p}(h_1)p(t_1, h_2) \geq 0$$

$$\bar{p}(t_1)p(h_1, t_2) - \underline{p}(h_1)p(t_1, t_2) \geq 0$$

$$-\underline{p}(t_1)p(h_1, t_2) + \bar{p}(h_1)p(t_1, t_2) \geq 0$$

$$\bar{p}(t_2)p(h_1, h_2) - \underline{p}(h_2)p(h_1, t_2) \geq 0$$

$$-\underline{p}(t_2)p(h_1, h_2) + \bar{p}(h_2)p(h_1, t_2) \geq 0$$

$$\bar{p}(t_2)p(t_1, h_2) - \underline{p}(h_2)p(t_1, t_2) \geq 0$$

$$-\underline{p}(t_2)p(t_1, h_2) + \bar{p}(h_2)p(t_1, t_2) \geq 0$$

$$K(X_i) = \left\{ p \in \Sigma_{\mathcal{X}_i} : p(h_i) \in [\underline{p}(h_i), \bar{p}(h_i)] \right\}$$

$$\bar{p}(t_i) := 1 - \underline{p}(h_i)$$

$$\underline{p}(t_i) := 1 - \bar{p}(h_i)$$

Independent natural extension for two binary variables

Analytical expressions for the extreme points

	$p(h_1, h_2) \Sigma$	$p(h_1, t_2) \Sigma$	$p(t_1, h_2) \Sigma$	$p(t_1, t_2) \Sigma$	Σ
p_{S1}	$\underline{p}(h_1)\underline{p}(h_2)$	$\underline{p}(h_1)\bar{p}(t_2)$	$\bar{p}(t_1)\underline{p}(h_2)$	$\bar{p}(t_1)\bar{p}(t_2)$	1
p_{S2}	$\underline{p}(h_1)\bar{p}(h_2)$	$\underline{p}(h_1)\underline{p}(t_2)$	$\bar{p}(t_1)\bar{p}(h_2)$	$\bar{p}(t_1)\underline{p}(t_2)$	1
p_{S3}	$\bar{p}(h_1)\underline{p}(h_2)$	$\bar{p}(h_1)\bar{p}(t_2)$	$\underline{p}(t_1)\underline{p}(h_2)$	$\underline{p}(t_1)\bar{p}(t_2)$	1
p_{S4}	$\bar{p}(h_1)\bar{p}(h_2)$	$\bar{p}(h_1)\underline{p}(t_2)$	$\underline{p}(t_1)\bar{p}(h_2)$	$\underline{p}(t_1)\underline{p}(t_2)$	1
p_{A1}	$\underline{p}(h_1)\bar{p}(h_1)\underline{p}(h_2)$	$\underline{p}(h_1)\bar{p}(h_1)\bar{p}(t_2)$	$\bar{p}(t_1)\bar{p}(h_1)\underline{p}(h_2)$	$\underline{p}(h_1)\underline{p}(t_1)\bar{p}(t_2)$	$\underline{p}(h_1)\bar{p}(t_2) + \bar{p}(h_1)\underline{p}(h_2)$
p_{A2}	$\underline{p}(h_1)\bar{p}(h_1)\bar{p}(h_2)$	$\underline{p}(h_1)\bar{p}(h_1)\underline{p}(t_2)$	$\underline{p}(h_1)\underline{p}(t_1)\bar{p}(h_2)$	$\bar{p}(t_1)\bar{p}(h_1)\underline{p}(t_2)$	$\underline{p}(h_1)\bar{p}(h_2) + \bar{p}(h_1)\underline{p}(t_2)$
p_{A3}	$\bar{p}(h_1)\bar{p}(t_1)\underline{p}(h_2)$	$\underline{p}(t_1)\underline{p}(h_1)\bar{p}(t_2)$	$\underline{p}(t_1)\bar{p}(t_1)\underline{p}(h_2)$	$\underline{p}(t_1)\bar{p}(t_1)\bar{p}(t_2)$	$\underline{p}(t_1)\bar{p}(t_2) + \bar{p}(t_1)\underline{p}(h_2)$
p_{A4}	$\underline{p}(t_1)\underline{p}(h_1)\bar{p}(h_2)$	$\bar{p}(h_1)\bar{p}(t_1)\underline{p}(t_2)$	$\underline{p}(t_1)\bar{p}(t_1)\bar{p}(h_2)$	$\underline{p}(t_1)\bar{p}(t_1)\underline{p}(t_2)$	$\underline{p}(t_1)\bar{p}(h_2) + \bar{p}(t_1)\underline{p}(t_2)$
p_{B1}	$\underline{p}(h_2)\bar{p}(h_2)\underline{p}(h_1)$	$\bar{p}(t_2)\bar{p}(h_2)\underline{p}(h_1)$	$\underline{p}(h_2)\bar{p}(h_2)\bar{p}(t_1)$	$\underline{p}(h_2)\underline{p}(t_2)\bar{p}(t_1)$	$\underline{p}(h_2)\bar{p}(t_1) + \bar{p}(h_2)\underline{p}(h_1)$
p_{B2}	$\bar{p}(h_2)\bar{p}(t_2)\underline{p}(h_1)$	$\underline{p}(t_2)\bar{p}(t_2)\underline{p}(h_1)$	$\underline{p}(t_2)\underline{p}(h_2)\bar{p}(t_1)$	$\underline{p}(t_2)\bar{p}(t_2)\bar{p}(t_1)$	$\underline{p}(t_2)\bar{p}(t_1) + \bar{p}(t_2)\underline{p}(h_1)$
p_{B3}	$\underline{p}(h_2)\bar{p}(h_2)\bar{p}(h_1)$	$\underline{p}(h_2)\underline{p}(t_2)\bar{p}(h_1)$	$\underline{p}(h_2)\bar{p}(h_2)\underline{p}(t_1)$	$\bar{p}(t_2)\bar{p}(h_2)\underline{p}(t_1)$	$\underline{p}(h_2)\bar{p}(h_1) + \bar{p}(h_2)\underline{p}(t_1)$
p_{B4}	$\underline{p}(t_2)\underline{p}(h_2)\bar{p}(h_1)$	$\underline{p}(t_2)\bar{p}(t_2)\bar{p}(h_1)$	$\bar{p}(h_2)\bar{p}(t_2)\underline{p}(t_1)$	$\underline{p}(t_2)\bar{p}(t_2)\underline{p}(t_1)$	$\underline{p}(t_2)\bar{p}(h_1) + \bar{p}(t_2)\underline{p}(t_1)$

Independent natural extension for two binary variables

Analytical expressions for the extreme points

	$p(h_1, h_2) \Sigma$	$p(h_1, t_2) \Sigma$	$p(t_1, h_2) \Sigma$	$p(t_1, t_2) \Sigma$	Σ
p_{S1}	$\underline{p}(h_1)\underline{p}(h_2)$	$\underline{p}(h_1)\bar{p}(t_2)$	$\bar{p}(t_1)\underline{p}(h_2)$	$\underline{p}(h_1)\bar{p}(t_1)\bar{p}(h_2)\bar{p}(t_2)$	$\underline{p}(h_1)\bar{p}(t_1)\bar{p}(h_2)\bar{p}(t_2)$
p_{S2}	$\underline{p}(h_1)\bar{p}(h_2)$	$\underline{p}(h_1)\bar{p}(t_2)$	$\bar{p}(t_1)\underline{p}(h_2)$	$\underline{p}(h_1)\bar{p}(t_1)\bar{p}(h_2)\bar{p}(t_2)$	$\underline{p}(h_1)\bar{p}(t_1)\bar{p}(h_2)\bar{p}(t_2)$
p_{S3}	$\bar{p}(h_1)\underline{p}(h_2)$	$\underline{p}(h_1)\bar{p}(t_2)$	$\bar{p}(t_1)\underline{p}(h_2)$	$\underline{p}(h_1)\bar{p}(t_1)\bar{p}(h_2)\bar{p}(t_2)$	$\underline{p}(h_1)\bar{p}(t_1)\bar{p}(h_2)\bar{p}(t_2)$
p_{S4}	$\bar{p}(h_1)\underline{p}(h_2)$	$\underline{p}(h_1)\bar{p}(t_2)$	$\bar{p}(t_1)\underline{p}(h_2)$	$\underline{p}(h_1)\bar{p}(t_1)\bar{p}(h_2)\bar{p}(t_2)$	$\underline{p}(h_1)\bar{p}(t_1)\bar{p}(h_2)\bar{p}(t_2)$
p_{A1}	$\underline{p}(h_1)\bar{p}(h_2)$	$\underline{p}(h_1)\bar{p}(t_2)$	$\bar{p}(t_1)\underline{p}(h_2)$	$\underline{p}(h_1)\bar{p}(t_1)\bar{p}(h_2)\bar{p}(t_2)$	$\underline{p}(h_1)\bar{p}(t_1)\bar{p}(h_2)\bar{p}(t_2)$
p_{A2}	$\underline{p}(h_1)\bar{p}(h_2)$	$\bar{p}(h_1)\underline{p}(t_2)$	$\underline{p}(t_1)\bar{p}(h_2)$	$\underline{p}(h_1)\bar{p}(t_1)\bar{p}(h_2)\bar{p}(t_2)$	$\underline{p}(h_1)\bar{p}(t_1)\bar{p}(h_2)\bar{p}(t_2)$
p_{A3}	$\bar{p}(h_1)\underline{p}(h_2)$	$\underline{p}(h_1)\bar{p}(t_2)$	$\underline{p}(t_1)\bar{p}(h_2)$	$\underline{p}(h_1)\bar{p}(t_1)\bar{p}(h_2)\bar{p}(t_2)$	$\underline{p}(h_1)\bar{p}(t_1)\bar{p}(h_2)\bar{p}(t_2)$
p_{A4}	$\underline{p}(t_1)\bar{p}(h_1)\bar{p}(h_2)$	$\bar{p}(h_1)\underline{p}(t_2)$	$\underline{p}(t_1)\bar{p}(h_1)\bar{p}(h_2)$	$\underline{p}(t_1)\bar{p}(h_1)\bar{p}(h_2)\bar{p}(t_2)$	$\underline{p}(t_1)\bar{p}(h_1)\bar{p}(h_2)\bar{p}(t_2)$
p_{B1}	$\underline{p}(h_2)\bar{p}(h_1)\underline{p}(h_2)$	$\bar{p}(t_2)\bar{p}(h_2)\underline{p}(h_1)$	$\underline{p}(h_2)\bar{p}(h_1)\bar{p}(h_2)$	$\underline{p}(h_2)\bar{p}(h_1)\bar{p}(h_2)\bar{p}(t_1)$	$\underline{p}(h_2)\bar{p}(h_1)\bar{p}(h_2)\bar{p}(t_1)$
p_{B2}	$\bar{p}(h_2)\bar{p}(h_1)\underline{p}(h_2)$	$\underline{p}(t_2)\bar{p}(h_2)\underline{p}(h_1)$	$\underline{p}(h_2)\bar{p}(h_1)\bar{p}(h_2)$	$\underline{p}(t_2)\bar{p}(h_1)\bar{p}(h_2)\bar{p}(t_1)$	$\underline{p}(t_2)\bar{p}(h_1)\bar{p}(h_2)\bar{p}(t_1)$
p_{B3}	$\underline{p}(h_2)\bar{p}(h_1)\bar{p}(h_2)$	$\underline{p}(h_2)\bar{p}(h_1)\bar{p}(h_2)$	$\bar{p}(h_2)\bar{p}(h_1)\underline{p}(h_2)$	$\bar{p}(t_2)\bar{p}(h_1)\bar{p}(h_2)\bar{p}(t_1)$	$\bar{p}(h_2)\bar{p}(h_1)\bar{p}(h_2)\bar{p}(t_1)$
p_{B4}	$\bar{p}(h_2)\bar{p}(h_1)\bar{p}(h_2)$	$\bar{p}(h_2)\bar{p}(h_1)\bar{p}(h_2)$	$\bar{p}(h_2)\bar{p}(h_1)\bar{p}(h_2)$	$\bar{p}(t_2)\bar{p}(h_1)\bar{p}(h_2)\bar{p}(t_1)$	$\bar{p}(h_2)\bar{p}(h_1)\bar{p}(h_2)\bar{p}(t_1)$

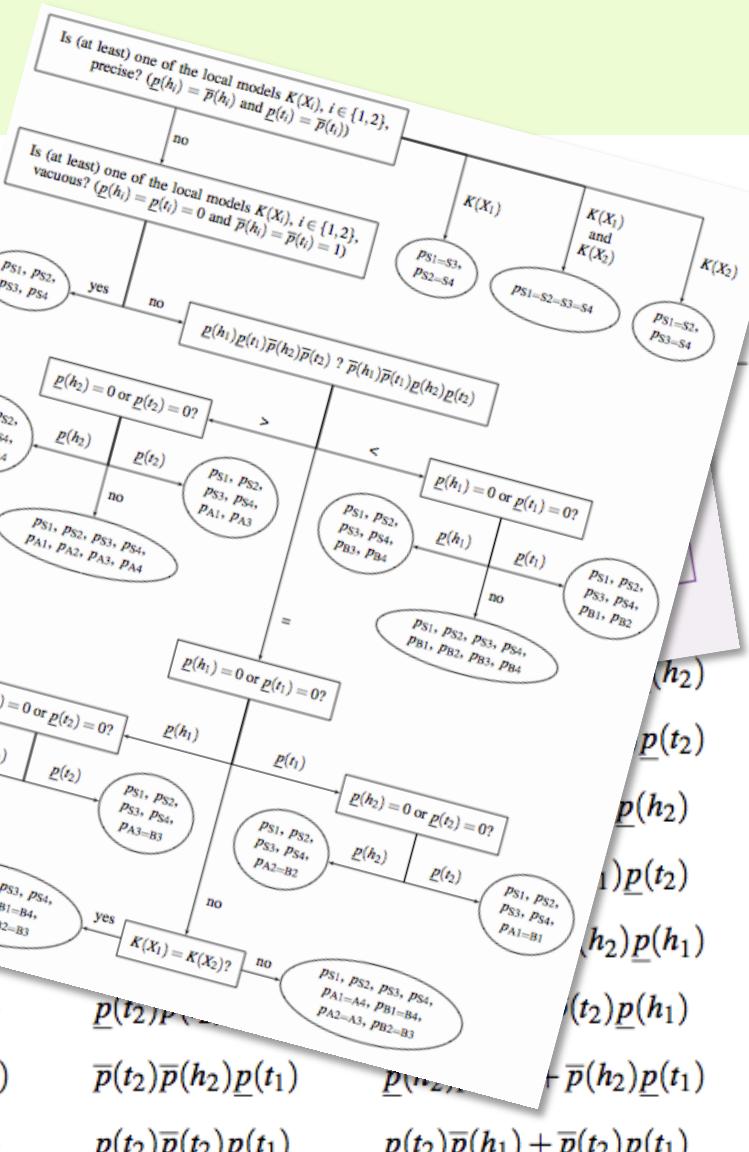
Annotations:

- Arrows point from the first four rows to boxes containing sets of parameters: $p_{A1}, p_{A2}, p_{A3}, p_{A4}$; $p_{B1}, p_{B2}, p_{B3}, p_{B4}$; and $p_{A1}=A4, p_{B1}=B4, p_{A2}=A3, p_{B2}=B3$.
- A box labeled "remove those for which $\Sigma = 0$ " has arrows pointing to the fifth row and the last four rows.
- A box on the right lists the remaining parameters: $p_{S1}, p_{S2}, p_{S3}, p_{S4}$.

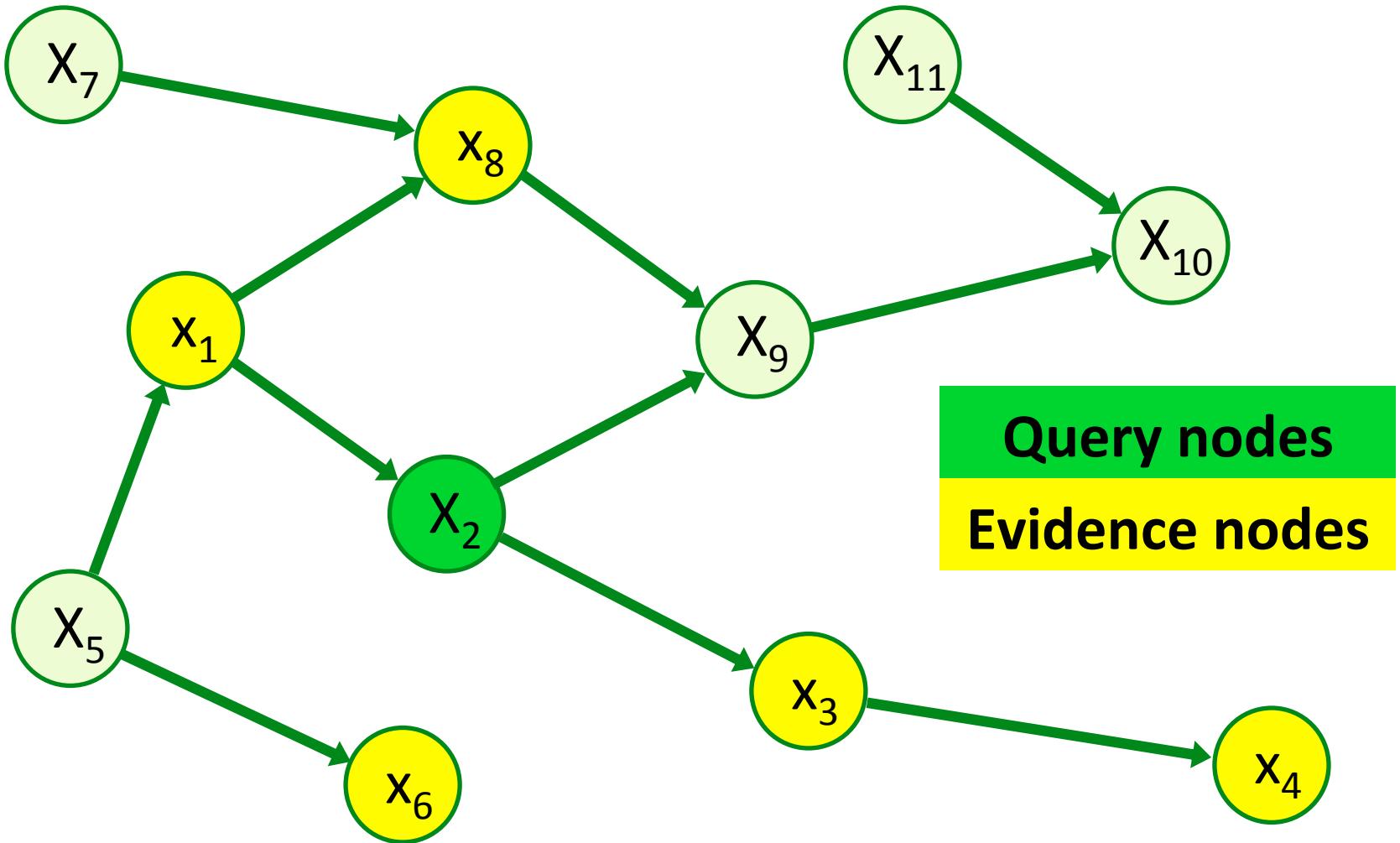
Independent natural extension for two binary variables

Analytical expressions for the ϵ

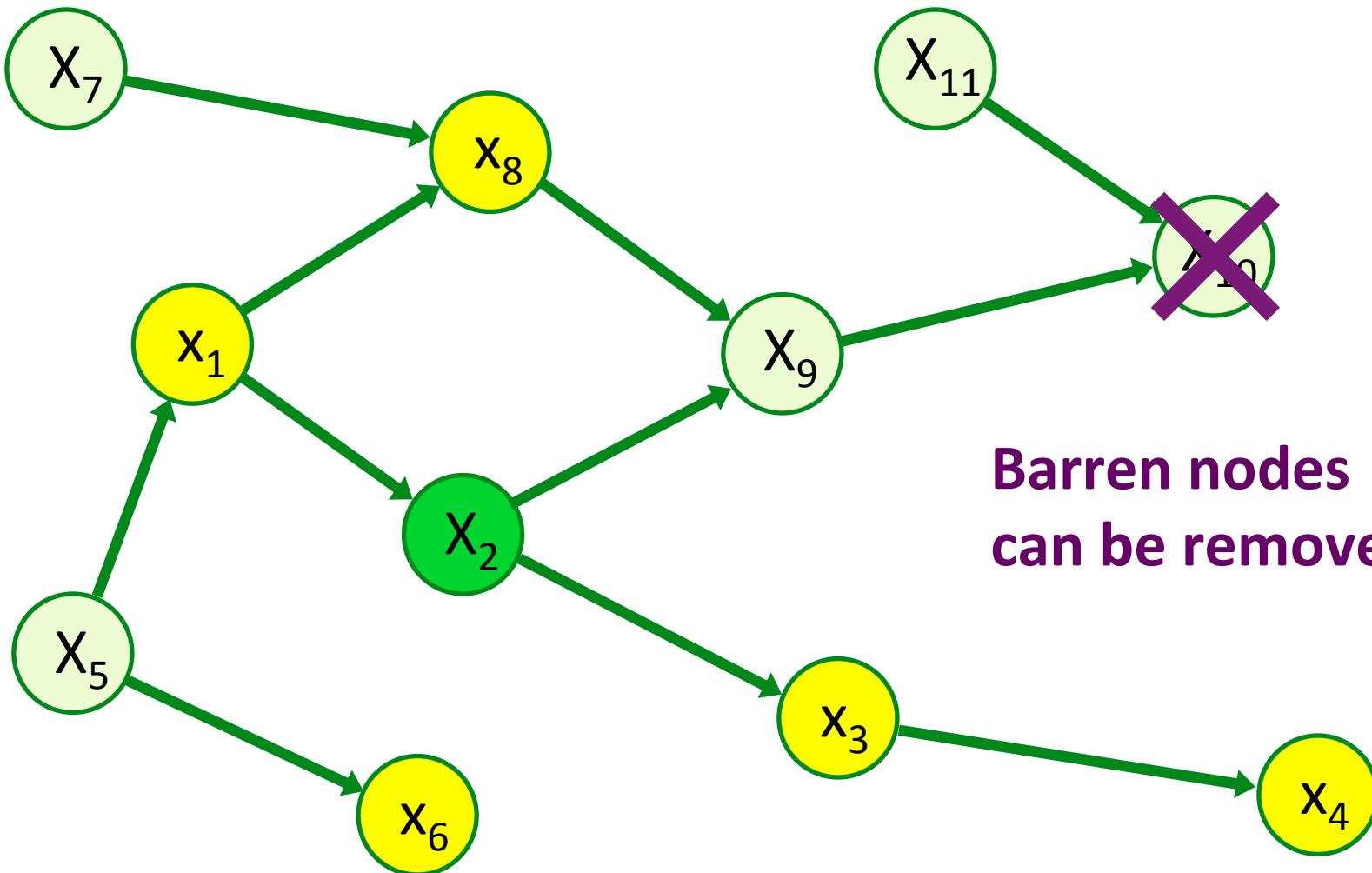
	$p(h_1, h_2) \Sigma$	$p(h_1, t_2) \Sigma$	$p(t_1, h_2) \Sigma$	
p_{S1}	$\underline{p}(h_1)\underline{p}(h_2)$	$\underline{p}(h_1)\bar{p}(t_2)$	$\bar{p}(t_1)p(h_2)$	
p_{S2}	$\underline{p}(h_1)\bar{p}(h_2)$	$\underline{p}(h_1)\bar{p}(t_2)$	$\bar{p}(t_1)\underline{p}(h_2)$	
p_{S3}	$\bar{p}(h_1)\underline{p}(t_1)\bar{p}(h_2)\bar{p}(t_2)$			
p_{S4}	$\bar{p}(h_1)\underline{p}(t_1)\bar{p}(h_2)\bar{p}(t_2)$			
p_{A1}	$\underline{p}(h_1)\underline{p}(h_2)$			
p_{A2}	$\underline{p}(h_1)\bar{p}(t_1)\bar{p}(h_2)\underline{p}(t_2)$			
p_{A3}	$\bar{p}(h_1)\underline{p}(t_1)\bar{p}(h_2)\underline{p}(t_2)$			
p_{A4}	$\underline{p}(t_1)\underline{p}(h_1)\underline{p}(t_2)$	$\bar{p}(h_1)\bar{p}(t_1)\underline{p}(t_2)$	$\underline{p}(t_1)\bar{p}(h_2)\underline{p}(t_2)$	
p_{B1}	$\underline{p}(h_2)\bar{p}(h_1)\underline{p}(h_1)$	$\bar{p}(t_2)\bar{p}(h_2)\underline{p}(h_1)$	$\underline{p}(h_2)\underline{p}(t_2)\bar{p}(h_1)$	$\underline{p}(h_2)\underline{p}(h_1)$
p_{B2}	$\bar{p}(h_2)\bar{p}(t_2)\underline{p}(h_1)$	$\underline{p}(t_2)\bar{p}(t_2)\underline{p}(h_1)$	$\underline{p}(t_2)\underline{p}(h_2)\bar{p}(t_1)$	$\bar{p}(t_2)\underline{p}(h_2)\underline{p}(t_1)$
p_{B3}	$\underline{p}(h_2)\bar{p}(h_2)\bar{p}(h_1)$	$\underline{p}(h_2)\underline{p}(t_2)\bar{p}(h_1)$	$\underline{p}(h_2)\bar{p}(h_2)\underline{p}(t_1)$	$\bar{p}(t_2)\bar{p}(h_2)\underline{p}(t_1)$
p_{B4}	$\underline{p}(t_2)\underline{p}(h_2)\bar{p}(h_1)$	$\underline{p}(t_2)\bar{p}(t_2)\bar{p}(h_1)$	$\bar{p}(h_2)\bar{p}(t_2)\underline{p}(t_1)$	$\underline{p}(t_2)\bar{p}(h_1) + \bar{p}(t_2)\underline{p}(t_1)$



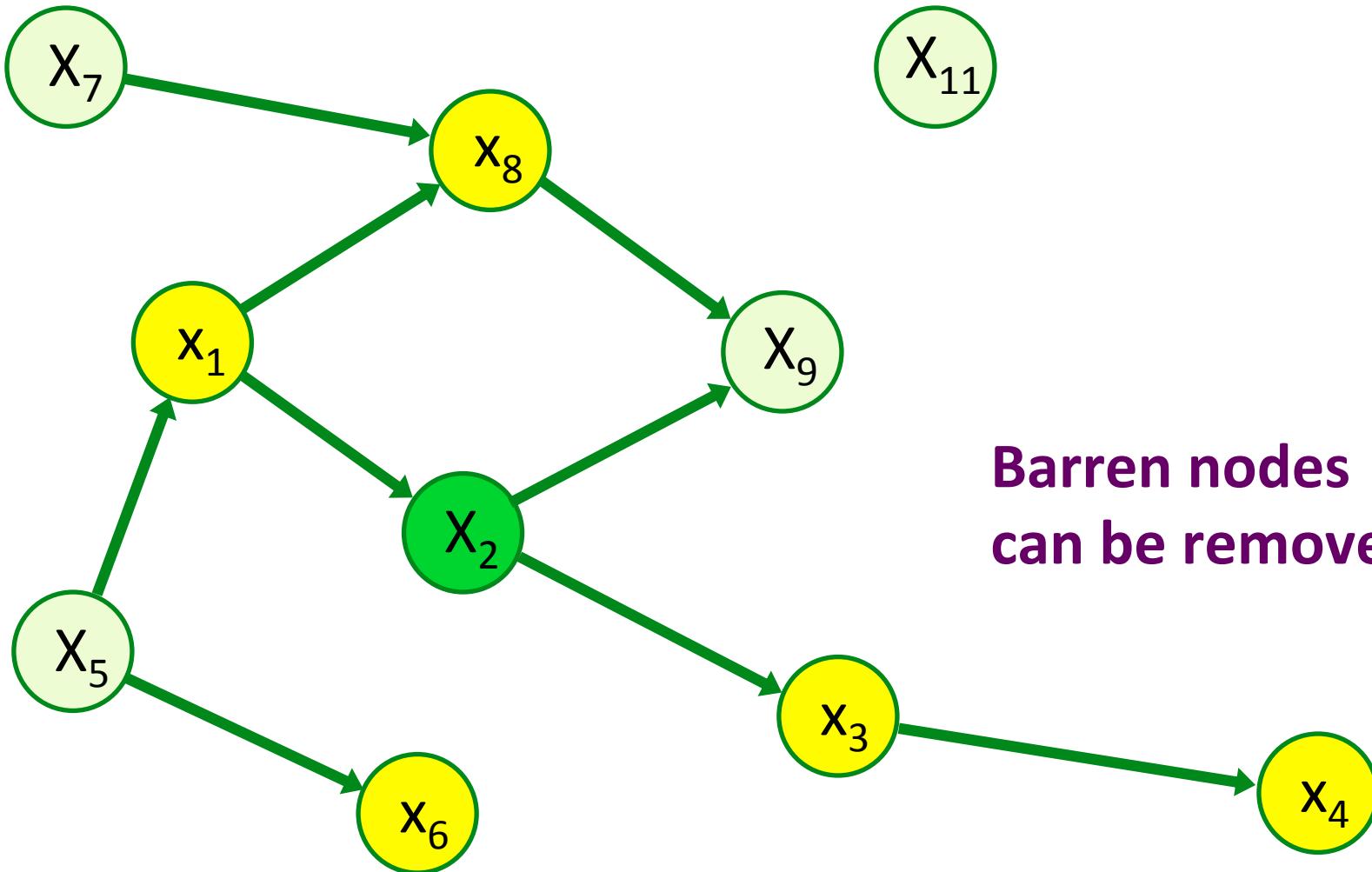
Bayesian networks: useful properties



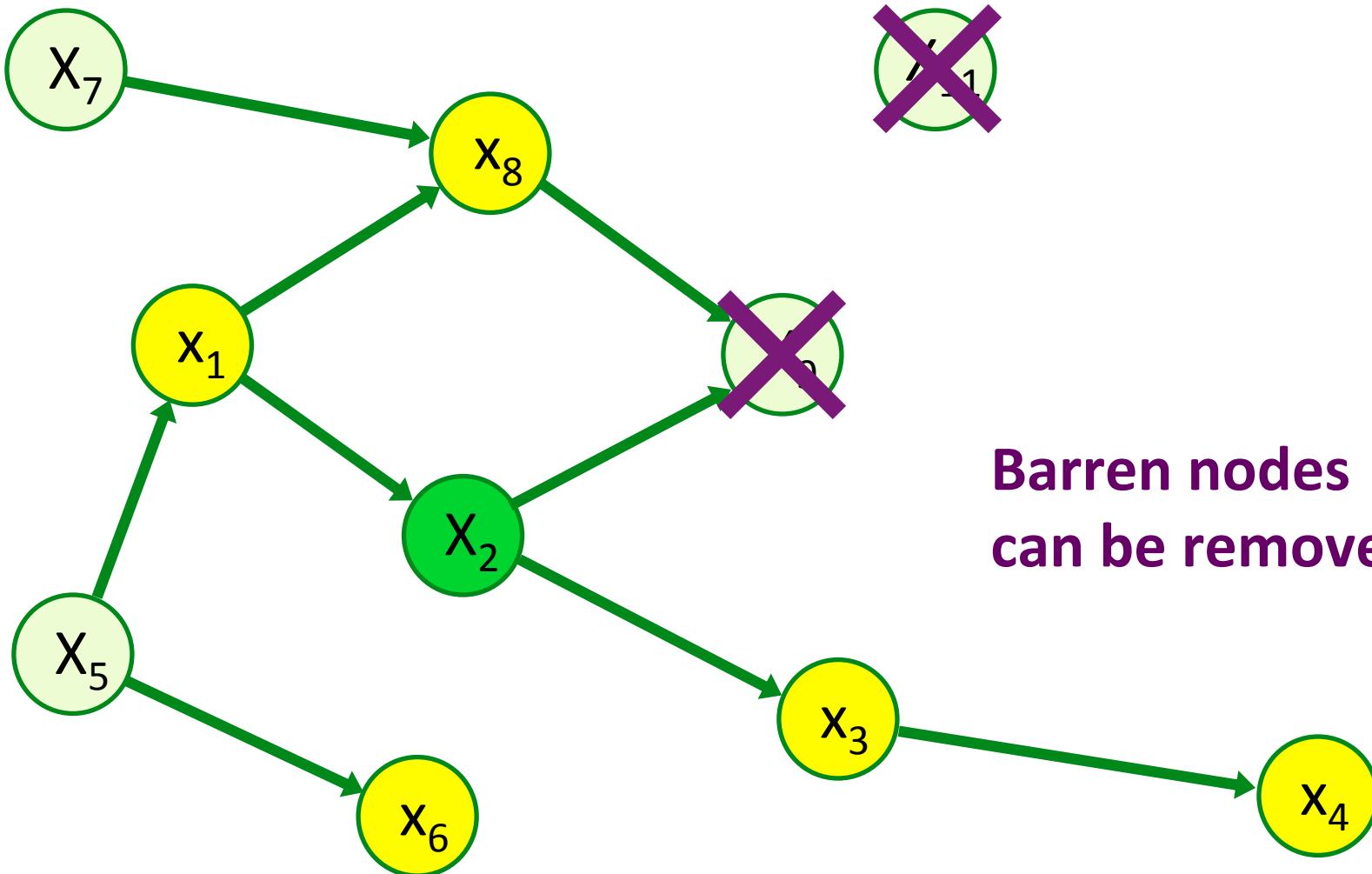
Bayesian networks: useful properties



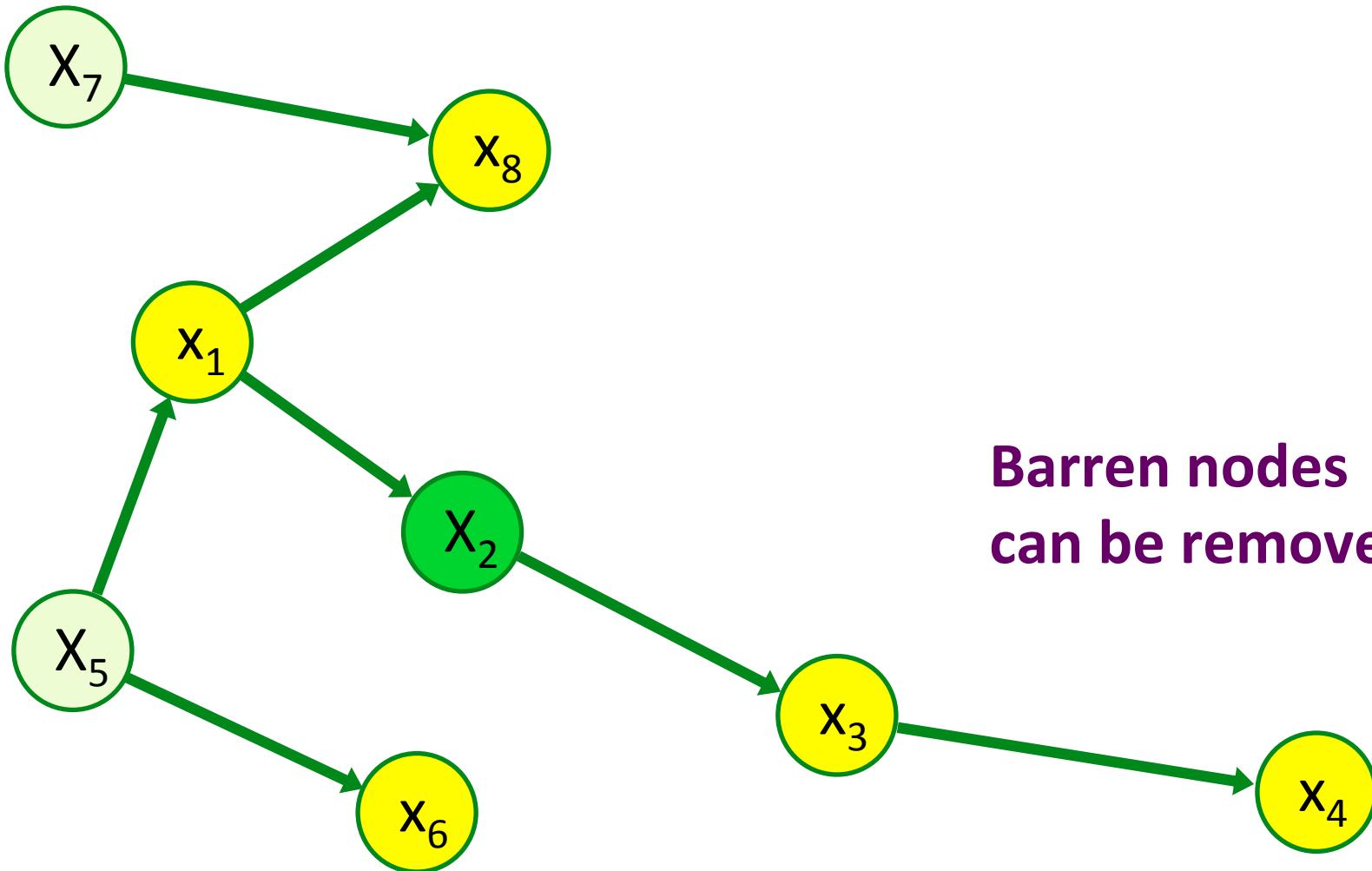
Bayesian networks: useful properties



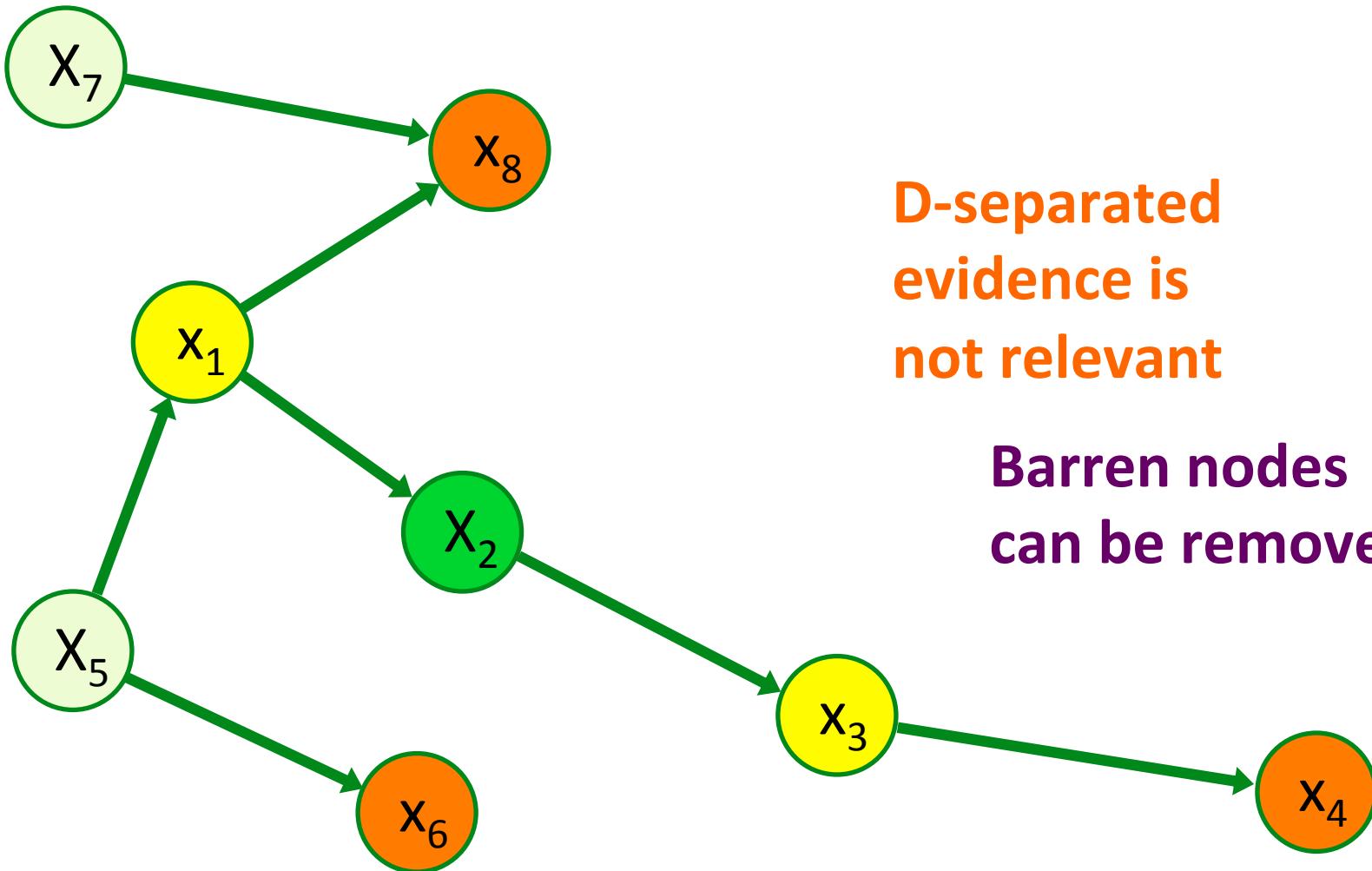
Bayesian networks: useful properties



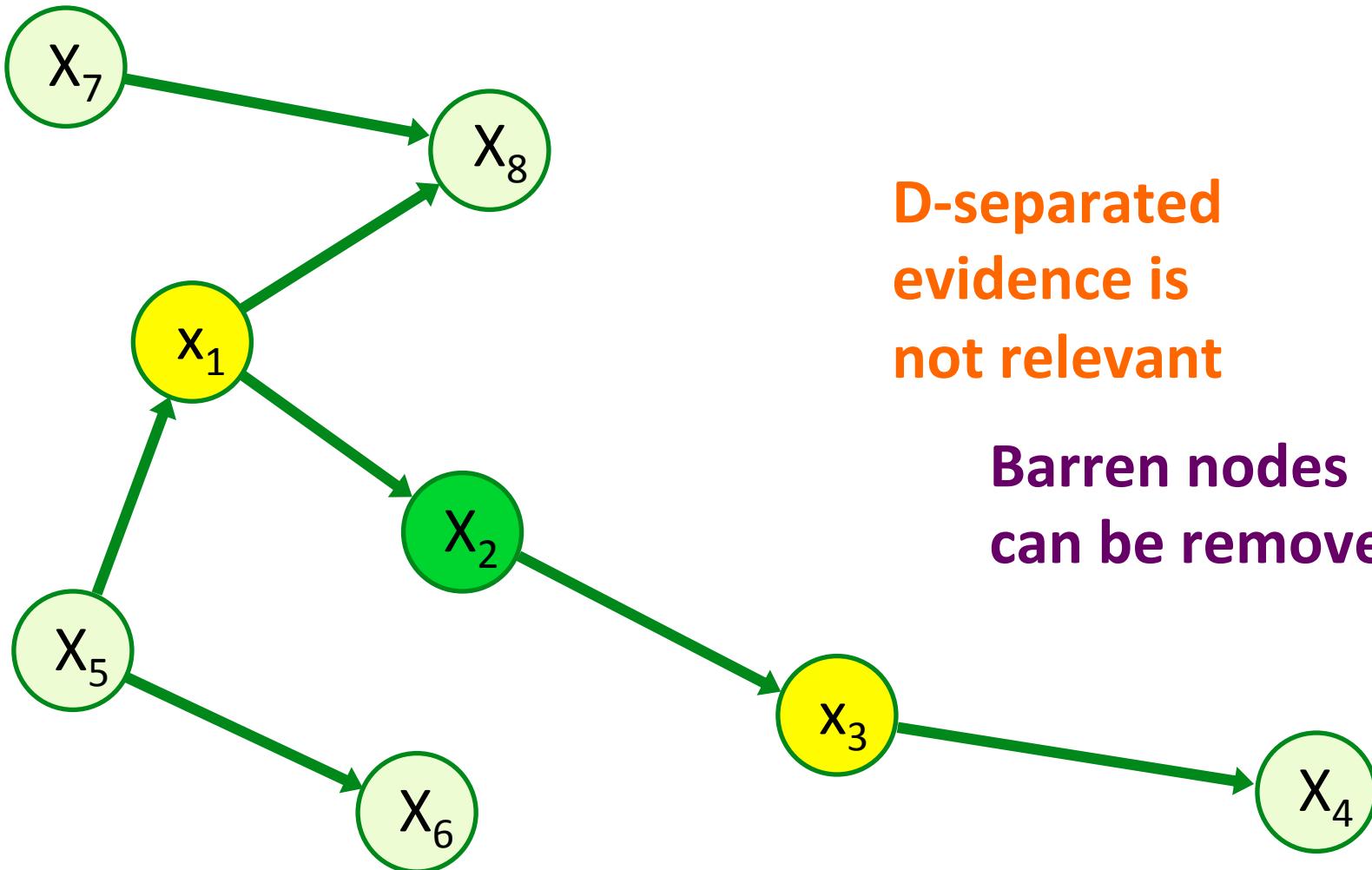
Bayesian networks: useful properties



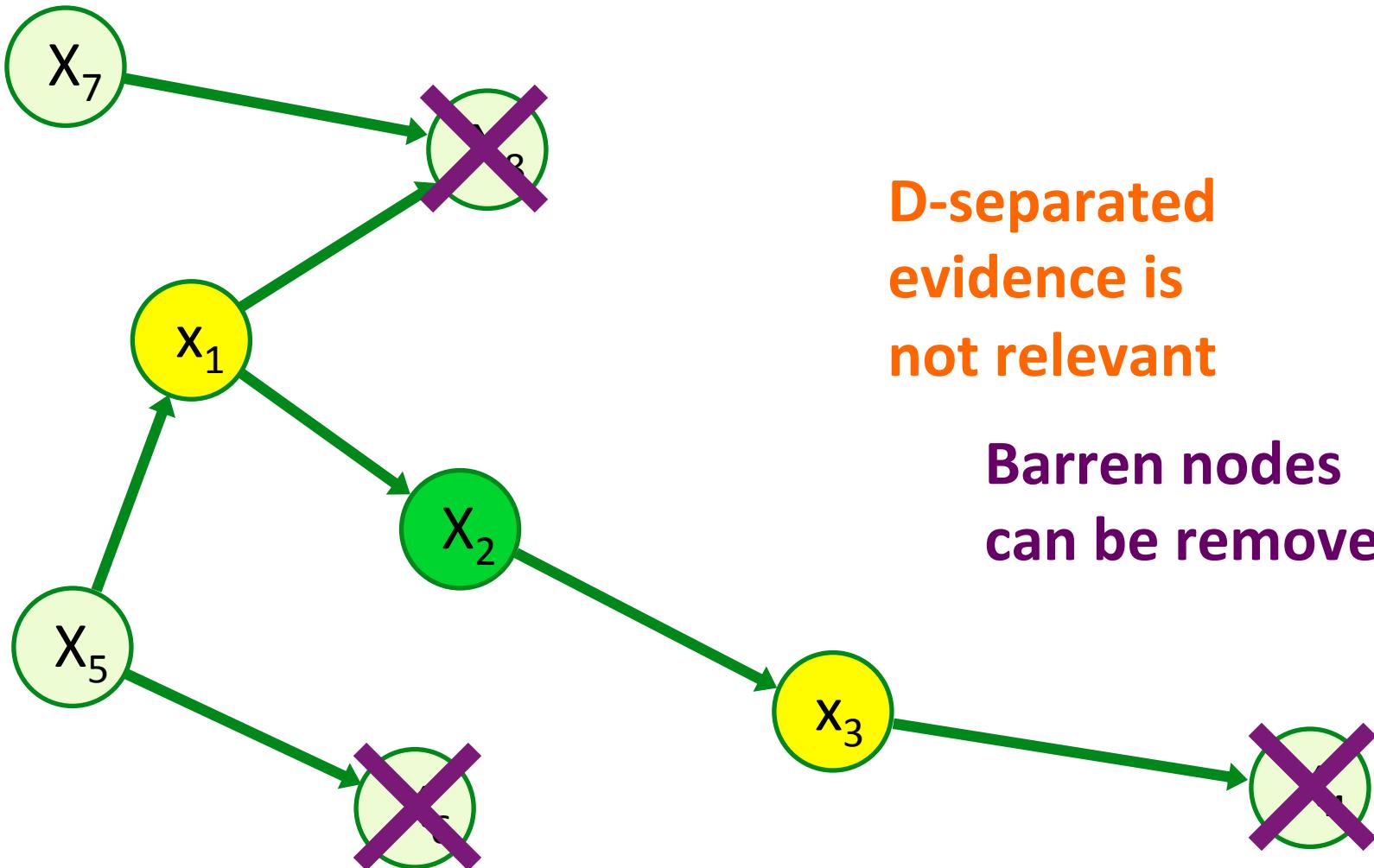
Bayesian networks: useful properties



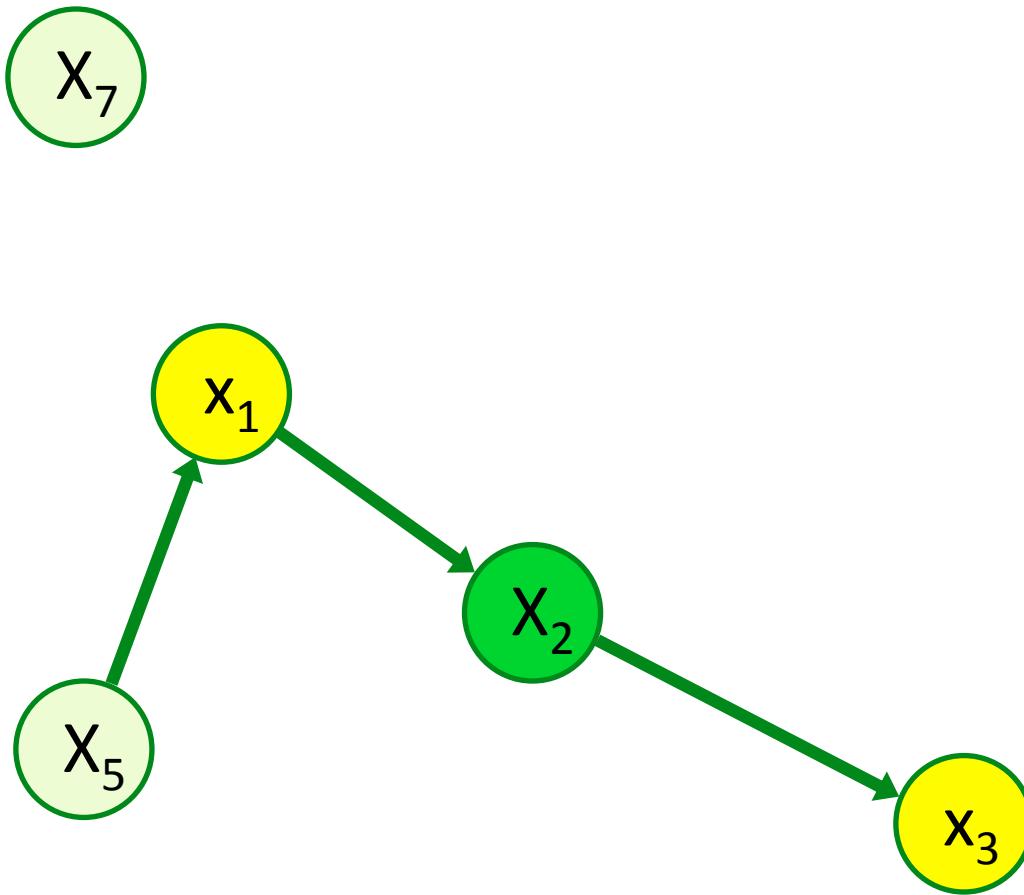
Bayesian networks: useful properties



Bayesian networks: useful properties



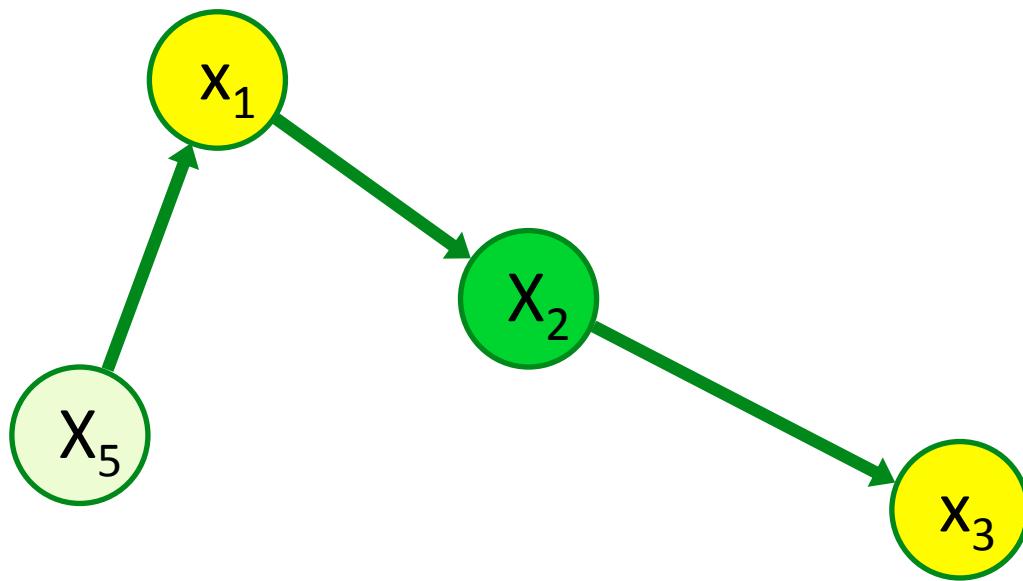
Bayesian networks: useful properties



D-separated
evidence is
not relevant

Barren nodes
can be removed

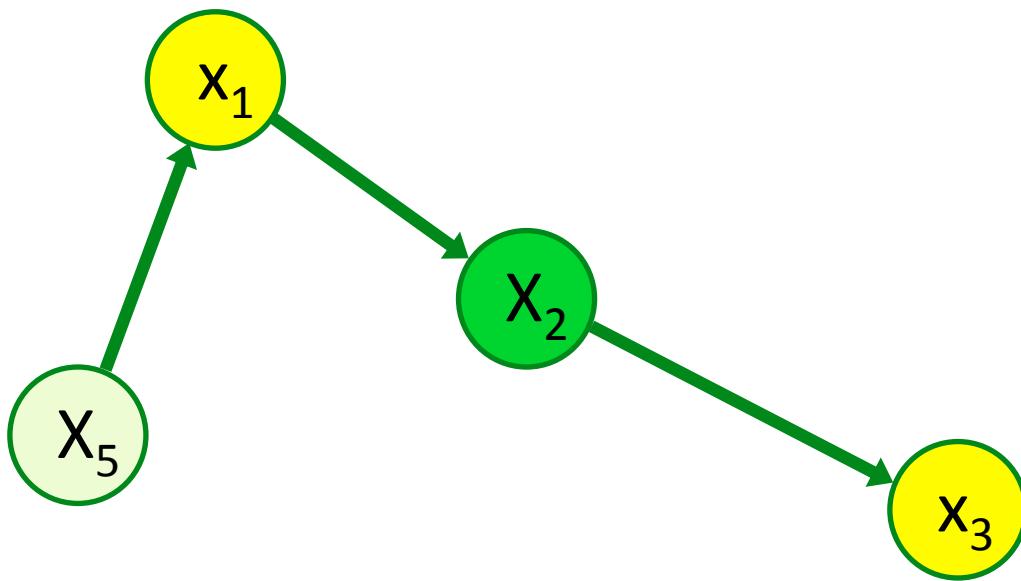
Bayesian networks: useful properties



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Bayesian networks: useful properties

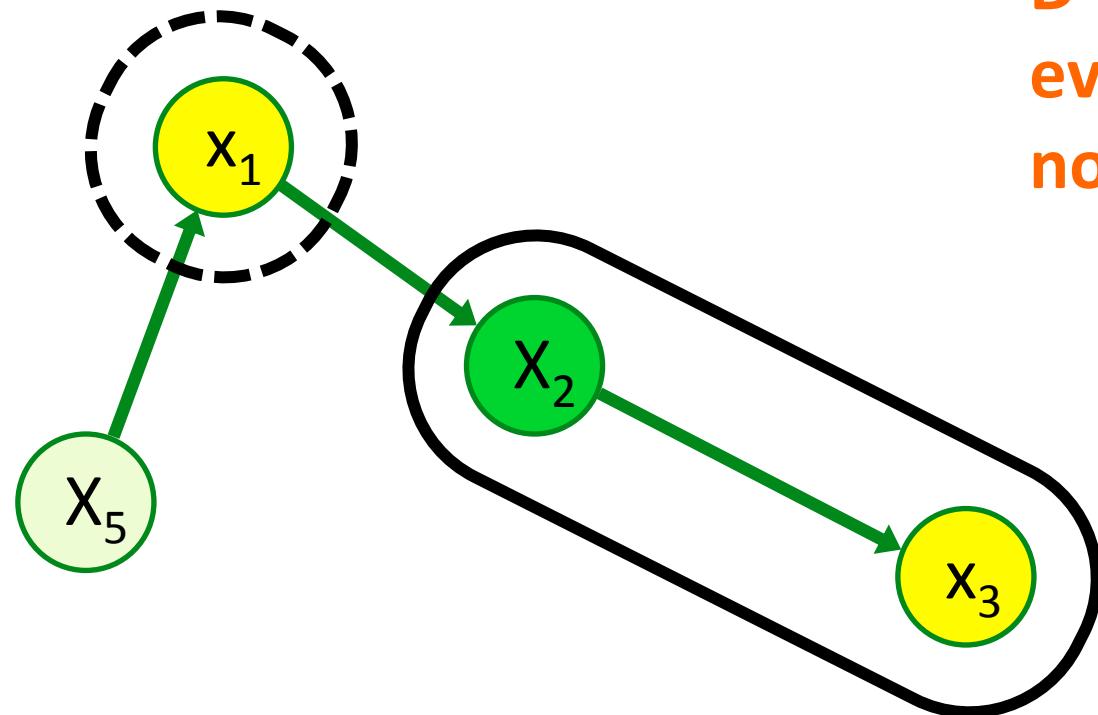


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Bayesian networks: useful properties

Conditional
marginalisation
properties

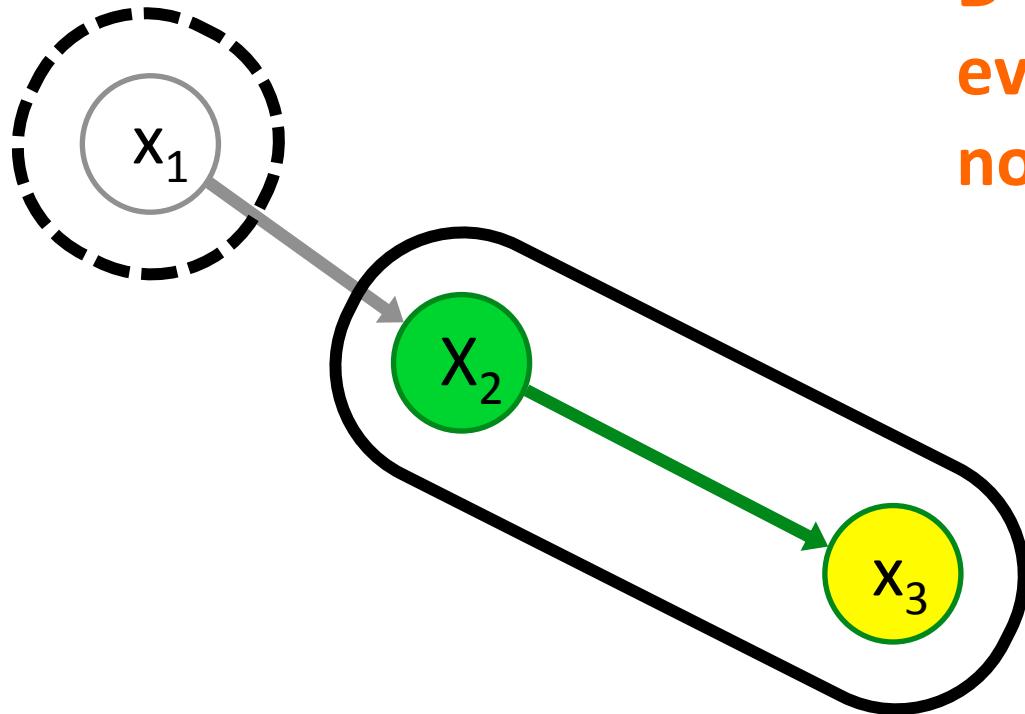


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Bayesian networks: useful properties

Conditional
marginalisation
properties

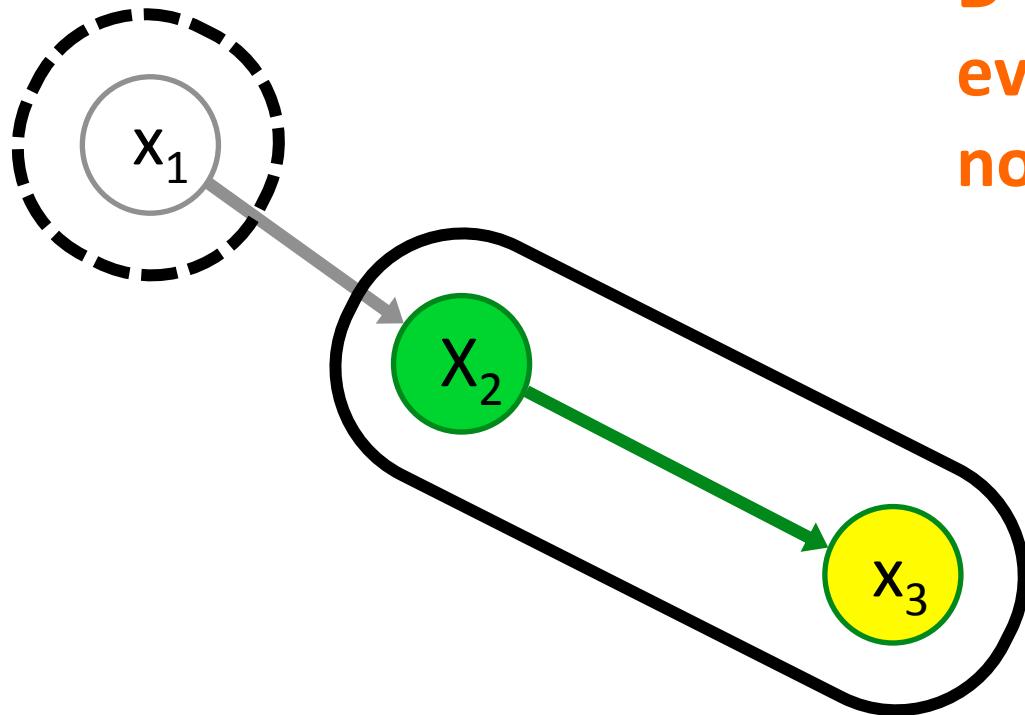


D-separated
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Credal networks: useful properties ?

Conditional
marginalisation
properties



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Credal networks: useful properties ?



Conditional
marginalisation
properties



D-separated
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Credal networks under
epistemic irrelevance

Credal networks: useful properties ?

?

Conditional
marginalisation
properties

D-separated
evidence is
not relevant

Barren nodes
can be removed

~~graphoid
axioms~~



?



Credal networks under
epistemic irrelevance

Credal networks: a joint model

Credal networks under epistemic irrelevance using sets of desirable gambles

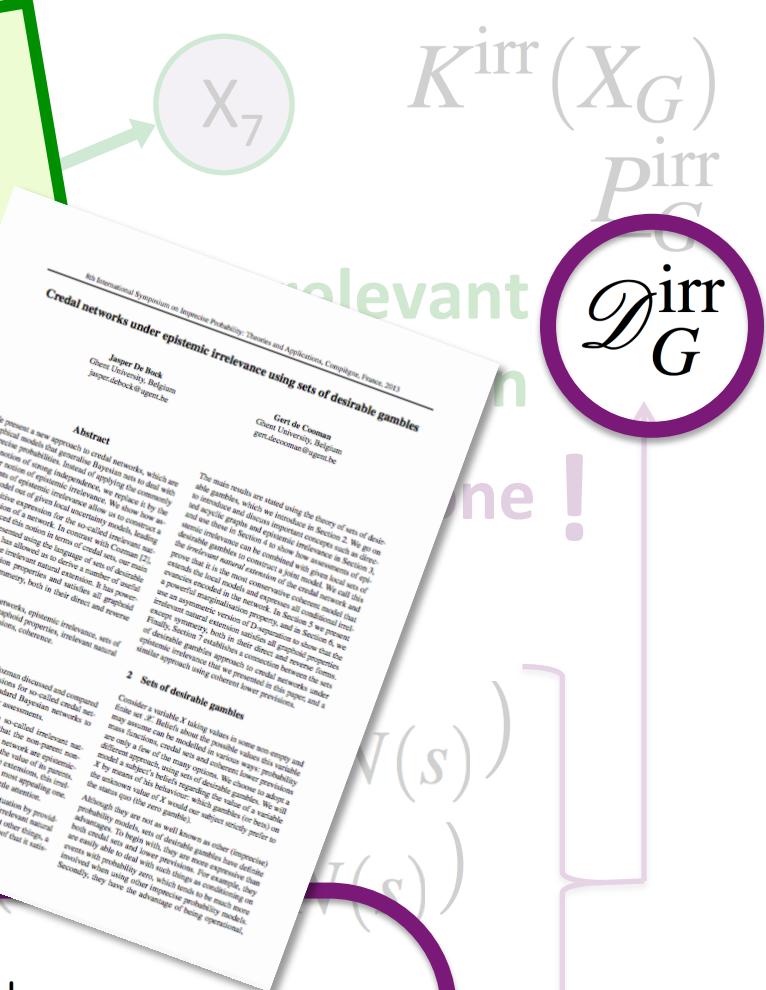
The m

X_1

$$K(X_s | x_{P(s)}) =$$

$$P_s(\cdot | x_{P(s)}) =$$

$$\mathcal{D}_s | x_{P(s)} = \mathcal{D}_s | x_{P(s) \cup N(s)}$$



Credal networks under epistemic irrelevance using sets of desirable gambles

Abstract of the paper

We present a new approach to credal networks, which are graphical models that generalise Bayesian nets to deal with imprecise probabilities. Instead of applying the commonly used notion of strong independence, we replace it by the weaker notion of epistemic irrelevance. We show how sets of desirable gambles allow us to construct local models of given local uncertainty models (\mathcal{G}), leading to an intuitive expression for the so-called irrelevant natural extension (III) of a network. In contrast with Cozman (2000) who introduced this notion in terms of credal sets, our main results are presented using the language of sets of desirable gambles (SDG). This has allowed us to derive a number of useful properties of the irrelevant natural extension. It has powerful marginalisation properties (IV) and satisfies all graphoid properties but symmetry, both in their direct and reverse forms (V & VI).

Sets of desirable gambles (SDG)

We will model a subject's beliefs about the value that a variable X , assumes in some set \mathcal{X} , by means of his behaviour, which generates (measurable) maps f on \mathcal{X} . f does not have to be the ratio or the ratio quo. This results in a set of desirable gambles $\mathcal{G} = \mathcal{G}(\mathcal{X})$, where $\mathcal{G}(\mathcal{X})$ is the set of all gambles on \mathcal{X} . \mathcal{G} is called coherent if it satisfies the rationality requirements D1—D4 for all $f_1, f_2 \in \mathcal{G}(\mathcal{X})$ and all real $\lambda > 0$:

$$\Rightarrow f_1 + f_2 \in \mathcal{G}$$

Although they are not as well known as other (imprecise) probability models, sets of desirable gambles have definite advantages. To give a few examples: they are operational, are easily able to deal with conditioning on events with probability zero, allow for intuitive geometrically flavoured proofs and are more expressive than both credal sets and lower previsions (see our papers or the second poster for credal networks under epistemic irrelevance that use these alternative models).

Local uncertainty models (I)

With every node x of a finite directed acyclic graph (DAG), we associate a variable X_x taking values in some finite, non-empty set \mathcal{X}_x . The set of all nodes is denoted by G . For every subset $S \subseteq G$, the joint variable X_S takes values in $\mathcal{X}_S := \times_{x \in S} \mathcal{X}_x$. For every $i \in G$, we denote by $P(i)$ the set consisting of the parent nodes of i . Similar to what is done in classical Bayesian networks, we attach local uncertainty models to the nodes of the network, conditioned on their parents. For all $s \in G$ and every instantiation $x_{P(s)} \in \mathcal{X}_{P(s)}$, we require a coherent set $\mathcal{G}_{s, P(s)}$ of desirable gambles (SDG) on \mathcal{X}_s .

AD-separation (V)

Consider any path x_1, \dots, x_n in G , with $n \geq 1$. We say that this path is **blocked** by a set of nodes $C \subseteq G$ whenever at least one of the following four conditions holds:

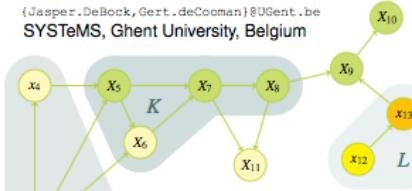
- B1 $x_1 \in C$;
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- B4 $x_n \in C$.

Now consider (not necessarily disjoint) subsets I, O and C of G . We say that O is **AD-separated** from I by C , denoted as $AD(I, O | C)$, if every path $i \rightarrow x_1, \dots, x_n \leftarrow o$, $n \geq 1$, from a node $i \in I$ to a node $o \in O$, is blocked by C .

This asymmetrical version of AD-separation is similar to, yet different from both Moral's (2005) version of AD-separation and the notion of *l*-separation, as introduced by Vantaggi (2002). Our reason for not using one of these existing concepts is that our version of AD-separation has stronger properties: it satisfies all graphoid properties except symmetry; it satisfies redundancy, decomposition, weak union, contraction and intersection both in their direct and reverse forms.

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Irrelevant natural extension (III)

We define the **irrelevant natural extension** $\mathcal{G}_G^{\text{IR}}$ of a credal network as the most conservative (smallest) coherent set of desirable gambles on \mathcal{X}_G that (a) marginalises to the global local models $\mathcal{G}_{s, P(s)}$ (1) and (b) for which, for any $s \in G$, its non-causal non-descendant variables $X_{N(s)}$ are epistemically irrelevant to X_s conditional on $X_{P(s)}$ (II). If we use $PN(s)$ as a shorthand notation for $P(s) \cup N(s)$, then requirements (a) and (b) are equivalent to requiring that for all $s \in G$ and $x_{PN(s)} \in \mathcal{X}_{PN(s)}$

$$\text{margin}_s(\mathcal{G}_G^{\text{IR}} | x_{PN(s)}) = \text{margin}_s(\mathcal{G}_G^{\text{IR}} | x_{P(s)}) = \mathcal{G}_{s, P(s)}.$$

We show that this irrelevant natural extension is a linear operator that generates the set of all finite positive linear combinations of elements in its argument set

$$\mathcal{G}_G^{\text{IR}} := \{ \mathbf{1}_{\{x_{PN(s)} \in \mathcal{X}_{PN(s)}, f \in \mathcal{G}_{s, P(s)}\}} : s \in G, x_{PN(s)} \in \mathcal{X}_{PN(s)}, f \in \mathcal{G}_{s, P(s)} \}.$$

Marginalisation properties (IV)

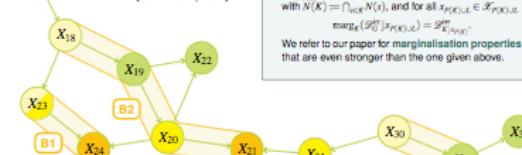
For any $K \subseteq G$, we construct a sub-DAG of the original DAG by eliminating the nodes $s \in G \setminus K$ and their associated edges. The parents of a node $s \in K$, with respect to this sub-DAG, are denoted by $P_K(s) := P(s) \cap K$. We derive local models $\mathcal{G}_{s, P_K(s)}$ for this sub-DAG from the original local models $\mathcal{G}_{s, P(s)}$ by fixing $x_{P(s) \setminus K}$. We do this consistently for all $s \in K$ at once by fixing $x_{P(K)}$, where $P(K) := (\bigcup_{s \in K} P(s)) \setminus K$. For any $x_{P(K)} \in \mathcal{X}_{P(K)}$, we use the resulting local models to construct an irrelevant natural extension of the sub-DAG and denote it by $\mathcal{G}_{K, P(K)}^{\text{IR}}$.

Suppose now that K is closed subset of G , meaning that $P(K) \cap D(K) = \emptyset$, where $D(K) := (\bigcup_{s \in K} D(s)) \setminus K$ and for all $s \in K$, $D(s)$ are the descendants of s . Then $\mathcal{G}_{K, P(K)}^{\text{IR}}$ is related to $\mathcal{G}_G^{\text{IR}}$ as follows: for all $L \subseteq N(K)$, with $N(K) := \bigcap_{s \in K} N(s)$, and for all $x_{P(K) \cup L} \in \mathcal{X}_{P(K) \cup L}$,

$$\text{margin}_K(\mathcal{G}_G^{\text{IR}} | x_{P(K) \cup L}) = \mathcal{G}_{K, P(K)}^{\text{IR}}.$$

We refer to our paper for marginalisation properties that are even stronger than the one given above.

AD($I, O | C$)
IR($I, O | C$)



Graphoid properties (VI)

In a credal network under epistemic irrelevance, every **AD-separation** corresponds to a conditional epistemic irrelevance statement that is satisfied by the corresponding irrelevant natural extension $\mathcal{G}_G^{\text{IR}}$ for all $I, O, C \subseteq G$ such that $AD(I, O | C)$: we have $IR(I, O | C)$ (II) (we refer to our paper for even stronger results regarding factorisation and subset-relevance (II)). Furthermore, the properties of **AD-separation** (V) imply that the corresponding family of irrelevance statements in their direct and reverse form are equivalent. We want to stress that our proof for this result is based solely on marginalisation results (IV) and properties of AD-separation (V). At no point does it invoke graphoid properties of epistemic irrelevance. We believe that this provides a more rather more positive perspective on the well-known fact that epistemic irrelevance, in general, can fail some of the graphoid axioms (Cozman 2005).

Advantages of sets of desirable gambles



Credal networks under epistemic irrelevance using sets of desirable gambles

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We present a new approach to credal networks, which are graphical models that generalise Bayesian nets to deal with imprecise probabilities. Instead of applying the commonly used notion of strong independence, we replace it by the weaker notion of epistemic irrelevance. We show how local models of epistemic irrelevance allow us to construct global models of given local uncertainty models (\mathcal{G}), leading to an intuitive expression for the so-called irrelevant natural extension (III) of a network. In contrast with Cozman (2000) who introduced this notion in terms of credal sets, our main results are presented using the language of sets of desirable gambles (SDG). This has allowed us to derive a number of useful properties of the irrelevant natural extension. It has powerful marginalisation properties (IV) and satisfies all graphoid properties but symmetry, both in their direct and reverse forms (V & VI).

Sets of desirable gambles (SDG)

We will model a subject's beliefs about the value that a variable X_i assumes in some set \mathcal{X}_i by means of his behaviour, which generates (subjective) maps. If \mathcal{X}_i does not have atoms, we refer to the ratio quo. This results in a set of desirable gambles $\mathcal{G} = \mathcal{G}(\mathcal{X})$, where $\mathcal{G}(\mathcal{X})$ is the set of all gambles on \mathcal{X} . \mathcal{G} is called coherent if it satisfies the rationality requirements D1—D4 for all $f_1, f_2 \in \mathcal{G}(\mathcal{X})$ and all real $\lambda > 0$:

$$\begin{aligned} \text{D1 } f \leq 0 \Rightarrow f \notin \mathcal{G} \\ \text{D2 } f > 0 \Rightarrow f \in \mathcal{G} \\ \text{D3 } f \in \mathcal{G} \Rightarrow \lambda f \in \mathcal{G} \\ \text{D4 } f_1, f_2 \in \mathcal{G} \Rightarrow f_1 + f_2 \in \mathcal{G} \end{aligned}$$

Although they are not as well known as other (imprecise) probability models, sets of desirable gambles have definite advantages. To give a few examples: they are operational, are easily able to deal with conditioning on events with probability zero, allow for intuitive geometrically flavoured proofs and are more expressive than both credal sets and lower previsions (see our papers or the second poster for credal networks under epistemic irrelevance that use these alternative models).

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We show that this irrelevant natural extension is simple (AD-separable), where the 'post' operator generates the set of all finite positive linear combinations of elements in $\mathcal{G}_G^{\text{IR}}$.

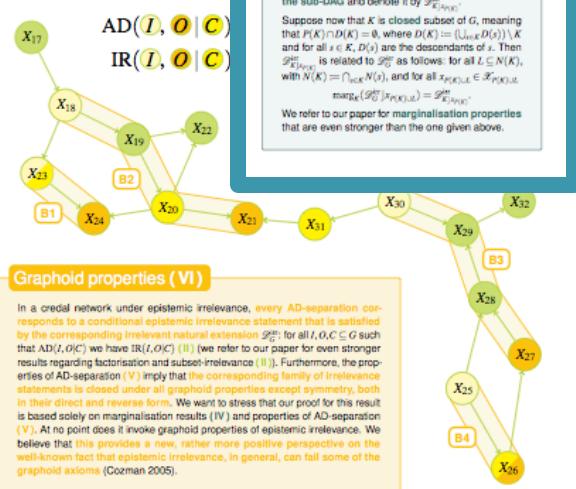
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Marginalisation properties

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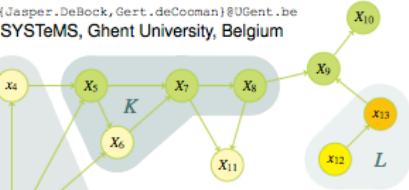
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$$\text{marg}_{\mathcal{G}}(\mathcal{G}_G^{\text{irr}} | x_{P(s)}) = \text{marg}_{\mathcal{G}}(\mathcal{G}_{(s,P(s))} | x_{P(s)})$$

We show that this irrelevant natural extension is simple to construct: $\mathcal{G}_G^{\text{irr}} = \text{posil}(\mathcal{A}_G^{\text{irr}})$, where the 'posil'-operator generates the set of all finite positive linear combinations of elements in its argument set

$$\mathcal{A}_G^{\text{irr}} := \{ \mathbf{I}_{(s,P(s))} f : s \in G, x_{P(s),s} \in \mathcal{X}_{P(s),s}, f \in \mathcal{G}_{(s,P(s))} \}$$

Epistemic irrelevance (II)

Consider a global set of desirable gambles $\mathcal{G}_G(\mathcal{X})$ on \mathcal{X}_G (I) and disjoint subsets S and K of G . Then the marginal model for X_S conditional on the information that X_K assumes a value x_K is given by

$$\text{marg}_{\mathcal{G}}(\mathcal{G}_G(x_K) - \{f \in \mathcal{G}(\mathcal{X}) : x_{K \setminus \{i\}} \in \mathcal{X}_i\})$$

Consider now three subsets $S, I, O \subseteq G$, with $I \subset C$ and $O \subset C$ disjoint. We say that X_I is epistemically irrelevant to X_O conditional on X_S , denoted as $IR(I, O | C)$, if and only if for all $x_{O \setminus I} \in \mathcal{X}_{O \setminus I}$ we have

$$\text{marg}_{\mathcal{G}_G}(\mathcal{G}_G(x_C) - \text{marg}_{\mathcal{G}}(\mathcal{G}_G | x_C))$$

Our paper also considers epistemical subset-irrelevance, which although interesting, is not discussed on this poster.

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AD($I, O | C$)

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Credal networks under epistemic irrelevance using sets of desirable gambles

Abstract of the paper

We present a new approach to credal networks, which are graphical models that generalise Bayesian nets to deal with imprecise probabilities. Instead of applying the commonly used notion of strong independence, we replace it by the weaker notion of epistemic irrelevance. We show how sets of desirable gambles allow us to construct global models of given local uncertainty models (I), leading to an intuitive expression for the so-called irrelevant natural extension (III) of a network. In contrast with Cozman (2000) who introduced this notion in terms of credal sets, our main results are presented using the language of sets of desirable gambles (SDG). This has allowed us to derive a number of useful properties of the irrelevant natural extension. It has powerful marginalisation properties (IV) and satisfies all graphoid properties but symmetry, both in their direct and reverse forms (V & VI).

Sets of desirable gambles (SDG)

We will model a subject's beliefs about the value that a variable X_i assumes in some set \mathcal{X}_i by means of his behaviour, which gambles (ordered maps) f . $\mathcal{G}(x)$ does not have to be the entire quo. This results in a set of desirable gambles $\mathcal{G} \subseteq \mathcal{G}(x)$, where $\mathcal{G}(x)$ is the set of all gambles on \mathcal{X} . \mathcal{G} is called coherent if it satisfies the rationality requirements D1—D4 for all $f_1, f_2 \in \mathcal{G}(x)$ and all real $\lambda > 0$.

- D1 $f \leq 0 \Rightarrow f \notin \mathcal{G}$
- D2 $f > 0 \Rightarrow f \in \mathcal{G}$
- D3 $f \in \mathcal{G} \Rightarrow \lambda f \in \mathcal{G}$
- D4 $f_1, f_2 \in \mathcal{G} \Rightarrow f_1 + f_2 \in \mathcal{G}$

Although they are not as well known as other (imprecise) probability models, sets of desirable gambles have definite advantages. To give a few examples: they are operational, are easily able to deal with conditioning on events with probability zero, allow for intuitive geometrically flavoured proofs and are more expressive than both credal sets and lower previsions (see our papers or the second poster for credal networks under epistemic irrelevance that use these alternative models).

Local uncertainty models (I)

With every node x of a finite directed acyclic graph (DAG), we associate a variable X_x taking values in some finite, non-empty set \mathcal{X}_x . The set of all nodes is denoted by G . For every subset $S \subseteq G$, the joint variable X_S takes values in $\mathcal{X}_S := \times_{x \in S} \mathcal{X}_x$. For every $i \in G$, we denote by $P(i)$ the set consisting of the parent nodes of x . Similar to what is done in classical Bayesian networks, we attach local uncertainty models to the nodes of the network, conditional on the values of their parents. For all $s \in G$ and every instantiation $x_{P(s)} \in \mathcal{X}_{P(s)}$, we require a coherent set $\mathcal{G}_{s, P(s)}$ of desirable gambles (SDG) on \mathcal{X}_s .

AD-separation (V)

Consider any path x_1, \dots, x_n in G , with $n \geq 1$. We say that this path is **blocked** by a set of nodes $C \subseteq G$ whenever at least one of the following four conditions holds:

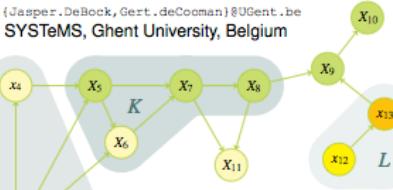
- B1 $x_1 \in C$;
- B2 there is some $1 < i < n$ such that $x_i \rightarrow x_{i+1}$ and $x_i \in C$;
- B3 there is some $1 < i < n$ such that $x_{i-1} \rightarrow x_i \leftarrow x_{i+1}$, $x_i \notin C$ and $D(x_i) \cap C = \emptyset$;
- B4 $x_n \in C$.

Now consider (not necessarily disjoint) subsets I, O and C of G . We say that O is AD-separated from I by C , denoted as $AD(I, O | C)$, if every path $i \rightarrow x_1, \dots, x_n \rightarrow o$, $n \geq 1$, from a node $i \in I$ to a node $o \in O$ is blocked by C .

This asymmetrical version of D-separation is similar to, yet different from both Moral's (2005) version of AD-separation and the notion of I-separation, as introduced by Vantaggi (2002). Our reason for not using one of these existing concepts is that our version of AD-separation has stronger properties: it satisfies all graphoid properties except symmetry; it satisfies redundancy, decomposition, weak union, contraction and intersection both in their direct and reverse forms.

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Irrelevant natural extension (III)

We define the **irrelevant natural extension** $\mathcal{G}_G^{\text{IR}}$ of a credal network as the most conservative (smallest) coherent set of desirable gambles on \mathcal{X}_G that (a) marginalises to the global model of the original model $\mathcal{G}_{s, P(s)}$ (1) and (b) for which, for any $x \in G$, its non-causal non-descendant variables $X_{N(x)}$ are epistemically irrelevant to X_x conditional on $X_{P(x)}$ (2). If we use $P(x)$ as a shorthand notation for $P(x) \cup N(x)$, then requirements (a) and (b) are equivalent to requiring that for all $s \in G$ and $x_{P(s)} \in \mathcal{X}_{P(s)}$,

$$\text{marg}_s(\mathcal{G}_G^{\text{IR}} | x_{P(s)}) = \text{marg}_s(\mathcal{G}_{s, P(s)}^{\text{IR}} | x_{P(s)}) = \mathcal{G}_{s, P(s)}^{\text{IR}}.$$

We show that this irrelevant natural extension is simple to construct: $\mathcal{G}_G^{\text{IR}} := \text{posl}(\mathcal{A}_G^{\text{IR}})$, where the 'posl'-operator generates the set of all finite positive linear combinations of elements in its argument set

$$\mathcal{A}_G^{\text{IR}} := \{ \mathbf{I}_{\{x \in G\}} f : s \in G, x_{P(s)} \in \mathcal{X}_{P(s)}, f \in \mathcal{G}_{s, P(s)} \}.$$

Epistemic irrelevance (II)

Consider a global set of desirable gambles $\mathcal{G}_G^{\text{SDG}}$ on \mathcal{X}_G (1) and disjoint subsets S and K of G . Then the marginal model for X_S conditional on the information that X_K assumes a value $x_K \in \mathcal{X}_K$, is given by

$$\text{marg}_S(\mathcal{G}_G^{\text{SDG}} | x_K) = \{ f \in \mathcal{G}(X_S) : \mathbf{I}_{\{x_K\}} f \in \mathcal{G}_S \}.$$

Consider now three subsets $S, I, O \subseteq G$, with $I \cup O \subseteq C$ disjoint. We say that X_I is epistemically irrelevant to X_O conditional on X_C , denoted as $IR(I, O | C)$, if and only if for all $x_{C \cup O} \in \mathcal{X}_{C \cup O}$ we have

$$\text{marg}_{O \cup C}(\mathcal{G}_G^{\text{SDG}} | x_C) = \text{marg}_{O \cup C}(\mathcal{G}_G^{\text{SDG}} | x_C).$$

Our paper also considers epistemical subset-irrelevance, which although interesting, is not discussed on this poster.

Marginalisation properties (IV)

For any $K \subseteq G$, we construct a sub-DAG of the original DAG by eliminating the nodes $s \in G \setminus K$ and their associated edges. The parents of a node $s \in K$, with respect to this sub-DAG, are denoted by $P_K(s) := P(s) \cap K$. We derive local models $\mathcal{G}_{s, P(s)}^{\text{IR}}$ for this sub-DAG from the original local models $\mathcal{G}_{s, P(s)}$ by fixing $x_{P(s) \setminus K}$. We do this consistently for all $s \in K$ at once by fixing $x_{P(K)}$, where $P(K) := (\bigcup_{s \in K} P(s)) \setminus K$. For any $x_{P(K)} \in \mathcal{X}_{P(K)}$, we use the resulting local models to construct an irrelevant natural extension of the sub-DAG and denote it by $\mathcal{G}_{K, P(K)}^{\text{IR}}$.

Suppose now that K is closed subset of G , meaning that $P(K) \cap D(K) = \emptyset$, where $D(K) := (\bigcup_{s \in K} D(s)) \setminus K$ and for all $s \in K$, $D(s)$ are the descendants of s . Then $\mathcal{G}_{K, P(K)}^{\text{IR}}$ is related to $\mathcal{G}_G^{\text{IR}}$ as follows: for all $L \subseteq K$, with $N(L) = \bigcap_{s \in L} N(s)$, and for all $x_{P(L) \cup L} \in \mathcal{X}_{P(L) \cup L}$,

$$\text{marg}_L(\mathcal{G}_G^{\text{IR}} | x_{P(L) \cup L}) = \mathcal{G}_{L, P(L)}^{\text{IR}}.$$

We refer to our paper for marginalisation properties that are even stronger than the one given above.

AD-separation (I)

IR (I, O | C)

Graphoid properties (VI)

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In a credal network under epistemic irrelevance, every AD-separation corresponds to a conditional epistemic irrelevance statement that is satisfied by the corresponding irrelevant natural extension $\mathcal{G}_G^{\text{IR}}$ for all $I, O, C \subseteq G$ such that $AD(I, O | C)$ (we refer to our paper for even stronger results regarding factorisation and subset-irrelevance (II)). Furthermore, the properties of AD-separation (V) imply that the corresponding kind of irrelevance statement is also satisfied in their direct and reverse form. We want to stress that our proof for this result is based solely on marginalisation results (IV) and properties of AD-separation (V). At no point does it invoke graphoid properties of epistemic irrelevance. We believe that this provides a more rather more positive perspective on the well-known fact that epistemic irrelevance, in general, can fail some of the graphoid axioms (Cozman 2005).

Advantages of sets of desirable gambles

Marginalisation properties

AD-separation satisfies every graphoid property except symmetry

AD-separation implies epistemic irrelevance

Credal networks: useful properties ✓



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~~graphoid
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Credal networks under
epistemic irrelevance

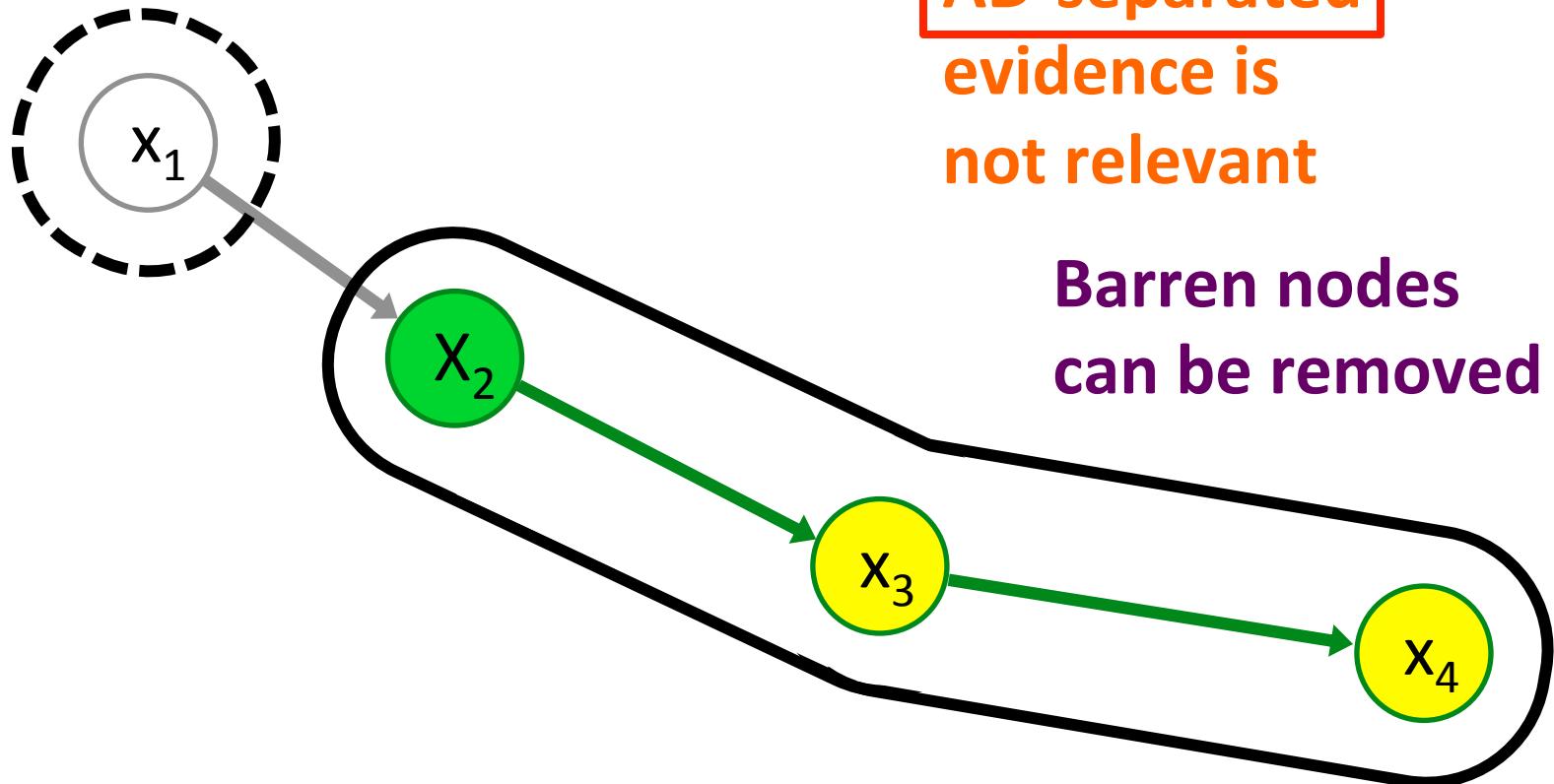


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Hope to see you at the poster session!

